

Supplementary Material for “On the Asymptotic Efficiency of Approximate Bayesian Computation Estimators”

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Technical lemmas and proofs of the main results are presented. In the following, the data are considered to be random, and $O(\cdot)$ and $\Theta(\cdot)$ denote the limiting behaviour when n goes to ∞ . For two sets A and B , the sum of integrals $\int_A f(x) dx + \int_B f(x) dx$ is written as $(\int_A + \int_B) f(x) dx$. Recall that $T_{\text{obs}} = a_n A(\theta_0)^{-1/2} \{s_{\text{obs}} - s(\theta_0)\}$ and by Condition 4, $T_{\text{obs}} \rightarrow N(0, I_d)$ in distribution, where I_d is the identity matrix with dimension d . For a constant $d \times p$ matrix A , let the minimum and maximum eigenvalues of $A^T A$ be $\lambda_{\min}^2(A)$ and $\lambda_{\max}^2(A)$. Obviously for any p -dimension vector x , $\lambda_{\min}(A) \|x\| \leq \|Ax\| \leq \lambda_{\max}(A) \|x\|$. For two matrices A and B , we say A is bounded by B if $\lambda_{\max}(A) \leq \lambda_{\min}(B)$.

1. PROOF OF RESULTS FROM SECTION 3

1.1. Overview and Notation

We first give an overview of the proof to Theorem 1. The convergence of the maximum likelihood estimator based on the summary follows almost immediately from Creel & Kristensen (2013). The minor extensions we used are summarized in Lemmas 1 and 2 below.

The main challenge with Theorem 1 are the results about the posterior mean of approximate Bayesian computation. For the convergence of posterior means of approximate Bayesian computation we need to consider convergence of integrals over the parameter space, \mathbb{R}^p . We will divide \mathbb{R}^p into $B_\delta = \{\theta : \|\theta - \theta_0\| < \delta\}$ and B_δ^c for some $\delta < \delta_0$, and introduce the notation $\pi(h) = \int h(\theta) \pi(\theta) f_{\text{ABC}}(s_{\text{obs}} | \theta) d\theta$. The posterior mean of approximate Bayesian computation is $h_{\text{ABC}} = \pi(h) / \pi(1)$. We can write $\pi(h)$, say, as $\pi(h) = \pi_{B_\delta}(h) + \pi_{B_\delta^c}(h)$, where

$$\pi_{B_\delta}(h) = \int_{B_\delta} h(\theta) \pi(\theta) f_{\text{ABC}}(s_{\text{obs}} | \theta) d\theta, \quad \pi_{B_\delta^c}(h) = \int_{B_\delta^c} h(\theta) \pi(\theta) f_{\text{ABC}}(s_{\text{obs}} | \theta) d\theta.$$

As $n \rightarrow \infty$ the posterior distribution of approximate Bayesian computation concentrates around θ_0 . The first step of our proof is to show that, as a result, the contribution that comes from integrating over B_δ^c can be ignored. Hence we need consider only $\pi_{B_\delta}(h) / \pi_{B_\delta}(1)$.

Second, we perform a Taylor expansion of $h(\theta)$ around θ_0 . Let $Dh(\theta)$ and $Hh(\theta)$ denote the vector of first derivatives and the matrix of second derivatives of $h(\theta)$ respectively. Then

$$h(\theta) = h(\theta_0) + Dh(\theta_0)^T (\theta - \theta_0) + \frac{1}{2} (\theta - \theta_0)^T Hh(\theta_*) (\theta - \theta_0),$$

49 for some θ_* , that depends on θ and that satisfies $\|\theta_* - \theta_0\| < \|\theta - \theta_0\|$. We plug this into
 50 $\pi_{B_\delta}(h)$, but re-express the integrals in term of the rescaled random vector

$$51 \quad t(\theta) = a_{n,\varepsilon}(\theta - \theta_0),$$

52 and let $t(B_\delta)$ be the set $\{\phi : \phi = t(\theta) \text{ for some } \theta \in B_\delta\}$. This gives

$$53 \quad \frac{\pi_{B_\delta}(h)}{\pi_{B_\delta}(1)} = h(\theta_0) + a_{n,\varepsilon}^{-1} Dh(\theta_0)^T \frac{\pi_{B_\delta}(t)}{\pi_{B_\delta}(1)} + \frac{1}{2} a_{n,\varepsilon}^{-2} \frac{\pi_{B_\delta}\{t^T H h(\theta_t)t\}}{\pi_{B_\delta}(1)}, \quad (1)$$

54 where we write t for $t(\theta)$, and θ_t is the value θ_* from remainder term in the Taylor expansion for
 55 $h(\theta)$. We use the notation θ_t to emphasize its dependence on t , and note that θ_t belongs to B_δ .

56 Let $\tilde{f}_{\text{ABC}}(s_{\text{obs}} | \theta) = \int \tilde{f}_n(s_{\text{obs}} + \varepsilon_n v | \theta) K(v) dv$, which is the likelihood approximation
 57 that we get if we replace the true likelihood by its Gaussian limit, and define $\tilde{\pi}_{B_\delta}(h) =$
 58 $\int_{B_\delta} h(\theta) \pi(\theta) \tilde{f}_{\text{ABC}}(s_{\text{obs}} | \theta) d\theta$. Our third step is to re-write (1) as

$$59 \quad \frac{\pi_{B_\delta}(h)}{\pi_{B_\delta}(1)} = h(\theta_0) + a_{n,\varepsilon}^{-1} Dh(\theta_0)^T \frac{\tilde{\pi}_{B_\delta}(t)}{\tilde{\pi}_{B_\delta}(1)} + a_{n,\varepsilon}^{-1} Dh(\theta_0)^T \left\{ \frac{\tilde{\pi}_{B_\delta}(t)}{\tilde{\pi}_{B_\delta}(1)} - \frac{\pi_{B_\delta}(t)}{\pi_{B_\delta}(1)} \right\}$$

$$60 \quad + \frac{1}{2} a_{n,\varepsilon}^{-2} \frac{\pi_{B_\delta}\{t^T H h(\theta_t)t\}}{\pi_{B_\delta}(1)}.$$

61 We bound the size of the last two terms, so that asymptotically h_{ABC} behaves as

$$62 \quad h(\theta_0) + a_{n,\varepsilon}^{-1} Dh(\theta_0)^T \frac{\tilde{\pi}_{B_\delta}(t)}{\tilde{\pi}_{B_\delta}(1)}.$$

63 If we introduce the density $g_n(t, v)$, defined as $g_n(t, v, \tau)$ in Section 4.3 of the main text but with
 64 $\tau = 0$, so

$$65 \quad g_n(t, v) \propto \begin{cases} N\left\{Ds(\theta_0)t; a_n \varepsilon_n v + A(\theta_0)^{1/2} T_{\text{obs}}, A(\theta_0)\right\} K(v), & a_n \varepsilon_n \rightarrow c < \infty, \\ N\left\{Ds(\theta_0)t; v + \frac{1}{a_n \varepsilon_n} A(\theta_0)^{1/2} T_{\text{obs}}, \frac{1}{a_n^2 \varepsilon_n^2} A(\theta_0)\right\} K(v), & a_n \varepsilon_n \rightarrow \infty, \end{cases}$$

66 then we can show that

$$67 \quad \frac{\tilde{\pi}_{B_\delta}(t)}{\tilde{\pi}_{B_\delta}(1)} \approx \frac{\int_{t(B_\delta)} \int_{\mathbb{R}^d} t g_n(t, v) dt dv}{\int_{t(B_\delta)} \int_{\mathbb{R}^d} g_n(t, v) dt dv},$$

68 with a remainder that can be ignored. Putting this together, we get that asymptotically h_{ABC} is

$$69 \quad h(\theta_0) + a_{n,\varepsilon}^{-1} Dh(\theta_0)^T \frac{\int_{t(B_\delta)} \int_{\mathbb{R}^d} t g_n(t, v) dt dv}{\int_{t(B_\delta)} \int_{\mathbb{R}^d} g_n(t, v) dt dv},$$

70 and the proof finishes by calculating the form of this.

71 A recurring theme in the proofs for the bounds on the various remainders is the need to bound
 72 expectations of polynomials of either the rescaled parameter t , or a rescaled difference in the
 73 summary statistic from s_{obs} , or both. Later we will present a lemma, stated in terms of a general
 74 polynomial, that is used repeatedly to obtain the bounds we need.

75 To define this we need to introduce a set of suitable polynomials. For any integer l and vector
 76 x , if a scalar function of x has the expression $\sum_{i=0}^l \alpha_i(x, n)^T x^i$, where for each i , x^i denotes the
 77 vector with all monomials of x with degree i as elements and $\alpha_i(x, n)$ is a vector of functions of
 78 x and n , we denote it by $P_l(x)$. Let $\mathbb{P}_{l,x}$ be the set

$$79 \quad \{P_l(x) : \text{for all } i \leq l, \text{ as } n \rightarrow \infty, \alpha_i(x, n) = O_p(1) \text{ holds uniformly in } x\}$$

To simplify the notations, for two vectors x_1 and x_2 , $P_l\{(x_1^T, x_2^T)^T\}$ and $\mathbb{P}_{l,(x_1^T, x_2^T)^T}$ are written as $P_l(x_1, x_2)$ and $\mathbb{P}_{l,(x_1, x_2)}$. Where the specific form of the polynomial does not matter, and we only use the fact that it lies in $\mathbb{P}_{l,x}$, we will often simplify expressions by writing it as $P_l(x)$.

1.2. Proof of Theorem 1

For the maximum likelihood estimator based on the summary, Creel & Kristensen (2013) gives the central limit theorem for $\hat{\theta}_{\text{MLEs}}$ when $a_n = n^{1/2}$ and \mathcal{P} is compact. According to the proof in Creel & Kristensen (2013), extending the result to the general a_n is straightforward. Additionally, we give the extension for general \mathcal{P} .

LEMMA 1. Assume Conditions 1,4-6. Then $a_n(\hat{\theta}_{\text{MLEs}} - \theta_0) \rightarrow N\{0, I^{-1}(\theta_0)\}$ in distribution as $n \rightarrow \infty$.

Given Condition 3, by Lemma 1 and the delta method (Lehmann, 2004), the convergence of the maximum likelihood estimator for general $h(\theta)$ holds as follows.

LEMMA 2. Assume the conditions of Lemma 1 and Condition 3. Then $a_n\{h(\hat{\theta}_{\text{MLEs}}) - h(\theta_0)\} \rightarrow N\{0, Dh(\theta_0)^T I^{-1}(\theta_0) Dh(\theta_0)\}$ in distribution as $n \rightarrow \infty$.

The following lemmas are used for the result about the posterior mean of approximate Bayesian computation, proofs of these are given in Section 1.3. Our first lemma is used to justify ignoring integrals over B_δ^c .

LEMMA 3. Assume Conditions 2, 3-6. Then for any $\delta < \delta_0$, $\pi_{B_\delta^c}(h) = O_p(e^{-a_n^{\alpha_\delta} c_\delta})$ for some positive constants c_δ and α_δ depending on δ .

The following lemma is used to calculate the form of

$$\frac{\int_{t(B_\delta)} \int_{\mathbb{R}^d} t g_n(t, v) dt dv}{\int_{t(B_\delta)} \int_{\mathbb{R}^d} g_n(t, v) dt dv},$$

which is the leading term for $\{h_{\text{ABC}} - h(\theta_0)\}$.

LEMMA 4. Assume Condition 2. Let c be a constant vector, $\{k_n\}$ be a series converging to $k_\infty \in (0, \infty]$ and $\{b'_n\}$ be a series converging to a non-negative constant. Let $b_n = \mathbb{1}_{\{k_\infty = \infty\}} + b'_n \mathbb{1}_{\{k_\infty < \infty\}}$. Then for any $d \times p$ constant matrix A and any $d \times d$ constant matrix B ,

$$\int_{\mathbb{R}^p} \int_{\mathbb{R}^d} t \frac{N(At; B_n v + \frac{1}{k_n} c, \frac{1}{k_n^2} I_d) K(v)}{\int_{\mathbb{R}^p} \int_{\mathbb{R}^d} N(At; B_n v + \frac{1}{k_n} c, \frac{1}{k_n^2} I_d) K(v) dt dv} dt dv = \frac{1}{k_n} \{(A^T A)^{-1} A^T c + R(A, B_n, k_n, c)\},$$

where $B_n = b_n B$, the expression of $R(c; A, B_n, k_n)$ is stated in the proof. Specifically, $R(A, B_n, k_n, c) = o(1)$ when $B_n = o(1)$ and $O(1)$ otherwise.

Our final two lemmas are used to bound the remainder terms in the expansion for h_{ABC} we presented in Section 1.1.

LEMMA 5. Assume Conditions 1, 2 and 4 hold. If $\varepsilon_n = o(a_n^{-1/2})$, there exists a $\delta < \delta_0$ such that

$$\begin{aligned} \tilde{\pi}_{B_\delta}(1) &= a_{n,\varepsilon}^{d-p} \left\{ \pi(\theta_0) \int_{t(B_\delta)} \int_{\mathbb{R}^d} g_n(t, v) dv dt + O_p(a_{n,\varepsilon}^{-1}) + O_p(a_n^2 \varepsilon_n^4) \right\}, \\ \int_{t(B_\delta)} \int_{\mathbb{R}^d} g_n(t, v) dt dv &= \Theta_p(1), \\ \frac{\tilde{\pi}_{B_\delta}(t)}{\tilde{\pi}_{B_\delta}(1)} &= \frac{\int_{t(B_\delta)} \int_{\mathbb{R}^d} t g_n(t, v) dt dv}{\int_{t(B_\delta)} \int_{\mathbb{R}^d} g_n(t, v) dt dv} + O_p(a_{n,\varepsilon}^{-1}) + O_p(a_n^2 \varepsilon_n^4), \end{aligned} \quad (2)$$

and $\tilde{\pi}_{B_\delta}\{P_2(t)\}/\tilde{\pi}_{B_\delta}(1) = O_p(1)$ for any $P_2(t) \in \mathbb{P}_{2,t}$.

LEMMA 6. Assume the conditions of Lemma 5 and Conditions 3 and 5. Then if $\varepsilon_n = o(a_n^{-1/2})$, there exists a $\delta < \delta_0$ such that

$$\frac{\pi_{B_\delta}(h)}{\pi_{B_\delta}(1)} = h(\theta_0) + a_{n,\varepsilon}^{-1} Dh(\theta_0)^T \left\{ \frac{\tilde{\pi}_{B_\delta}(t)}{\tilde{\pi}_{B_\delta}(1)} + O_p(\alpha_n^{-1}) \right\} + \frac{1}{2} a_{n,\varepsilon}^{-2} \left[\frac{\tilde{\pi}_{B_\delta}\{t^T H h(\theta_t) t\}}{\tilde{\pi}_{B_\delta}(1)} + O_p(\alpha_n^{-1}) \right], \quad (3)$$

Now we are ready to prove Theorem 1.

Proof of Theorem 1. The convergence of the maximum likelihood estimator based on the summary is given by Lemma 1 and Lemma 2.

We now focus on the convergence for the posterior mean of approximate Bayesian computation. The convergence of the posterior mean given the summaries follows from a similar, but simpler, argument and is omitted.

We can bound $t^T H(\theta_t) t$ for θ in B_δ by the quadratic $t^T H_{max} t$, where H_{max} is an upper bound on $H(\theta_t)$ for θ_t in B_δ . This means that

$$\tilde{\pi}_{B_\delta}\{t^T H h(\theta_t) t\} = O(1).$$

Together with Lemmas 3, 5 and 6, we then have the expansion

$$h_{ABC} = h(\theta_0) + a_{n,\varepsilon}^{-1} Dh(\theta_0)^T \left\{ \frac{\int_{t(B_\delta) \times \mathbb{R}^d} t g_n(t, v) dt dv}{\int_{t(B_\delta) \times \mathbb{R}^d} g_n(t, v) dt dv} + O_p(a_{n,\varepsilon}^{-1}) + O_p(a_n^2 \varepsilon_n^4) + O_p(\alpha_n^{-1}) \right\}.$$

The analytical form of the integral in the above expansion, which we will denote by $E_{g_n}(t)$, can be obtained by applying Lemma 4 with $A = A(\theta_0)^{-1/2} DS(\theta_0)$, $c = T_{obs}$,

$$B_n = \begin{cases} a_n \varepsilon_n A(\theta_0)^{-1/2}, & c_\varepsilon < \infty, \\ A(\theta_0)^{-1/2}, & c_\varepsilon = \infty, \end{cases} \quad k_n = \begin{cases} 1, & c_\varepsilon < \infty, \\ a_n \varepsilon_n, & c_\varepsilon = \infty. \end{cases}$$

It can be seen that $E_{g_n}(t)$ is $\Theta_p(k_n^{-1})$, and the remainder term, $O_p(a_{n,\varepsilon}^{-1}) + O_p(a_n^2 \varepsilon_n^4) + O_p(\alpha_n^{-1})$, is $o_p(1)$ as $\varepsilon_n = o(a_n^{-3/5})$ and $\alpha_n^{-1} = o(a_n^{-2/5})$. Then since $a_{n,\varepsilon}^{-1} k_n^{-1} = a_n^{-1}$, we have

$$\begin{aligned} & a_n \{h_{ABC} - h(\theta_0)\} \\ &= Dh(\theta_0)^T \left[\left\{ Ds(\theta_0)^T A(\theta_0)^{-1} Ds(\theta_0) \right\}^{-1} Ds(\theta_0)^T A(\theta_0)^{-1/2} T_{obs} + R_n(a_n \varepsilon_n, T_{obs}) \right] + o_p(1), \end{aligned} \quad (4)$$

193 where $R_n(a_n\varepsilon_n, T_{\text{obs}})$ is $Dh(\theta_0)^T R(A, B_n, k_n, c)$ with $R(A, B_n, k_n, c)$ defined in Lemma 4.
 194 We can interpret $R_n(a_n\varepsilon_n, T_{\text{obs}})$ as the extra variation brought by ε_n : $a_n[h_{\text{ABC}} - E\{h(\theta) |$
 195 $s_{\text{obs}}\}]$.

196 By the delta method, the first term in the right hand side of (4) converges to $I(\theta_0)^{-1/2}Z$. For
 197 the second term, since $A(A^T A)^{-1}A^T$ is a projection matrix, by eigen decomposition

$$198 \quad I - A(A^T A)^{-1}A^T = U \begin{pmatrix} 0 & 0 \\ 0 & I_{d-p} \end{pmatrix} U^T, \quad (A^T A)^{-1/2}A^T = \begin{pmatrix} I_p & 0 \\ 0 & 0 \end{pmatrix} U^T,$$

199 where U is an orthogonal matrix. For a vector x , let $x_{k_1:k_2}$ be the $(k_2 - k_1 + 1)$ -dimension vector
 200 containing the k_1 th– k_2 th coordinates of x . Let $v' = U^T A(\theta_0)^{-1/2}v$, and $T'_{\text{obs}} = U^T T_{\text{obs}}$. Then
 201 $R_n(a_n\varepsilon_n, T_{\text{obs}})$ can be written as

$$202 \quad R_n(a_n\varepsilon_n, T_{\text{obs}}) \\
 203 = Dh(\theta_0)^T (A^T A)^{-1/2} a_n \varepsilon_n \frac{\int v'_{1:p} N\{v'_{(p+1):d}; -\frac{1}{a_n \varepsilon_n} T'_{\text{obs},(p+1):d}, \frac{1}{a_n^2 \varepsilon_n^2} I_{d-p}\} K\{A(\theta_0)^{1/2} U v'\} dv'}{\int N\{v'_{(p+1):d}; -\frac{1}{a_n \varepsilon_n} T'_{\text{obs},(p+1):d}, \frac{1}{a_n^2 \varepsilon_n^2} I_{d-p}\} K\{A(\theta_0)^{1/2} U v'\} dv'}. \quad (5)$$

204 Denote the weak limit of $R_n(a_n\varepsilon_n, T_{\text{obs}})$ as $R(c_\varepsilon, Z)$. When $d = p$, obviously $R_n(a_n\varepsilon_n, T_{\text{obs}}) =$
 205 0 and therefore $R(c_\varepsilon, Z) = 0$. When $d > p$, if $\varepsilon_n = o(1/a_n)$, $R_n(a_n\varepsilon_n, T_{\text{obs}}) = o_p(1)$ by
 206 Lemma 4 and therefore $R(c_\varepsilon, Z) = 0$. When the covariance matrix of $K(\cdot)$ is $c^2 A(\theta_0)$, for con-
 207 stant $c > 0$, $K(v) \propto \bar{K}\{c\|A(\theta_0)^{-1/2}v\|^2\}$. Then $K\{A(\theta_0)^{1/2}Uv'\}$ in (5) can be replaced by
 208 $\bar{K}\{c\|v'\|^2\}$ and for fixed $v'_{(p+1):d}$, the integrand in the numerator, as a function of $v'_{1:p}$, is sym-
 209 metric around zero. Therefore $R_n(a_n\varepsilon_n, T_{\text{obs}}) = 0$ and $R(c_\varepsilon, Z) = 0$.

210 Otherwise, $R_n(a_n\varepsilon_n, z)$ is not necessarily zero. Since for any n , $R_n(a_n\varepsilon_n, z)$ as a function of
 211 z is symmetric around 0, $R(c_\varepsilon, z)$ is also symmetric and $R(c_\varepsilon, Z)$ has mean zero. Since $I^{-1}(\theta_0)$
 212 is the Cramer-Rao lower bound, $\text{var}\{I(\theta_0)^{-1/2}Z + R(c_\varepsilon, Z)\} \geq I^{-1}(\theta_0)$.

213 For (i), the asymptotic normality holds for $h(\hat{\theta})$ by Lemma 2. □

214 1.3. Proof of Lemmas

215 Here we give the proofs of lemmas from Section 1.2.

216 *Proof of Lemma 3.* It is sufficient to show that for any δ , $\sup_{\theta \in B_\delta^c} f_{\text{ABC}}(s_{\text{obs}} | \theta) =$
 217 $O_p(e^{-a_n^{\alpha_\delta} c \delta})$. By dividing \mathbb{R}^d into $\{v : \|\varepsilon_n v\| \leq \delta'/3\}$ and its complement, we have

$$218 \quad \sup_{\theta \in B_\delta^c} f_{\text{ABC}}(s_{\text{obs}} | \theta) = \sup_{\theta \in B_\delta^c} \int_{\mathbb{R}^d} f_n(s_{\text{obs}} + \varepsilon_n v | \theta) K(v) dv \\
 219 \leq \sup_{\theta \in B_\delta^c \setminus \mathcal{P}_0^c} \left\{ \sup_{\|s - s_{\text{obs}}\| \leq \delta'/3} f_n(s | \theta) \right\} + \sup_{\theta \in \mathcal{P}_0^c} \left\{ \sup_{\|s - s_{\text{obs}}\| \leq \delta'/3} f_n(s | \theta) \right\} + \bar{K}(\lambda_{\min}(\Lambda) \varepsilon_n^{-1} \delta'/3) \varepsilon_n^{-d},$$

220 where $\lambda_{\min}(\Lambda)$ is positive. In the above, as $n \rightarrow \infty$, the third term is exponentially decreasing
 221 by Conditions 2(iv). For the second term, by Condition 4, with probability 1,

$$222 \quad \|s - s(\theta)\| = \|\{s(\theta_0) - s(\theta)\} + \{s_{\text{obs}} - s(\theta_0)\} + \varepsilon_n v\| \\
 223 \geq \delta' - \delta'/3 - \delta'/3 = \delta'/3.$$

224 Recall that $W_n(s) = a_n A(\theta)^{-1/2}\{s - s(\theta)\}$. Then by Condition 6, the second term is expo-
 225 nentially decreasing. For the first term, when $\theta \in B_\delta^c \setminus \mathcal{P}_0^c$ and $\|s - s_{\text{obs}}\| \leq \delta'/3$, $\|W_n(s)\| \geq$
 226 $a_n \delta' r$ for some constant r . By Condition 5 and 6, $f_{W_n}(w | \theta)$ is bounded by the

241 sum of a normal density and $\alpha_n^{-1} r_{\max}(w)$, which are both exponentially decreasing, so
 242 $\sup_{\theta \in B_\delta^c \setminus \mathcal{P}_0^c} \sup_{\|s - s_{\text{obs}}\| \leq \delta'/3} f_n(s | \theta)$ is also exponentially decreasing. Finally, the sum of all
 243 the above is $O(e^{-a_n \varepsilon^{\alpha_n \delta} c \delta})$ by noting that $a_{n, \varepsilon} \leq \min(\varepsilon_n^{-1}, a_n)$. \square
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246 The following additional lemma will be used repeatedly to bound error terms that appear in
 247 Lemmas 5 and 6.
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250 LEMMA 7. Assume Condition 2. For $t \in \mathbb{R}^p$ and $v \in \mathbb{R}^d$, let $\{A_n(t)\}$ be a series of $d \times p$
 251 matrix functions, $\{C_n(t)\}$ be a series of $d \times d$ matrix functions, Q be a positive definite matrix
 252 and $g_1(v)$ and $g_2(v)$ be probability densities in \mathbb{R}^d . Let c be a random vector, $\{k_n\}$ be a series
 253 converging to $k_\infty \in (0, \infty]$ and $\{b'_n\}$ be a series converging to a non-negative constant. Let
 254 $b_n = \mathbb{1}_{\{k_\infty = \infty\}} + b'_n \mathbb{1}_{\{k_\infty < \infty\}}$. If

- 255 (i) $g_1(v)$ and $g_2(v)$ are bounded in \mathbb{R}^d ;
 256 (ii) $g_1(v)$ and $g_2(v)$ depend on v only through $\|v\|$ and are decreasing functions of $\|v\|$;
 257 (iii) there exists an integer l such that $\int \prod_{k=1}^{l+p} v_{i_k} g_j(v) dv < \infty$, $j = 1, 2$, for any coordinates
 258 $(v_{i_1}, \dots, v_{i_l})$ of v ;
 259 (iv) there exists a positive constant m such that for any $t \in \mathbb{R}^p$ and n , $\lambda_{\min}\{A_n(t)\}$ and
 260 $\lambda_{\min}\{C_n(t)\}$ are greater than m ;
 261 then for any $P_l(t, v) \in \mathbb{P}_{l, (t, v)}$,
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$$264 \int_{\mathbb{R}^p} \int_{\mathbb{R}^d} P_l(t, v) k_n^d g_1[k_n C_n(t) \{A_n(t)t - b_n v - k_n^{-1} c\}] g_2(Qv) dv dt = O_p(1),$$

$$265 \int_{\mathbb{R}^p} \int_{\mathbb{R}^d} k_n^d g_1[k_n C_n(t) \{A_n(t)t - b_n v - k_n^{-1} c\}] g_2(Qv) dv dt = \Theta_p(1).$$

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 270 *Proof.* For simplicity, here \int denotes the integration over the whole Euclidean space. Accord-
 271 ing to (ii), $g_1(v)$ can be written as $\bar{g}_1(\|v\|)$. When $k_\infty < \infty$, assume $k_n = 1$ without loss of gen-
 272 erality. For any $P_l(t, v) \in \mathbb{P}_{l, (t, v)}$, by Cauchy–Schwarz inequality, there exists a $P_l(\|t\|, \|v\|) \in$
 273 $\mathbb{P}_{l, (\|t\|, \|v\|)}$ with coefficient functions taking positive values such that $|P_l(t, v)|$ is bounded by
 274 $P_l(\|t\|, \|v\|)$ almost surely. Therefore for the first equality, it is sufficient to consider the equal-
 275 ity where $P_l(t, v)$ is replaced by $P_l(\|t\|, \|v\|)$ and the coefficient functions of $P_l(\|t\|, \|v\|)$ are
 276 positive almost surely. For each n , divide \mathbb{R}^p into $V = \{t : \|A_n(t)t\|/2 \geq \|b'_n v + c\|\}$ and V^c .
 277 In V , $\|C_n(t)\{A_n(t)t - b'_n v - c\}\| \geq m^2 \|t\|/2$; in V^c , $\|t\| \leq 2m^{-1} \|b'_n v + c\|$. With probability
 278 tending to 1,
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$$281 \int P_l(\|t\|, \|v\|) g_1[C_n(t) \{A_n(t)t - b'_n v - c\}] g_2(Qv) dv dt \leq$$

$$282 \int P_l(\|t\|, \|v\|) \bar{g}_1(m^2 \|t\|/2) g_2(Qv) dv dt + \sup_{v \in \mathbb{R}^d} g_1(v) \int \int_{V^c} dt P_l(2m^{-1} \|b'_n v + c\|, \|v\|) g_2(Qv) dv.$$

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 287 In the above, $\int_{V^c} dt$ is the volume of V^c in \mathbb{R}^p and is proportional to $\|b'_n v + c\|^p$. By (iii), the
 288 right hand side of the above inequality is $O_p(1)$.

When $k_\infty = \infty$, let $v^* = k_n\{A(t)t - v - k_n^{-1}c\}$. Then for any $P_l(t, v) \in \mathbb{P}_{l,(t,v)}$, with probability 1,

$$\begin{aligned} & \left| \int P_l(t, v) k_n^d g_1[k_n C_n(t)\{A(t)t - v - k_n^{-1}c\}] g_2(Qv) dv dt \right| \\ &= \left| \int P_l(t, v^*) g_2[Q\{A(t)t - k_n^{-1}v^* - k_n^{-1}c\}] g_1(C_n(t)v^*) dv^* dt \right|, \\ &\leq \int P_l(\|t\|, \|v^*\|) g_2[Q\{A(t)t - k_n^{-1}v^* - k_n^{-1}c\}] \bar{g}_1(m\|v^*\|) dv^* dt \end{aligned}$$

for some $P_l(t, v^*) \in \mathbb{P}_{l,(t,v^*)}$ and $P_l(\|t\|, \|v^*\|) \in \mathbb{P}_{l,(\|t\|, \|v^*\|)}$. The right hand side of the above inequality is similar to the integral when $k_\infty < \infty$ with $g_1(\cdot)$ and $g_2(\cdot)$ replaced by $g_2(\cdot)$ and $\bar{g}_1(\cdot)$ respectively. Therefore it is $O_p(1)$ by the same reasoning.

For $P_l(t, v) = 1$, by considering only the integral in a compact region, it is easy to see the target integral is larger than 0. Therefore the lemma holds. \square

Proof of Lemma 4. Let $P = A^T A$. By matrix algebra,

$$N\left(At; B_n v + \frac{1}{k_n}c, \frac{1}{k_n^2}I_d\right) K(v) = N\left\{t; P^{-1}A^T\left(B_n v + \frac{1}{k_n}c\right), \frac{1}{k_n^2}P^{-1}\right\} r(v; A, B_n, k_n, c),$$

where

$$r(v; A, B_n, k_n, c) = \frac{k_n^{d-p}}{(2\pi)^{(d-p)/2}} \exp\left\{-\frac{k_n^2}{2}\left(B_n v + \frac{c}{k_n}\right)^T (I - AP^{-1}A^T)\left(B_n v + \frac{c}{k_n}\right)\right\} K(v).$$

Then the target integral can be expanded as

$$\begin{aligned} \int t \frac{N(At; B_n v + \frac{1}{k_n}c, \frac{1}{k_n^2}I_d) K(v)}{\int N(At; B_n v + \frac{1}{k_n}c, \frac{1}{k_n^2}I_d) K(v) dt dv} dt dv &= \int P^{-1}A^T\left(\frac{1}{k_n}c + B_n v\right) \frac{r(v; A, B_n, k_n, c)}{\int r(v; A, B_n, k_n, c) dv} dv \\ &= \frac{1}{k_n} \left\{ (A^T A)^{-1} A^T c + R(A, B_n, k_n, c) \right\}, \end{aligned}$$

where

$$R(A, B_n, k_n, c) = (A^T A)^{-1} A^T B_n \int k_n v \frac{r(v; A, B_n, k_n, c)}{\int r(v; A, B_n, k_n, c) dv} dv.$$

The remainder term $R(A, B_n, k_n, c)$ depends on the mean of the probability density proportional to $r(v; A, B_n, k_n, c)$ in the directions of $(A^T A)^{-1} A^T B$. If B_n does not degenerate to 0 as $n \rightarrow \infty$, then in the directions orthogonal to those of $(I - A(A^T A)^{-1} A^T)^{1/2} B$, $r(v; A, B_n, k_n, c)$ is symmetric around 0; in the directions of $(I - A(A^T A)^{-1} A^T)^{1/2} B$, $r(v; A, B_n, k_n, c)$ is a product of a normal density whose mean is $O(1/k_n)$ and a rescaled $K(v)$, which is symmetric around 0, so its mean value is $O(1/k_n)$. Therefore when the spaces expanded by $(A^T A)^{-1} A^T B$ and $\{I - A(A^T A)^{-1} A^T\} B$ are orthogonal, $R(A, B_n, k_n, c) = 0$; when it is not the case, $R(A, B_n, k_n, c) = O(1)$.

If $B_n = o(1)$ as $n \rightarrow \infty$, which implies $k_n \rightarrow c \in (0, \infty)$, it is easy to see that $\int k_n v r(v; A, B_n, k_n, c) dv / \int r(v; A, B_n, k_n, c) dv$ is upper bounded as $n \rightarrow \infty$ and hence $R(A, B_n, k_n, c)$ is $o(1)$. \square

In the following lemmas, to deal with the case where $K(x) = \bar{K}(\|x\|_\Lambda)$ with Λ not the identity, we use the property that such a $K(x)$ can be bounded above by a function that depends only on $\|x\|$. We refer to this bound as $K(\cdot)$ rescaled to have identity covariance matrix.

337 *Proof of Lemma 5.* First consider $\tilde{\pi}_{B_\delta}(1)$. With the transformation $t = t(\theta)$,

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339
340
$$\tilde{\pi}_{B_\delta}(1) = a_{n,\varepsilon}^{-p} \int_{t(B_\delta)} \int_{\mathbb{R}^d} \pi(\theta_0 + a_{n,\varepsilon}^{-1}t) \tilde{f}_n(s_{\text{obs}} + \varepsilon_n v \mid \theta_0 + a_{n,\varepsilon}^{-1}t) K(v) dv dt. \quad (6)$$

341
342 We can obtain an expansion of $\tilde{\pi}_{B_\delta}(1)$ by expanding $\tilde{f}_n(s_{\text{obs}} + \varepsilon_n v \mid \theta_0 + a_{n,\varepsilon}^{-1}t)K(v)$ as fol-
343 lows. The expansion needs to be discussed separately for two cases, depending on whether the
344 limit of $a_n \varepsilon_n$ is finite or infinite.

345 When $a_n \varepsilon_n \rightarrow c_\varepsilon < \infty$, $a_{n,\varepsilon} = a_n$. We apply a Taylor expansion to $s(\theta_0 + a_n^{-1}t)$ and $A(\theta_0 +$
346 $a_n^{-1}t)^{-1/2}$ and have

347
348
$$\tilde{f}_n(s_{\text{obs}} + \varepsilon_n v \mid \theta_0 + a_n^{-1}t) = \frac{a_n^d}{|A(\theta_0 + a_n^{-1}t)|^{1/2}}$$

349
350
$$\times N \left(\left\{ A(\theta_0)^{-1/2} + a_n^{-1} r_A(t, \varepsilon_2) \right\} \left[A(\theta_0)^{1/2} T_{\text{obs}} + a_n \varepsilon_n v - \{Ds(\theta_0) + a_n^{-1} r_s(t, \varepsilon_1)\} t \right]; 0, I_d \right),$$

351
352 (7)

353
354 where $r_s(t, \varepsilon_1)$ is the $d \times p$ matrix whose i th row is $t^T H s_i \{\theta_0 + \varepsilon_1(t)\}$, $r_A(t, \varepsilon_2)$ is the $d \times d$
355 matrix $\sum_{k=1}^p \frac{d}{d\theta_k} A\{\theta_0 + \varepsilon_2(t)\}^{-1/2} t_k$, and $\varepsilon_1(t)$ and $\varepsilon_2(t)$ are from the remainder terms of
356 the Taylor expansions and satisfy $\|\varepsilon_1(t)\| \leq \delta$ and $\|\varepsilon_2(t)\| \leq \delta$. For a $d \times d$ matrix τ_2 , let
357 $g_n(t, v; \tau_1, \tau_2)$ be the function $g_n(t, v; \tau_1)$, defined in Section 4.3 of the main text, with $A(\theta_0)$
358 replaced by $\{A(\theta_0)^{-1/2} + \tau_2\}^{-2}$. Applying a Taylor expansion to the normal density in (7), we
359 have

360
361
$$\tilde{f}_n(s_{\text{obs}} + \varepsilon_n v \mid \theta_0 + a_n^{-1}t) K(v)$$

362
$$= \frac{a_n^d |A(\theta_0)|^{1/2}}{|A(\theta_0 + a_n^{-1}t)|^{1/2}} \left[g_n(t, v) + a_n^{-1} P_3(t, v) g_n\{t, v; e_{n1} r_s(t, \varepsilon_1), e_{n1} r_A(t, \varepsilon_2)\} \right], \quad (8)$$

363
364

365 where $P_3(t, v)$ is the function

366
367
$$\frac{1}{2|A(\theta_0)^{-1/2} + r_2(a_n^{-1}t)|}$$

368
$$\times \frac{d}{dx} \left\| \left\{ A(\theta_0)^{-1/2} + x r_A(t, \varepsilon_2) \right\} \left[A(\theta_0)^{1/2} T_{\text{obs}} + a_n \varepsilon_n v - \{Ds(\theta_0) + x r_s(t, \varepsilon_1)\} t \right] \right\|_{x=e_{n1}}^2,$$

369
370
371

372 and e_{n1} is from the remainder term of Taylor expansion and satisfies $|e_{n1}| \leq a_n^{-1}$. Since
373 $\|e_{n1}t\| \leq \delta$ and $r_s(t, \varepsilon_1)$ and $r_A(t, \varepsilon_2)$ belong $\mathbb{P}_{1,t}$, this $P_3(t, v)$ belongs to $\mathbb{P}_{3,(t,v)}$. Furthermore,
374 since $r_s(t, \varepsilon_1)$ and $r_A(t, \varepsilon_2)$ have no constant term, for any small σ , $e_{n1} r_s(t, \varepsilon_1)$ and $e_{n1} r_A(t, \varepsilon_2)$
375 can be bounded by σI_d and σI_p uniformly in n and t , if δ is small enough.

376
377 When $a_n \varepsilon_n \rightarrow \infty$, $a_{n,\varepsilon} = \varepsilon_n^{-1}$. Let $v^*(v) = A(\theta_0)^{1/2} T_{\text{obs}} + a_n \varepsilon_n v - a_n \varepsilon_n Ds(\theta_0)t$. Under
378 the transformation $v^* = v^*(v)$, the expansion of $\tilde{f}_n(s_{\text{obs}} + \varepsilon_n v \mid \theta_0 + \varepsilon_n t)$ obtained by applying
379 a Taylor expansion to $s(\theta_0 + \varepsilon_n t)$ and $A(\theta_0 + \varepsilon_n t)^{-1/2}$ is

380
381
$$\tilde{f}_n(s_{\text{obs}} + \varepsilon_n v \mid \theta_0 + \varepsilon_n t)$$

382
$$= \frac{a_n^d}{|A(\theta_0 + a_n^{-1}t)|^{1/2}} N \left[\left\{ A(\theta_0)^{-1/2} + a_n \varepsilon_n^2 \frac{r_A(t, \varepsilon_4)}{a_n \varepsilon_n} \right\} \left\{ v^* - a_n \varepsilon_n^2 r_s(t, \varepsilon_3) t \right\}; 0, I_d \right],$$

383
384

where $\epsilon_3(t)$ and $\epsilon_4(t)$ are from the remainder terms of the Taylor expansion and satisfy $\|\epsilon_3(t)\| \leq \delta$ and $\|\epsilon_4(t)\| \leq \delta$. Let $g_n^*(t, v^*; \tau_1, \tau_2)$ be the function

$$g_n^*(t, v^*; \tau_1, \tau_2) = N \left[v^*; a_n \epsilon_n \tau_1 t, \{A(\theta_0)^{-1/2} + \tau_2\}^{-2} \right] K \left\{ Ds(\theta_0)t + \frac{1}{a_n \epsilon_n} v^* - \frac{1}{a_n \epsilon_n} A(\theta_0)^{1/2} T_{\text{obs}} \right\},$$

so that $(a_n \epsilon_n)^d g_n^*(t, v^*; \tau_1, \tau_2)$ is $g_n(t, v; \tau_1, \tau_2)$ with transformed variable $v^* = v^*(v)$, and $g_n^*(t, v^*) = g_n^*(t, v^*; 0, 0)$. Denote a $k_1 \times k_2$ matrix with element being $P_l(t)$ by $P_l^{(k_1 \times k_2)}(t)$. Then by applying a Taylor expansion to the normal density in the expansion above,

$$\begin{aligned} & \tilde{f}_n(s_{\text{obs}} + \epsilon_n v \mid \theta_0 + \epsilon_n t) K(v) \\ &= \frac{\epsilon_n^{-d} |A(\theta_0)|^{1/2}}{|A(\theta_0 + \epsilon_n t)|^{1/2}} \left[g_n^*(t, v^*) + a_n \epsilon_n^2 \left\{ P_2^{(d \times 1)}(t) v^* + \frac{1}{a_n \epsilon_n} v^{*T} P_1^{(d \times d)}(t) v^* \right\} g_n^*(t, v^*) \right. \\ & \quad \left. + (a_n \epsilon_n^2)^2 P_4(t, v^*) g_n^*\{t, v^*; e_{n2} r_s(t, \epsilon_3), e_{n2} r_A(t, \epsilon_4)\} \right] (a_n \epsilon_n)^d, \end{aligned} \quad (9)$$

where $P_2^{(d \times 1)}(t)$ is the function $t^T r_s(t, \epsilon_3)^T A(\theta_0)^{-1/2} / 2$, $P_1^{(d \times d)}(t)$ is the function $-A(\theta_0)^{-1/2} r_A(t, \epsilon_4)$, $e_{n2} = e'_{n2} / (a_n \epsilon_n)$, e'_{n2} is from the remainder term of the Taylor expansion and satisfies $|e'_{n2}| \leq a_n \epsilon_n^2$, and $P_4(t, v^*)$ is a linear combination of $\{d\rho(w)/dw\}^2$ and $d^2\rho(w)/dw^2$ at $w = e'_{n2}$ with $\rho(w)$ being the function

$$\left\| \left\{ A(\theta_0)^{-1/2} + w \frac{r_A(t, \epsilon_4)}{a_n \epsilon_n} \right\} \left\{ v^* - w r_s(t, \epsilon_3) t \right\} \right\|^2.$$

Obviously elements of $P_2^{(d \times 1)}(t)$ and $P_1^{(d \times d)}(t)$ belong to $\mathbb{P}_{2,t}$ and $\mathbb{P}_{1,t}$ respectively. Since $\|e_{n2} t\| \leq \delta$, the function $P_4(t, v^*)$ belongs to $\mathbb{P}_{4,(t,v^*)}$ and, similar to before, $e_{n2} r_s(t, \epsilon_3)$ and $e_{n2} r_A(t, \epsilon_4)$ can be bounded by σI_d and σI_p uniformly in n and t for any small σ , if δ is small enough.

For $\pi(\theta_0 + a_{n,\varepsilon}^{-1} t)$ in the integral of $\tilde{\pi}_{B_\delta}(1)$ in (6), a Taylor expansion gives that

$$\frac{\pi(\theta_0 + a_{n,\varepsilon}^{-1} t)}{|A(\theta_0 + a_{n,\varepsilon}^{-1} t)|^{1/2}} = \frac{\pi(\theta_0)}{|A(\theta_0)|^{1/2}} + a_{n,\varepsilon}^{-1} D_\theta \frac{\pi\{\theta_0 + \epsilon_5(t)\}}{|A\{\theta_0 + \epsilon_5(t)\}|^{1/2}} t, \quad |\epsilon_5(t)| \leq \delta. \quad (10)$$

As mentioned before, δ can be selected such that $Ds(\theta_0) + e_{n1} r_s(t, \epsilon_1)$ and $Ds(\theta_0) + e_{n2} r_s(t, \epsilon_3)$ are lower bounded by $m_1 I_p$ and $A(\theta_0)^{-1/2} + e_{n1} r_A(t, \epsilon_2)$ and $A(\theta_0)^{-1/2} + e_{n2} r_A(t, \epsilon_4)$ are lowered bounded by $m_2 I_d$ for some positive constant m_1 and m_2 . We choose δ satisfying these and, since $\|a_{n,\varepsilon}^{-1} t\| \leq \delta$, this means $\pi(\theta_0 + a_{n,\varepsilon}^{-1} t) / |A(\theta_0 + a_{n,\varepsilon}^{-1} t)|^{1/2}$ is bounded uniformly in t and n .

By plugging (8)–(10) into (6), it can be seen that the leading term of $\tilde{\pi}_{B_\delta}(1)$ is $a_{n,\varepsilon}^{d-p} \pi(\theta_0) \int_{t(B_\delta) \times \mathbb{R}^d} g_n(t, v) dt dv$. The remainder terms are given in the following,

$$\begin{aligned}
& a_{n,\varepsilon}^{p-d} \tilde{\pi}_{B_\delta}(1) - \pi(\theta_0) \int_{t(B_\delta) \times \mathbb{R}^d} g_n(t, v) dt dv \\
&= a_{n,\varepsilon}^{-1} \int_{t(B_\delta) \times \mathbb{R}^d} |A(\theta_0)|^{1/2} D \frac{\pi(\theta_0 + \varepsilon_5)}{|A(\theta_0 + \varepsilon_5)|^{1/2}} t g_n(t, v) dv dt \\
&+ a_n^{-1} \int_{t(B_\delta) \times \mathbb{R}^d} P_3(t, v) g_n\{t, v; e_{n1} r_s(t, \varepsilon_1), e_{n1} r_A(t, \varepsilon_2)\} dv dt \mathbb{1}_{\{\lim a_n \varepsilon_n < \infty\}} \\
&+ a_n \varepsilon_n^2 \int_{t(B_\delta)} P_2^{(d \times 1)}(t) \int_{\mathbb{R}^d} v^* g_n^*(t, v^*) dv^* dt \mathbb{1}_{\{\lim a_n \varepsilon_n = \infty\}} \\
&+ \varepsilon_n \int_{t(B_\delta) \times \mathbb{R}^d} v^{*T} P_1^{(d \times d)}(t) v^* g_n^*(t, v^*) dv^* dt \mathbb{1}_{\{\lim a_n \varepsilon_n = \infty\}} \\
&+ a_n^2 \varepsilon_n^4 \int_{t(B_\delta) \times \mathbb{R}^d} P_4(t, v^*) g_n^*\{t, v^*; e_{n2} r_s(t, \varepsilon_3), e_{n2} r_A(t, \varepsilon_4)\} dv^* dt \mathbb{1}_{\{\lim a_n \varepsilon_n = \infty\}}, \quad (11)
\end{aligned}$$

where $P_3(t, v)$, $P_2^{(d \times 1)}(t)$, $P_1^{(d \times d)}(t)$ and $P_4(t, v^*)$ are products of $\pi(\theta_0 + a_{n,\varepsilon}^{-1}t)/|A(\theta_0 + a_{n,\varepsilon}^{-1}t)|^{1/2}$ and corresponding terms in expansions (8) and (9). In the above, there are five remainder terms. For the integrals in the first two terms, it is easy to write them in the form of the first integral in Lemma 7 and conditions therein are satisfied, where $g_1(\cdot)$ is the standard normal density and $g_2(\cdot)$ is $K(v)$ rescaled to have identity covariance. Then the first two terms are $O_p(a_{n,\varepsilon}^{-1})$ and $O_p(a_n^{-1})$. The integral in the fourth term can also be written in this form where $g_1(\cdot)$ is the rescaled $K(v)$ and $g_2(\cdot)$ is the standard normal density. The integral in the fifth term needs to use the transformation $v^{**} = v^* - a_n \varepsilon_n e_{n2} r_s(t, \varepsilon_3)t$, after which it can be written in a similar form, as $P_5\{t, v^{**} + a_n \varepsilon_n e_{n2} r_s(t, \varepsilon_3)t\} \in \mathbb{P}_{5,(t,v^{**})}$ by the expression of $P_4(t, v^*)$ in (9). Thus the fourth and fifth term are $O_p(\varepsilon_n)$ and $O_p(a_n^2 \varepsilon_n^4)$.

The third term is somewhat different as the center of $g_n^*(t, v^*)$ in the direction of v^* degenerates to zero as $n \rightarrow \infty$. Let ψ_k be the d -dimension unit vector with 1 at the k th coordinate. Then

$$\begin{aligned}
\int_{-\infty}^{\infty} v_k^* g_n^*(t, v^*) dv_k^* &= \int_0^{\infty} v_k^* \{g_n^*(t, v^*) - g_n^*(t, v^* - 2v_k^* \psi_k)\} dv_k^* \\
&= \int_0^{\infty} v_k^* N\{v^*; 0, A(\theta_0)\} [K\{v(v^*)\} - K\{v(v^* - 2v_k^* \psi_k)\}] dv_k^*,
\end{aligned}$$

which by a Taylor expansion is bounded by $(a_n \varepsilon_n)^{-1} c$ for some constant c . Hence the third term is $O_p(\varepsilon_n)$. Combining the orders of all remainder terms, the expansion of $\tilde{\pi}_{B_\delta}(1)$ in the lemma holds.

For any $P_2(t) \in \mathbb{P}_{2,t}$, $\tilde{\pi}_{B_\delta}\{P_2(t)\}$ can be expanded similarly to $\tilde{\pi}_{B_\delta}(1)$ in (11), simply by multiplying $P_2(t)$ into every integral in (11). This gives that

$$\tilde{\pi}_{B_\delta}\{P_2(t)\} = a_{n,\varepsilon}^{d-p} \left\{ \pi(\theta_0) \int_{t(B_\delta) \times \mathbb{R}^d} P_2(t) g_n(t, v) dt dv + O_p(a_{n,\varepsilon}^{-1}) + O_p(a_n^2 \varepsilon_n^4) \right\}.$$

Then since $\int_{t(B_\delta) \times \mathbb{R}^d} g_n(t, v) dt dv = \Theta_p(1)$ by the second result of Lemma 7, $\tilde{\pi}_{B_\delta}\{P_2(t)\}/\tilde{\pi}_{B_\delta}(1) = O_p(1)$ and (2) holds by taking $P_2(t) = t$. \square

481 *Proof of Lemma 6.* Let $r_n(s | \theta)$ be the scaled remainder $\alpha_n\{f_n(s | \theta) - \tilde{f}_n(s | \theta)\}$. The error
 482 of using $\tilde{\pi}_{B_\delta}\{P_l(t)\}$ to approximate $\pi_{B_\delta}\{P_l(t)\}$ is
 483

$$484 \pi_{B_\delta}\{P_l(t)\} - \tilde{\pi}_{B_\delta}\{P_l(t)\} = \alpha_n^{-1} \int_{B_\delta} \int P_l\{t(\theta)\} \pi(\theta) r_n(s_{\text{obs}} + \varepsilon_n v | \theta) K(v) dv d\theta.$$

486 If this approximation error satisfies

$$487 \frac{\pi_{B_\delta}\{P_l(t)\} - \tilde{\pi}_{B_\delta}\{P_l(t)\}}{\tilde{\pi}_{B_\delta}(1)} = O_p(\alpha_n^{-1}), \quad (12)$$

489 then, since $a_{n,\varepsilon}^{p-d} \tilde{\pi}_{B_\delta}(1) = \Theta_p(1)$ by Lemma 5,

$$492 \pi_{B_\delta}(1) = \tilde{\pi}_{B_\delta}(1)\{1 + O_p(\alpha_n^{-1})\}, \quad \frac{\pi_{B_\delta}\{P_l(t)\}}{\pi_{B_\delta}(1)} = \frac{\tilde{\pi}_{B_\delta}\{P_l(t)\}}{\tilde{\pi}_{B_\delta}(1)} + O_p(\alpha_n^{-1}). \quad (13)$$

494 By plugging (12) into (1),

$$495 \frac{\pi_{B_\delta}(h)}{\pi_{B_\delta}(1)} = h(\theta_0) + a_{n,\varepsilon}^{-1} Dh(\theta_0)^T \left\{ \frac{\tilde{\pi}_{B_\delta}(t)}{\tilde{\pi}_{B_\delta}(1)} + O_p(\alpha_n^{-1}) \right\} + \frac{1}{2} a_{n,\varepsilon}^{-2} \left[\frac{\tilde{\pi}_{B_\delta}\{t^T H h(\theta_t) t\}}{\tilde{\pi}_{B_\delta}(1)} + O_p(\alpha_n^{-1}) \right]. \quad (14)$$

499 Verification of (12) is given by the following argument. With the transformation $t = t(\theta)$ we
 500 have

$$501 \pi_{B_\delta}\{P_l(t)\} - \tilde{\pi}_{B_\delta}\{P_l(t)\} = \alpha_n^{-1} a_{n,\varepsilon}^{-p} \int_{t(B_\delta)} \int P_l(t) \pi(\theta_0 + a_{n,\varepsilon}^{-1} t) r_n(s_{\text{obs}} + \varepsilon_n v | \theta_0 + a_{n,\varepsilon}^{-1} t) K(v) dv dt.$$

504 Let $r_{W_n}(w | \theta) = \alpha_n\{f_{W_n}(w | \theta) - \tilde{f}_{W_n}(w | \theta)\}$, and we have

$$505 r_n(s | \theta) = a_n^d |A(\theta)|^{-1/2} r_{W_n}[a_n A(\theta)^{-1/2}\{s - s(\theta)\} | \theta].$$

508 For the value of δ , we choose the smaller value of the one from Lemma 5 and the one such
 509 that $Ds(\theta)$ is lower bounded and $A(\theta)^{-1/2}$ is upper bounded by MI_d in B_δ for some $M > 0$.
 510 Since $r_{W_n}(w | \theta)$ is upper bounded by $r_{\max}(w)$ according to Condition 5, by applying a Taylor
 511 expansion to $s(\theta_0 + a_{n,\varepsilon}^{-1} t)$ we have

$$512 |\pi_{B_\delta}\{P_l(t)\} - \tilde{\pi}_{B_\delta}\{P_l(t)\}| \leq \alpha_n^{-1} a_{n,\varepsilon}^{d-p} \sup_{\theta \in B_\delta} |\pi(\theta) A(\theta)^{-1/2}| \int_{t(B_\delta)} \int |P_l(t)| (a_n a_{n,\varepsilon}^{-1})^d$$

$$513 r_{\max} \left[a_n a_{n,\varepsilon}^{-1} M \left\{ Ds(\theta_0 + \varepsilon_t) t - a_{n,\varepsilon} \varepsilon_n v - \frac{1}{a_n a_{n,\varepsilon}^{-1}} A(\theta_0)^{1/2} T_{\text{obs}} \right\} \right] K(v) dv dt,$$

514 where ε_t is from the remainder term of the Taylor expansion and satisfies $|\varepsilon_t| \leq \delta$. Since
 515 $\tilde{\pi}_{B_\delta}(1) = \Theta_p(a_{n,\varepsilon}^{d-p})$ by Lemma 5, it is sufficient to show that the above integral is $O_p(1)$. This is
 516 immediate by noting that when either $\lim a_n \varepsilon_n \rightarrow \infty$ or $\lim a_n \varepsilon_n \rightarrow c_\varepsilon < \infty$, the above integral
 517 can be written in the form of the first integral in Lemma 7 and conditions therein are satisfied,
 518 where $g_1(\cdot)$ and $g_2(\cdot)$ are $r_{\max}(\cdot)$ and $K(\cdot)$ rescaled to have identity covariance matrix. \square

524 2. PROOF OF RESULTS FROM SECTION 4

525 2.1. Proof of Proposition 2

526 The proof of Proposition 2 follows the standard asymptotic argument of importance sampling.
 527 In the following we use the convention that for a vector x , the matrix xx^T is denoted by x^2 .
 528

529 *Proof of Proposition 2.* Algorithm 1 generates independent, identically distributed triples,
 530 $(\phi_i, \theta_i, s_n^{(i)})$, where $(\theta_i, s_n^{(i)})$ is generated from $g_n(\theta)f(s_n | \theta)$, and, conditional on $s_n = s_n^{(i)}$,
 531 ϕ_i is generated from a Bernoulli distribution with probability $K_{\varepsilon_n}(s_n - s_{\text{obs}})$.

532 Now \hat{h} can be expressed as a ratio of sample means of functions of these independent, indenti-
 533 cally distributed random variables. Thus we can use the standard delta method (Lehmann, 2004)
 534 for ratio statistics to show that the central limit theorem holds. Further we obtain that the limiting
 535 distribution has mean

$$536 \frac{E\{h(\theta_1)w_1\phi_1\}}{E\{w_1\phi_1\}} = \frac{E\{h(\theta_1)w_1K_{\varepsilon_n}(s_n^{(1)} - s_{\text{obs}})\}}{E\{w_1K_{\varepsilon_n}(s_n^{(1)} - s_{\text{obs}})\}} = \frac{\int h(\theta)\pi(\theta)f_n(s_n | \theta)K_{\varepsilon_n}(s_n - s_{\text{obs}}) ds_n d\theta}{\int \pi(\theta)f_n(s_n | \theta)K_{\varepsilon_n}(s_n - s_{\text{obs}}) ds_n d\theta},$$

537 which is equal to h_{ABC} . Its variance is

$$538 \frac{1}{E^2(w_1\phi_1)} \text{var}\{h(\theta_1)w_1\phi_1\} + \frac{E^2\{h(\theta_1)w_1\phi_1\}}{E^4(w_1\phi_1)} \text{var}(w_1\phi_1) - 2 \frac{E\{h(\theta_1)w_1\phi_1\}}{E^3(w_1\phi_1)} \text{cov}\{h(\theta_1)w_1\phi_1, w_1\phi_1\}^T$$

$$539 = p_{\text{acc},\pi}^{-2} [E\{h(\theta_1)^2 w_1^2 \phi_1\} - h_{\text{ABC}}^2 p_{\text{acc},\pi}^2 + h_{\text{ABC}}^2 \{E(w_1^2 \phi_1) - p_{\text{acc},\pi}^2\}$$

$$540 - 2h_{\text{ABC}} \{E\{h(\theta_1)w_1^2 \phi_1\} - h_{\text{ABC}} p_{\text{acc},\pi}^2\}^T]$$

$$541 = p_{\text{acc},\pi}^{-2} E[\{h(\theta_1)^2 - 2h_{\text{ABC}}h(\theta_1)^T + h_{\text{ABC}}^2\} w_1^2 K_{\varepsilon_n}(s_n^{(1)} - s_{\text{obs}})]$$

$$542 = p_{\text{acc},\pi}^{-1} E_{\pi_{\text{ABC}}} \left\{ (h(\theta) - h_{\text{ABC}})^2 \frac{\pi(\theta)}{q_n(\theta)} \right\}.$$

543 In the above expression we used $p_{\text{acc},\pi} = E(w_1\phi_1)$. It is easy to verify that

$$544 \Sigma_{\text{ABC},n} = p_{\text{acc},\pi}^{-1} E_{\pi_{\text{ABC}}} \left\{ (h(\theta) - h_{\text{ABC}})^2 \frac{\pi(\theta)}{q_n(\theta)} \right\}, \quad (15)$$

545 as required. \square

546 2.2. Proof of Theorem 2

547 For simplicity, a consider one-dimensional function $h(\theta)$. For multi-dimensional functions,
 548 the extension is trivial by considering each element of $\Sigma_{\text{IS},n}$ separately. Denote $\{h(\theta) - h_{\text{ABC}}\}^2$
 549 by $G_n(\theta)$. In Theorem 2(i), $\Sigma_{\text{IS},n}$ is just the ABC posterior variance of $h(\theta)$, and the derivation
 550 of its order is similar to that of h_{ABC} in Section 1 of this supplementary material. The result is
 551 stated in the following lemma.

552 LEMMA 8. Assume the conditions of Theorem 1. Then $\text{var}_{\pi_{\text{ABC}}}\{h(\theta)\} = O_p(a_{n,\varepsilon}^{-2})$.

553 *Proof.* Using the notation of Section 1, $\text{var}_{\pi_{\text{ABC}}}[h(\theta)] = \pi(G_n)/\pi(1)$. It follows immediately
 554 from Lemma 3 that

$$555 \text{var}_{\pi_{\text{ABC}}}\{h(\theta)\} = \frac{\pi_{B_\delta}(G_n)}{\pi_{B_\delta}(1)} \{1 + o_p(1)\}.$$

556 Applying a first order Taylor expansion of $h(\theta)$ around $\theta = \theta_0$ gives

$$557 \frac{\pi_{B_\delta}(G_n)}{\pi_{B_\delta}(1)} = G_n(\theta_0) + 2a_{n,\varepsilon}^{-1} \{h(\theta_0) - h_{\text{ABC}}\} \frac{\pi_{B_\delta}\{Dh(\theta_t)^T t\}}{\pi_{B_\delta}(1)} + a_{n,\varepsilon}^{-2} \frac{\pi_{B_\delta}\{t^T Dh(\theta_t) Dh(\theta_t)^T t\}}{\pi_{B_\delta}(1)}, \quad (16)$$

558 where θ_t is from the remainder term and belongs to B_δ . In the above decomposition, $G_n(\theta_0)$
 559 and $a_{n,\varepsilon}^{-1} \{h(\theta_0) - h_{\text{ABC}}\}$ are $O_p(a_{n,\varepsilon}^{-2})$ by Theorem 1. Since $Dh(\theta_t)^T t$ and $t^T Dh(\theta_t) Dh(\theta_t)^T t$
 560 belong to $\mathbb{P}_{2,t}$, the two ratios in the above are $O_p(1)$ by Lemma 5 and Lemma 6. \square

The following lemma states that moments of $K(v)^\gamma$ exist for any positive constant γ .

LEMMA 9. Assume Condition 2. For any constant $\gamma \in (0, \infty)$ and coordinates $(v_{i_1}, \dots, v_{i_l})$ of v with $l \leq p + 6$, $\int \prod_{k=1}^l v_{i_k} K(v)^\gamma dv < \infty$.

Proof. By Condition 2 (iv), for some positive constant M there exists $x_0 \in (0, \infty)$ such that when $\|v\| > x_0$, $K(v) < M e^{-c_1 \|v\|^{\alpha_1}}$. Then consider the integration in two regions $\{v : \|v\| \leq x_0\}$ and $\{v : \|v\| > x_0\}$ separately. In the first region, since $K(v) \leq 1$, we have

$$\int_{\|v\| \leq x_0} \prod_{k=1}^l v_{i_k} K(v)^\gamma dv \leq x_0^l V_{x_0},$$

where V_{x_0} is the volume of the d -dimension sphere with radius x_0 , and is finite. In the second region,

$$\int_{\|v\| > x_0} \prod_{k=1}^l v_{i_k} K(v)^\gamma dv \leq M \int_{\|v\| > x_0} \|v\|^l e^{-c_1 \gamma \|v\|^{\alpha_1}} dv.$$

The right hand side of this is proportional to $\exp\{-c_1 \gamma x_0^{\alpha_1/(l+d)}\}$ by integrating in spherical coordinates. \square

Proof of Theorem 2. For (i), since $p_{\text{acc}, \pi} = \varepsilon_n^d \pi(1)$ and $\pi(1) = \Theta_p(a_{n, \varepsilon}^{d-p})$ by Lemmas 3, 5 and 6, then $p_{\text{acc}, \pi} = \Theta_p(\varepsilon_n^d a_{n, \varepsilon}^{d-p})$. Together with Lemma 8, (i) holds.

For (ii), if we can show that $p_{\text{acc}, q} = \Theta_p(\varepsilon_n^d a_{n, \varepsilon}^d)$, then the order of $\Sigma_{\text{IS}, n}$ is obvious from (15) and the definition of $\Sigma_{\text{ABC}, n}$. Similar to the expansion of $\pi(1)$ from Lemma 3 and (13),

$$\begin{aligned} p_{\text{acc}, q} &= \varepsilon_n^d \int \pi_{\text{ABC}}(\theta \mid s_{\text{obs}}, \varepsilon_n) f_{\text{ABC}}(s_{\text{obs}} \mid \theta) d\theta \\ &= \varepsilon_n^d \left\{ \frac{\int_{B_\delta} \pi(\theta) \tilde{f}_{\text{ABC}}(s_{\text{obs}} \mid \theta)^2 d\theta}{\tilde{\pi}_{B_\delta}(1)} + O_p(\alpha_n^{-1}) \right\} \{1 + o_p(1)\}. \end{aligned}$$

The integral in the above differs from $\tilde{\pi}_{B_\delta}(1)$ by the square power of $\tilde{f}_{\text{ABC}}(s_{\text{obs}} \mid \theta)$ in the integrand. We will show that this integral has order $\Theta_p(a_{n, \varepsilon}^{2d-p})$, from which $p_{\text{acc}, q} = \Theta_p(\varepsilon_n^d a_{n, \varepsilon}^d)$ trivially holds. Let $g_n^{**}(t, v; \tau_1, \tau_2)$ be the function

$$g_n^{**}(t, v; \tau_1, \tau_2) = N[v; 0, \{A(\theta_0)^{-1/2} + \tau_2\}^{-2}] K \left[\{Ds(\theta_0) + \tau_1\}t + \frac{1}{a_n \varepsilon_n} v^* - \frac{1}{a_n \varepsilon_n} A(\theta_0)^{1/2} T_{\text{obs}} \right],$$

and $g_n^{**}(t, v; \tau_1, \tau_2) = g_n^*(t, v + a_n \varepsilon_n \tau_1 t; \tau_1, \tau_2)$. Here expansions (8) and (9) of $\tilde{f}_n(s_{\text{obs}} + \varepsilon_n v \mid \theta_0 + a_n^{-1} t) K(v)$ are to be used in the form of

$$\frac{a_{n, \varepsilon}^d |A(\theta_0)|^{1/2}}{|A(\theta_0 + a_n^{-1} t)|^{1/2}} \begin{cases} \{g_n(t, v) + a_n^{-1} P_3(t, v) g_{n,r}(t, v)\}, & \lim_{n \rightarrow \infty} a_n \varepsilon_n < \infty, \\ \{g_n^*(t, v^*) + a_n \varepsilon_n^2 P_3(t, v^*) g_n^*(t, v^*) \\ + (a_n \varepsilon_n^2)^2 P_4(t, v^{**}) g_{n,r}^*(t, v^{**})\} (a_n \varepsilon_n)^d, & \lim_{n \rightarrow \infty} a_n \varepsilon_n = \infty, \end{cases} \quad (17)$$

where $P_3(t, v^*) \in \mathbb{P}_{3, (t, v^*)}$, $g_{n,r}(t, v) = g_n\{t, v; e_{n1} r_s(t, \varepsilon_1), e_{n1} r_A(t, \varepsilon_2)\}$, $g_{n,r}^*(t, v^{**})$ is $g_n^*\{t, v^{**}; e_{n2} r_s(t, \varepsilon_3), e_{n2} r_A(t, \varepsilon_4)\}$ and $P_4(t, v^{**})$ is $P_4(t, v^*)$ with the transformation $v^{**} = v^* - a_n \varepsilon_n e_{n2} r_s(t, \varepsilon_3) t$, and the expansion of $\pi(\theta)/|A(\theta)|$ similar to (10) is to be used.

By the expression of $P_4(t, v^*)$ in (9), it can be seen that $P_4(t, v^{**}) \in \mathbb{P}_{4,(t,v^{**})}$. Basic inequalities $(a + \varepsilon b)^2 \leq \varepsilon a^2 + (\varepsilon + \varepsilon^2)b^2$ and $(a + \varepsilon b + \varepsilon^2 c)^2 \leq (\varepsilon + \varepsilon^2)a^2 + (\varepsilon + \varepsilon^2 + \varepsilon^3)b^2 + (\varepsilon^2 + \varepsilon^3 + \varepsilon^4)c^2$ for any real constants a, b, c and ε , from the fact that $2ab \leq a^2 + b^2$, are also to be used. Then by the above expansions and inequalities, an expansion of the target integral similar to (11) can be obtained, with the leading term $a_{n,\varepsilon}^{2d-p}\pi(\theta_0) \int_{t(B_\delta)} \left\{ \int g_n(t, v) dv \right\}^2 dt$ and remainder term with the following upper bound

$$\begin{aligned}
& \left| a_{n,\varepsilon}^{p-2d} \int_{B_\delta} \pi(\theta) \tilde{f}_{\text{ABC}}(s_{\text{obs}} | \theta)^2 d\theta - \pi(\theta_0) \int_{t(B_\delta)} \left\{ \int g_n(t, v) dv \right\}^2 dt \right| \\
& \leq a_{n,\varepsilon}^{-1} \int_{t(B_\delta)} |A(\theta_0)| D_\theta \frac{\pi(\theta_0 + \varepsilon_6)}{|A(\theta_0 + \varepsilon_6)|} t \left\{ \int g_n(t, v) dv \right\}^2 dt \\
& \quad + M \int_{t(B_\delta)} \left[a_n^{-1} \left\{ \int g_n(t, v) dv \right\}^2 + (a_n^{-1} + a_n^{-2}) \left\{ \int P_3(t, v) g_{n,r}(t, v) dv \right\}^2 \right] dt \mathbb{1}_{\{\lim a_n \varepsilon_n < \infty\}} \\
& \quad + M \int_{t(B_\delta)} \left[\{a_n \varepsilon_n^2 + (a_n \varepsilon_n^2)^2\} \left\{ \int g_n^*(t, v^*) dv^* \right\}^2 \right. \\
& \quad \quad + \{a_n \varepsilon_n^2 + (a_n \varepsilon_n^2)^2 + (a_n \varepsilon_n^2)^3\} \left\{ \int P_3(t, v^*) g_n^*(t, v^*) dv^* \right\}^2 \\
& \quad \quad \left. + \{(a_n \varepsilon_n^2)^2 + (a_n \varepsilon_n^2)^3 + (a_n \varepsilon_n^2)^4\} \left\{ \int P_4(t, v^{**}) g_{n,r}^*(t, v^{**}) dv^{**} \right\}^2 \right] dt \mathbb{1}_{\{\lim a_n \varepsilon_n = \infty\}},
\end{aligned}$$

where M is the upper bound of $\pi(\theta)|A(\theta_0)|/|A(\theta)|$ for $\theta \in B_\delta$ with δ chosen so that M exists. Then if we can show that for any $P_4(t, v) \in \mathbb{P}_{5,(t,v)}$, $d \times p$ matrix function $r_{n1}(t)$ and $d \times d$ matrix function $r_{n2}(t)$ which can be bounded by σI_d and σI_p uniformly in n and t for any small δ if δ is small enough, (a) $\int_{t(B_\delta)} \left\{ \int_{\mathbb{R}^d} g_n(t, v) dv \right\}^2 dt$ is $\Theta_p(1)$; (b) $\int_{t(B_\delta)} \left[\int_{\mathbb{R}^d} P_4(t, v) g_n \{t, v; r_{n1}(t), r_{n2}(t)\} dv \right]^2 dt$ is $O_p(1)$ when $\lim_{n \rightarrow \infty} a_n \varepsilon_n < \infty$; (c) $\int_{t(B_\delta)} \left[\int_{\mathbb{R}^d} P_4(t, v) g_n^{**} \{t, v; r_{n1}(t), r_{n2}(t)\} dv \right]^2 dt$ is $O_p(1)$ when $\lim_{n \rightarrow \infty} a_n \varepsilon_n = \infty$, the lemma would hold.

Here δ is selected such that $Ds(\theta_0) + r_{n1}(t)$ is bounded bounded by $m_1 I_p$ and $m_2 I_d \leq A(\theta_0)^{-1/2} + r_{n2}(t) \leq M_2 I_d$, for some positive constants m_1, m_2 and M_2 , uniformly in n and t . For the purpose of bounding integrals, we can assume that $A(\theta_0) = I_d$ and $r_{n2}(t) = 0$ without loss of generality by the following inequality when $\lim_{n \rightarrow \infty} a_n \varepsilon_n < \infty$,

$$g_n \{t, v; r_{n1}(t), r_{n2}(t)\} \leq \frac{M_2^d}{(2\pi)^{d/2}} \exp \left[-\frac{m_2^2}{2} \|a_n \varepsilon_n v + A(\theta_0)^{1/2} T_{\text{obs}} - \{Ds(\theta_0) + r_{n1}(t)\} t\|^2 \right] K(v),$$

and a similar one for $g_n^{**} \{t, v; r_{n1}(t), r_{n2}(t)\}$.

Consider any $P_4(t, v) \in \mathbb{P}_{4,(t,v)}$. When $\lim_{n \rightarrow \infty} a_n \varepsilon_n < \infty$, let $E_1 = \{v : \|a_n \varepsilon_n v\|^2 \leq \beta_1 \| \{Ds(\theta_0) + r_{n1}(t)\} t - A(\theta_0)^{1/2} T_{\text{obs}} \|^2\}$ for some $\beta_1 \in (0, 1)$. Then for any $\beta_2 \in (0, 1)$ we

673 have

$$\begin{aligned}
674 & \int_{\mathbb{R}^d} P_4(t, v) g_n \{t, v; r_{n1}(t), r_{n2}(t)\} dv \\
675 & \leq \left(\int_{E_1} + \int_{E_1^c} \right) P_4(t, v) \frac{M_2^d}{(2\pi)^{d/2}} \exp \left[-\frac{m_2^2}{2} \|a_n \varepsilon_n v - \{Ds(\theta_0) + r_{n1}(t)\}t + A(\theta_0)^{1/2} T_{\text{obs}}\|^2 \right] K(v) dv \\
676 & \\
677 & \leq P_4(t) \left(\exp \left[-\frac{m_2^2(1-\beta_1)}{2} \|\{Ds(\theta_0) + r_{n1}(t)\}t - A(\theta_0)^{1/2} T_{\text{obs}}\|^2 \right] \right. \\
678 & \\
679 & \left. + \bar{K}^{\beta_2} \left[\frac{\lambda_{\min}^2(\Lambda) \beta_1}{a_n^2 \varepsilon_n^2} \|\{Ds(\theta_0) + r_{n1}(t)\}t - A(\theta_0)^{1/2} T_{\text{obs}}\|^2 \right] \right), \tag{18}
\end{aligned}$$

684 where $P_4(t) \in \mathbb{P}_{4,t}$ and the above inequality uses Lemma 9. Then using $(a+b)^2 \leq 2(a^2+b^2)$,

$$\begin{aligned}
685 & \int_{t(B_\delta)} \left[\int_{\mathbb{R}^d} P_4(t, v) g_n \{t, v; r_{n1}(t), r_{n2}(t)\} dv \right]^2 dt \\
686 & \\
687 & \leq \int_{t(B_\delta)} P_8(t) \exp \left[-m_2^2(1-\beta_1) \|\{Ds(\theta_0) + r_{n1}(t)\}t - A(\theta_0)^{1/2} T_{\text{obs}}\|^2 \right] dt \\
688 & \\
689 & + \int_{t(B_\delta)} P_8(t) \bar{K}^{2\beta_2} \left[\frac{\lambda_{\min}^2(\Lambda) \beta_1}{a_n^2 \varepsilon_n^2} \|\{Ds(\theta_0) + r_{n1}(t)\}t - A(\theta_0)^{1/2} T_{\text{obs}}\|^2 \right] dt,
\end{aligned}$$

694 where $P_8(t) \in \mathbb{P}_{8,t}$.

695 When $a_n \varepsilon_n \rightarrow \infty$, let $E_2 = \{v : \|(a_n \varepsilon_n)^{-1} v\|^2 \leq \beta_1 \|\{Ds(\theta_0) + r_{n1}(t)\}t -$
696 $(a_n \varepsilon_n)^{-1} A(\theta_0)^{1/2} T_{\text{obs}}\|^2\}$ for some $\beta_1 \in (0, 1)$. Then for any $\beta_2 \in (0, 1)$ we have

$$\begin{aligned}
697 & \int_{\mathbb{R}^d} P_4(t, v) g_n^{**} \{t, v; r_{n1}(t), r_{n2}(t)\} dv \\
698 & \\
699 & \leq \left(\int_{E_2} + \int_{E_2^c} \right) P_4(t, v) K \left[\frac{1}{a_n \varepsilon_n} v + \{Ds(\theta_0) + r_{n1}(t)\}t - \frac{1}{a_n \varepsilon_n} A(\theta_0)^{1/2} T_{\text{obs}} \right] \tag{19} \\
700 & \\
701 & \times \frac{M_2^d}{(2\pi)^{d/2}} \exp \left(-\frac{m_2^2}{2} \|v\|^2 \right) dv \\
702 & \\
703 & \leq P_4(t) \left(\bar{K} \left[\lambda_{\min}^2(\Lambda) (1-\beta_1) \|\{Ds(\theta_0) + r_{n1}(t)\}t - \frac{1}{a_n \varepsilon_n} A(\theta_0)^{1/2} T_{\text{obs}}\|^2 \right] \right. \\
704 & \\
705 & \left. + \exp \left[-\frac{a_n^2 \varepsilon_n^2 \beta_1 m_2^2 \beta_2}{2} \|\{Ds(\theta_0) + r_{n1}(t)\}t - \frac{1}{a_n \varepsilon_n} A(\theta_0)^{1/2} T_{\text{obs}}\|^2 \right] \right), \tag{20}
\end{aligned}$$

710 where $P_4(t) \in \mathbb{P}_{4,t}$. Then using $(a+b)^2 \leq 2(a^2+b^2)$,

$$\begin{aligned}
711 & \int_{t(B_\delta)} \left[\int_{\mathbb{R}^d} P_4(t, v) g_n^{**} \{t, v; r_{n1}(t), r_{n2}(t)\} dv \right]^2 dt \\
712 & \\
713 & \leq \int_{t(B_\delta)} P_8(t) \bar{K}^2 \left[\frac{\lambda_{\min}^2(\Lambda) (1-\beta_1)}{2} \|\{Ds(\theta_0) + r_{n1}(t)\}t - \frac{1}{a_n \varepsilon_n} A(\theta_0)^{1/2} T_{\text{obs}}\|^2 \right] dt \\
714 & \\
715 & + \int_{t(B_\delta)} P_8(t) \exp \left[-\frac{a_n^2 \varepsilon_n^2 \beta_1 m_2^2 \beta_2}{2} \|\{Ds(\theta_0) + r_{n1}(t)\}t - \frac{1}{a_n \varepsilon_n} A(\theta_0)^{1/2} T_{\text{obs}}\|^2 \right] dt
\end{aligned}$$

719 Applying Lemma 7 on these upper bounds, (b) and (c) hold.

For (a), to see that the limit of $\int_{t(B_\delta)} \left\{ \int_{\mathbb{R}^d} g_n(t, v) dv \right\}^2 dt$ is lower bounded away from zero, just use the positivity of the limit of the integrand and Fatou's lemma to interchange the order of limit and integral. \square

2.3. Proof of Theorem 3

Now let $w_n(\theta)$ be the importance weight $\pi(\theta)/q_n(\theta)$, define $\pi_{B_\delta, \text{IS}}(h) = \int_{B_\delta} h(\theta) \pi(\theta) f_{\text{ABC}}(s_{\text{obs}} | \theta) w_n(\theta) d\theta$ and define $\pi_{B_\delta^c, \text{IS}}(h)$ correspondingly. Then by (15), we have

$$\Sigma_{\text{ABC}, n} = p_{\text{acc}, \pi}^{-1} \frac{\pi_{B_\delta, \text{IS}}(G_n) + \pi_{B_\delta^c, \text{IS}}(G_n)}{\pi_{B_\delta}(1) + \pi_{B_\delta^c}(1)}. \quad (21)$$

Proof of Theorem 3. For p_{acc, q_n} , we only need to consider the case when $\beta = 0$. Recall that $t(\theta) = a_{n, \varepsilon}(\theta - \theta_0)$. By the transformation $t = t(\theta)$, since $a_{n, \varepsilon} \sigma_n = 1$, $q_n(\theta) = a_{n, \varepsilon}^p |\Sigma|^{-1/2} q\{\Sigma^{-1/2}(t - c_\mu)\}$. Then, similar to the expansion of $\pi(1)$ from Lemma 3,

$$\begin{aligned} p_{\text{acc}, q_n} &= \varepsilon_n^d \int q_n(\theta) f_{\text{ABC}}(s_{\text{obs}} | \theta) d\theta \\ &= \varepsilon_n^d |\Sigma|^{-1/2} \int_{t(B_\delta)} q\{\Sigma^{-1/2}(t - c_\mu)\} \tilde{f}_{\text{ABC}}(s_{\text{obs}} | \theta_0 + a_{n, \varepsilon}^{-1}t) dt \{1 + o_p(1)\}. \end{aligned}$$

The above integral differs from $\tilde{\pi}_{B_\delta}(1)$ by replacing $\pi(\theta_0 + a_{n, \varepsilon}^{-1}t)$ with the density $q\{\Sigma^{-1/2}(t - c_\mu)\}$ which does not degenerate to a constant as $n \rightarrow \infty$. We will show that this integral has order $\Theta_p(1)$. Plugging in the expansion (17) of $\tilde{f}_{\text{ABC}}(s_{\text{obs}} | \theta_0 + a_{n, \varepsilon}^{-1}t)$ into p_{acc, q_n} , we can obtain an expansion similar to (11), differing in that parts from expanding $\pi(\theta_0 + a_{n, \varepsilon}^{-1}t)/|A(\theta_0 + a_{n, \varepsilon}^{-1}t)|^{1/2}$ are replaced by the Taylor expansion

$$\frac{q\{\Sigma^{-1/2}(t - c_\mu)\}}{|A(\theta_0 + a_{n, \varepsilon}^{-1}t)|^{1/2}} = q\{\Sigma^{-1/2}(t - c_\mu)\} \left[1 + a_{n, \varepsilon}^{-1} D_\theta \frac{1}{|A\{\theta_0 + \varepsilon_6(t)\}|^{1/2}} t \right],$$

where $\|\varepsilon_6(t)\| \leq \delta$. The explicit form is omitted here to avoid repetition. It can be seen that $p_{\text{acc}, q_n} = \Theta_p(a_{n, \varepsilon}^d \varepsilon_n^d)$ if (a) $\int_{\mathbb{R}^d \times t(B_\delta)} q\{\Sigma^{-1/2}(t - c_\mu)\} g_n(t, v) dv dt = \Theta_p(1)$; (b) $\int_{\mathbb{R}^d \times t(B_\delta)} P_3(t, v) q\{\Sigma^{-1/2}(t - c_\mu)\} g_n\{t, v; r_{n1}(t), r_{n2}(t)\} dv dt = O_p(1)$ when $\lim_{n \rightarrow \infty} a_n \varepsilon_n < \infty$; and (c) $\int_{\mathbb{R}^d \times t(B_\delta)} P_3(t, v) q\{\Sigma^{-1/2}(t - c_\mu)\} g_n^{**}\{t, v; r_{n1}(t), r_{n2}(t)\} dv dt = O_p(1)$ when $\lim_{n \rightarrow \infty} a_n \varepsilon_n = \infty$, where $r_{n1}(t)$ and $r_{n2}(t)$ are defined as in the proof of Theorem 2. Since $q\{\Sigma^{-1/2}(t - c_\mu)\}$ is uniformly upper bounded for $t \in \mathbb{R}^p$, (b) and (c) hold and the integral in (a) is $O_p(1)$ following the arguments for the similar cases in the proof of Theorem 2. By the positivity of the limit of the integrand and Fatou's lemma, the limit of the integral in (a) is lower bounded away from 0. Therefore $p_{\text{acc}, q_n} = \Theta_p(a_{n, \varepsilon}^d \varepsilon_n^d)$ holds.

As $\Sigma_{\text{IS}, n}$ is equal to $p_{\text{acc}, q_n} \Sigma_{\text{ABC}, n}$, by (15) we have

$$\Sigma_{\text{IS}, n} = \frac{p_{\text{acc}, q_n} \pi_{B_\delta, \text{IS}}(G_n) + \pi_{B_\delta^c, \text{IS}}(G_n)}{p_{\text{acc}, \pi} \pi_{B_\delta}(1) + \pi_{B_\delta^c}(1)} = \frac{p_{\text{acc}, q_n} \pi_{B_\delta, \text{IS}}(G_n)}{p_{\text{acc}, \pi} \pi_{B_\delta}(1)} \{1 + o_p(1)\},$$

where the second equality holds by noting that $\omega_n(\theta) \leq \beta^{-1}$. Given the obtained orders of p_{acc, q_n} and $p_{\text{acc}, \pi}$, $\Sigma_{\text{IS}, n} = O_p(a_{n, \varepsilon}^{-2})$ if $\pi_{B_\delta, \text{IS}}(G_n)/\pi_{B_\delta}(1) = O_p(a_{n, \varepsilon}^{-p-2})$. Similar to (16), we have the

769 following expansion

$$770 \frac{\pi_{B_\delta, \text{IS}}(G_n)}{\pi_{B_\delta}(1)} = G(\theta_0) \frac{\pi_{B_\delta, \text{IS}}(1)}{\pi_{B_\delta}(1)}$$

$$771 + 2a_{n, \varepsilon}^{-1} \{h(\theta_0) - h_{\text{ABC}}\} \frac{\pi_{B_\delta, \text{IS}}\{Dh(\theta_t)^T t\}}{\pi_{B_\delta}(1)} + a_{n, \varepsilon}^{-2} \frac{\pi_{B_\delta, \text{IS}}\{t^T Dh(\theta_t) Dh(\theta_t)^T t\}}{\pi_{B_\delta}(1)},$$

775 and we only need $\pi_{B_\delta, \text{IS}}\{P_2(t)\}/\pi_{B_\delta}(1) = O_p(a_{n, \varepsilon}^{-p})$ for any $P_2(t) \in \mathbb{P}_{2, t}$. Since $w_n(\theta) \leq (1 - \beta)^{-1} w_{n, 0}(\theta)$, where $w_{n, 0}(\theta)$ is the weight when $\beta = 0$, it is sufficient to consider the case $\beta = 0$.

777 Similar to the proof of Theorem 1, first the normal counterpart $\tilde{\pi}_{B_\delta, \text{IS}}\{P_2(t)\}/\tilde{\pi}_{B_\delta}(1)$ of $\pi_{B_\delta, \text{IS}}\{P_2(t)\}/\pi_{B_\delta}(1)$, where $f_{\text{ABC}}(s_{\text{obs}} | \theta)$ is replaced by $\tilde{f}_{\text{ABC}}(s_{\text{obs}} | \theta)$, is considered, then it is shown that their difference can be ignored. Using the transformation $t = t(\theta)$ and plugging in expansion (17) of $\tilde{f}_{\text{ABC}}(s_{\text{obs}} | \theta_0 + a_{n, \varepsilon}^{-1} t)$ into $\tilde{\pi}_{B_\delta, \text{IS}}\{P_2(t)\}$, we obtain an expansion similar to (11), differing in that parts from expanding $\pi(\theta_0 + a_{n, \varepsilon}^{-1} t)/|A(\theta_0 + a_{n, \varepsilon}^{-1} t)|^{1/2}$ are replaced by the Taylor expansion

$$778 \frac{1}{q_n(\theta)} \frac{\pi(\theta_0 + a_{n, \varepsilon}^{-1} t)^2}{|A(\theta_0 + a_{n, \varepsilon}^{-1} t)|^{1/2}}$$

$$779 = \frac{1}{a_{n, \varepsilon}^p |\Sigma|^{-1/2} q\{\Sigma^{-1/2}(t - c_\mu)\}} \left[\pi(\theta_0)^2 + a_{n, \varepsilon}^{-1} D_\theta \frac{\pi\{\theta_0 + \varepsilon_7(t)\}^2}{|A\{\theta_0 + \varepsilon_7(t)\}|^{1/2}} t \right],$$

782 where $\|\varepsilon_7(t)\| \leq \delta$. The explicit form is omitted here to avoid repetition. Then it can be seen that if we can show that

$$783 (d) \int_{t(B_\delta)} \frac{\int_{\mathbb{R}^d} P_5(t, v) g_n\{t, v; r_{n1}(t), r_{n2}(t)\} dv}{q\{\Sigma^{-1/2}(t - c_\mu)\}} dt = O_p(1) \text{ when } \lim_{n \rightarrow \infty} a_n \varepsilon_n < \infty,$$

$$784 (e) \int_{t(B_\delta)} \frac{\int_{\mathbb{R}^d} P_5(t, v) g_n^{**}\{t, v; r_{n1}(t), r_{n2}(t)\} dv}{q\{\Sigma^{-1/2}(t - c_\mu)\}} dt = O_p(1) \text{ when } \lim_{n \rightarrow \infty} a_n \varepsilon_n = \infty,$$

787 where $r_{n1}(t)$ and $r_{n2}(t)$ are defined as in the proof of Theorem 2, $\tilde{\pi}_{B_\delta, \text{IS}}\{P_2(t)\} = O_p(a_{n, \varepsilon}^{d-2p})$ and $\tilde{\pi}_{B_\delta, \text{IS}}\{P_2(t)\}/\tilde{\pi}_{B_\delta}(1) = O_p(a_{n, \varepsilon}^{-p})$ by Lemma 5. By (18) and the following equality for $d \times p$ full column-rank matrix A and vector c ,

$$788 \|At - c\| = \|P^{1/2}(t - P^{-1}Ac)\|^2 + c^T(I - AP^{-1}A^T)c,$$

789 where $P = A^T A$ and $P^{1/2}P^{1/2} = P$, for (d) we have

$$790 \frac{\int_{\mathbb{R}^d} P_5(t, v) g_n\{t, v; r_{n1}(t), r_{n2}(t)\} dv}{q\{\Sigma^{-1/2}(t - c_\mu)\}}$$

$$791 \leq P_5(t) \frac{\exp\left\{-\frac{m_2^2 m_2^2 \gamma}{2} \|t - P(\theta_0, t)T_{\text{obs}}\|^2\right\}}{q\{\Sigma^{-1/2}(t - c_\mu)\}} \exp\left[-\frac{m_2^2 \Delta}{2} \|\{Ds(\theta_0) + r_{n1}(t)\}t - A(\theta_0)^{1/2}T_{\text{obs}}\|^2\right]$$

$$792 + P_5(t) \frac{\bar{K}^\alpha \left\{\frac{\lambda_{\min}^2(\Lambda)(1-\gamma-\Delta)m_1^2}{a_n^2 \varepsilon_n^2} \|t - P(\theta_0, t)T_{\text{obs}}\|^2\right\}}{q\{\Sigma^{-1/2}(t - c_\mu)\}}$$

$$793 \times \bar{K}^\Delta \left[\frac{\lambda_{\min}^2(\Lambda)(1-\gamma-\Delta)}{a_n^2 \varepsilon_n^2} \|\{Ds(\theta_0) + r_{n1}(t)\}t - A(\theta_0)^{1/2}T_{\text{obs}}\|^2 \right],$$

794 where $P(\theta_0, t) = [\{Ds(\theta_0) + r_{n1}(t)\}^T \{Ds(\theta_0) + r_{n1}(t)\}]^{-1} \{Ds(\theta_0) + r_{n1}(t)\}^T A(\theta_0)^{1/2}$, both $P_5(t)$ belong to $\mathbb{P}_{5, t}$ and Δ is chosen such that $\gamma + \Delta \in (0, 1)$ and $\alpha + \Delta \in (0, 1)$ for γ

and α in Condition 7. Then since both ratios on the right hand side of the above inequality are $O_p(1)$ by Condition 7, by Lemma 7 and Lemma 9, (d) holds. Similarly by (20), for (e) we have

$$\begin{aligned}
& \frac{\int_{\mathbb{R}^d} P_5(t, v) g_n^{**}\{t, v; r_{n1}(t), r_{n2}(t)\} dv}{q\{\Sigma^{-1/2}(t - c_\mu)\}} \\
& \leq P_5(t) \frac{\exp\left\{-\frac{m_1^2 m_2^2 \gamma}{2} \|t - \frac{1}{a_n \varepsilon_n} P(\theta_0, t) T_{\text{obs}}\|^2\right\}}{q\{\Sigma^{-1/2}(t - c_\mu)\}} \\
& \quad \times \exp\left[-\frac{(a_n^2 \varepsilon_n^2 \beta_1 \beta_2 - \gamma) m_2^2}{2} \|\{Ds(\theta_0) + r_{n1}(t)\}t - A(\theta_0)^{1/2} T_{\text{obs}}\|^2\right] \\
& \quad + P_5(t) \frac{\bar{K}^\alpha \left\{\lambda_{\min}^2(\Lambda)(1 - \beta_1) m_1^2 \|t - \frac{1}{a_n \varepsilon_n} P(\theta_0, t) T_{\text{obs}}\|^2\right\}}{q\{\Sigma^{-1/2}(t - c_\mu)\}} \\
& \quad \times \bar{K}^{1-\alpha} \left[\lambda_{\min}^2(\Lambda)(1 - \beta_1) \|\{Ds(\theta_0) + r_{n1}(t)\}t - A(\theta_0)^{1/2} T_{\text{obs}}\|^2\right],
\end{aligned}$$

where both $P_5(t)$ belong to $\mathbb{P}_{5,t}$. Thus by Condition 7, Lemma 7 and Lemma 9, (e) holds. Therefore $\tilde{\pi}_{B_\delta, \text{IS}}\{P_2(t)\}/\tilde{\pi}_{B_\delta}(1) = O_p(a_{n,\varepsilon}^{-p})$.

To show that $\pi_{B_\delta, \text{IS}}\{P_2(t)\}/\pi_{B_\delta}(1) = O_p(a_{n,\varepsilon}^{-p})$, similar to the discussion of (13), it is sufficient to show that

$$\frac{\pi_{B_\delta, \text{IS}}\{P_2(t)\} - \tilde{\pi}_{B_\delta, \text{IS}}\{P_2(t)\}}{\tilde{\pi}_{B_\delta}(1)} = O_p(\alpha_n^{-1} a_{n,\varepsilon}^{-p}). \quad (22)$$

With the transformation $t = t(\theta)$ we have $\pi_{B_\delta, \text{IS}}\{P_2(t)\} - \tilde{\pi}_{B_\delta, \text{IS}}\{P_2(t)\}$ is equal to

$$\alpha_n^{-1} a_{n,\varepsilon}^{-2p} \int_{t(B_\delta)} \int P_2(t) \pi(\theta_0 + a_{n,\varepsilon}^{-1} t)^2 \frac{r_n(s_{\text{obs}} + \varepsilon_n v \mid \theta_0 + a_{n,\varepsilon}^{-1} t) K(v)}{|\Sigma|^{-1/2} q\{\Sigma^{-1/2}(t - c_\mu)\}} dv dt.$$

Then by following the arguments of the proof of Lemma 6, we have

$$\begin{aligned}
& |\pi_{B_\delta, \text{IS}}\{P_2(t)\} - \tilde{\pi}_{B_\delta, \text{IS}}\{P_2(t)\}| \leq \alpha_n^{-1} a_{n,\varepsilon}^{d-2p} \sup_{\theta \in B_\delta} |\pi(\theta)^2 A(\theta)^{-1/2}| \\
& \quad \times \int_{t(B_\delta)} \int |P_2(t)| \frac{(a_n a_{n,\varepsilon}^{-1})^d r_{\max}\left[a_n a_{n,\varepsilon}^{-1} M\left\{Ds(\theta_0 + \varepsilon t) - a_{n,\varepsilon} \varepsilon_n v - \frac{1}{a_n a_{n,\varepsilon}^{-1}} A(\theta_0)^{1/2} T_{\text{obs}}\right\}\right] K(v)}{q\{\Sigma^{-1/2}(t - c_\mu)\}} dv dt.
\end{aligned}$$

The ratio above is similar to the ratio of $g_n\{t, v; r_1(t), r_2(t)\}/q\{\Sigma^{-1/2}(t - c_\mu)\}$ except that the normal density is replaced by $r_{\max}(\cdot)$. Then by Condition 7, previous arguments for proving (iv) and (v) can be followed. Hence $\pi_{B_\delta, \text{IS}}\{P_2(t)\} - \tilde{\pi}_{B_\delta, \text{IS}}\{P_2(t)\} = O_p(\alpha_n a_{n,\varepsilon}^{d-2p})$ and (22) holds. Therefore $\Sigma_{\text{IS},n} = O_p(a_{n,\varepsilon}^{-2})$. \square

REFERENCES

- CREEL, M. & KRISTENSEN, D. (2013). Indirect likelihood inference (revised). UFAE and IAE working papers, Unitat de Fonaments de l'Anlisi Econmica (UAB) and Institut d'Anlisi Econmica (CSIC).
 LEHMANN, E. L. (2004). *Elements of large-sample theory*. Springer Science & Business Media.