

1 **Prekindergarten and Kindergarten Predictors of Reading Comprehension in Monolingual**
2 **English and Spanish-English Bilingual 6th Graders**

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23 **Keywords:**

24 reading comprehension; children; predictors

25 **Preregistration**

26 We completed an initial pre-registration of the analysis of Grade 6 reading comprehension in
27 monolingual English and Spanish-English bilingual children (referred to as Dual Language
28 Learners in that document). Prior to data analyses we revised our pre-registration plan. The
29 decision to use item-level Bayesian mixed-effects models was specified in the revised
30 preregistration prior to analysis. Links to preregistration documents include: <https://osf.io/4sfnu>,
31 <https://osf.io/whacq>, and <https://osf.io/v9ntx> and the addendum <https://osf.io/c2pwz>. The revised
32 pre-registration plan that we adhered to in these analyses is here: <https://osf.io/dt5gu>.

33 **Conflict of Interest**

34 The authors declare there are no real or potential conflicts of interest that could be seen as having
35 an influence on this research.

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Abstract

Purpose: Poor reading comprehension performance by children in the US is a continuing concern. Early identification and intervention can reduce the number of children with significant reading problems. We documented prekindergarten and kindergarten predictors of Grade 6 reading comprehension for English monolingual and Spanish-English bilingual groups and identified measures available to educators to flag children at risk for future reading comprehension problems.

Method: In Grade 6, children in the monolingual (N = 88) and bilingual groups (N = 95) completed a reading comprehension measure. These children were in a longitudinal study with previously completed code-related, vocabulary, grammar, listening comprehension, higher-level language, working memory, and nonverbal IQ measures in prekindergarten and kindergarten.

Results: Data were analyzed separately by language group. We fit a series of Bayesian mixed-effects models incorporating prekindergarten and kindergarten measures as predictors. For the monolingual group, the most promising prekindergarten predictor was letter identification, and the most promising kindergarten predictors were vocabulary, grammar/morphology, and listening comprehension. For the bilingual group, the most promising prekindergarten predictors were English vocabulary, listening comprehension in Spanish, and memory updating in Spanish, and the most promising kindergarten predictor was English vocabulary.

Conclusion: We suggest measures that could be administered in prekindergarten and kindergarten to flag students who may be at risk for future reading comprehension problems. We review other steps that schools and families may take to prevent reading comprehension problems.

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Introduction

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Given the relatively unchanged National Assessment of Educational Progress (NAEP) reading scores over the past two decades, it is fair to ask whether anything can be done to improve reading comprehension in the U.S., especially for bilingual children who, as a group, score significantly lower than their monolingual English peers (e.g., Friesen et al., 2022; Proctor et al., 2012; Silverman et al., 2015). Prevention is one potential solution (McCardle et al., 2001; Pikulski, 1994; Snow et al., 1998), with convincing evidence that early identification and intervention can reduce the number of children with significant reading problems (Linan-Thompson et al., 2006; McMaster et al., 2005). Because Head Start offers prekindergarten (PK) screening and 73% of public schools offer kindergarten (K) screening (Shields et al., 2016), accurate early identification is a promising approach for identifying children who need tailored instruction early in schooling; however, the best combination of predictors for later reading comprehension has not yet been identified for either monolingual or bilingual children.

Research shows that reading comprehension is a complex skill set involving multiple constructs (e.g., LARRC & Logan, 2017; Oakhill & Cain, 2012; Ortiz et al., 2012). To identify early indicators of later reading comprehension, it is necessary to assess a range of theoretically motivated constructs in early childhood. For three decades the Simple View of Reading (Gough & Tunmer, 1986) has provided a valuable, testable, framework for understanding reading comprehension that has been supported in English (e.g., Hoover & Gough, 1990; Kershaw & Schatschneider, 2012; LARRC, 2015c; Oakhill et al., 2003; Tilstra et al., 2009) and many other languages including Spanish (e.g., Mancilla-Martinez & Lesaux, 2010) and Greek (Protopapas et al., 2012). In general, word reading and linguistic comprehension, the two central components of the framework, have explained from 40% to 80% of the variance in reading comprehension in

81 English-speaking children (Bell & Wheldall, 2022; Joshi et al., 2015; Nation, 2019), with the
82 majority of these studies being conducted in grades 1-3.

83 Only a small number of longitudinal studies have examined predictors of reading
84 comprehension that extend beyond the code-related and oral language skills emphasized in the
85 Simple View of Reading (Gough & Tunmer, 1986) by incorporating additional cognitive factors
86 such as working memory and nonverbal IQ. In these studies, “code-related skills” included
87 foundational processes that support word-reading development, including phonological
88 awareness, letter identification, decoding, and orthographic knowledge. To provide the most
89 appropriate points of comparison for our work, we reviewed studies in which children were
90 assessed beginning in K or first grade, and in which those early skills were used to predict
91 reading comprehension outcomes measured from Grade 2 through high school. Apart from
92 LARRC and Chiu (2018) there were no longitudinal studies that assessed children in preschool
93 to predict later reading comprehension. We included K and grade 1 as predictor grades when
94 children are first learning to read to support early screening for language and literacy disorders.

95 Two studies included poor readers (Fuchs et al., 2012; Peng et al., 2019). Each study
96 measured code-related skills and oral language plus cognitive abilities such as non-verbal
97 reasoning, verbal memory, working memory, rapid naming, or IQ in English. Across all of these
98 studies code-related skills, oral language, and cognitive measures were significant predictors of
99 subsequent reading comprehension from one to nine years later. These results suggest that to best
100 predict reading comprehension, cognitive as well as code-related and oral language measures
101 should be assessed in early grades.

102 Across reviewed studies participants were from the US, UK, Norway, Finland, and
103 Canada meaning that they spoke multiple languages and represented a wide variety of school
104 systems.

105 We know that bilingual and monolingual children differ in language experience,
106 vocabulary, and developmental timing; however, prior work suggests that many core processes
107 underlying reading comprehension operate similarly within each group. For example, studies of
108 Spanish–English bilingual learners have shown that vocabulary, syntax, and word-reading skills
109 predict English reading comprehension, and that Spanish language skills can support English
110 comprehension through cross-linguistic transfer (Farnia & Geva, 2013; Proctor et al., 2006;
111 Mancilla-Martinez & Lesaux, 2010). However, it is also well established that bilingual and
112 monolingual children do not share identical underlying measurement structures for oral language
113 or reading comprehension. Studies conducted with the same or overlapping longitudinal samples
114 used in the present study have demonstrated that oral language skills, including vocabulary,
115 grammar, and listening comprehension, are not measurement-invariant across Spanish–English
116 bilingual and English monolingual children (LARRC, Yeomans-Maldonado, Bengochea, &
117 Mesa, 2018). This lack of invariance makes direct statistical comparisons across the two
118 language groups inappropriate. Accordingly, our aim is not to compare predictors between
119 bilingual and monolingual children, but rather to identify, within each group separately, the early
120 language, code-related, and cognitive skills most strongly associated with Grade 6 reading
121 comprehension. This approach permits careful interpretation of group-specific developmental
122 patterns while acknowledging that similar processes may underlie reading development even
123 when constructs do not behave identically across populations.

124 In these longitudinal studies, significant code-related predictors of later reading
125 comprehension included letter knowledge (Manu et al., 2021 in Finnish monolinguals; Peng et
126 al., 2019 in English monolinguals; Psyridou et al., 2021 in Finnish monolinguals), phonological
127 awareness (Farnia & Geva, 2013 in English bilinguals; Psyridou et al., 2021 in Finnish
128 monolinguals), word reading (Farnia & Geva, 2013 in English bilinguals; Fuchs et al., 2012 in
129 English monolinguals; Lervag et al., 2019 in Norwegian monolinguals; Peng et al., 2019 in
130 English monolinguals; Psyridou et al., 2021 in Finnish monolinguals; Sparks et al., 2014 in
131 English monolinguals), and reading fluency (Fuchs et al., 2012 in English monolinguals).

132 Oral language skills were also significant predictors of later reading comprehension using
133 an oral language factor that included vocabulary, grammar, and discourse (LARRC & Chiu,
134 2018) as well as separate indicators of grammar (Lervag et al., 2018), syntax (Farnia & Geva,
135 2013), vocabulary (Farnia & Geva, 2013; Lervag et al., 2019; Manu et al., 2021; Peng et al.,
136 2019; Psyridou et al., 2021; Sparks et al., 2014), reading inferencing (Lervag et al., 2019; Oakhill
137 & Cain, 2012), listening comprehension (Fuchs et al., 2012; Lervag et al., 2019; Manu et al.,
138 2021), comprehension monitoring (Oakhill & Cain, 2012), and passage comprehension (Sparks
139 et al., 2014). Of these, vocabulary was a significant predictor in 6 of 7 studies with reading
140 comprehension measured as early as second grade and as late as high school.

141 The longitudinal studies assessed cognitive development using a wide range of measures.
142 Significant predictors of later reading comprehension included nonverbal reasoning (Fuchs et al.,
143 2012; Peng et al., 2019), verbal memory (Lervag et al., 2018), rapid naming (Farnia & Geva,
144 2013; Fuchs et al., 2012; Manu et al., 2021; Psyridou et al., 2021), working memory (Farnia &
145 Geva, 2013) and IQ (Sparks et al., 2014).

146 Together, this literature suggests that while many early skills relate to later reading
147 comprehension, fewer studies have examined how predictors operate across adjacent early
148 grades within linguistically distinct groups. The present study addresses this gap by examining
149 PK and K predictors within monolingual and bilingual groups separately.

150 Assessments in the present longitudinal study included multiple code-related, oral
151 language, and cognitive measures for two groups of children: those who entered PK speaking
152 English and those who entered PK speaking primarily Spanish. As described below, those who
153 entered PK speaking primarily Spanish became bilingual as they were educated in English only
154 schools beginning in K.

155 Although the present study was not designed to evaluate diagnostic classification or
156 screening accuracy, identifying early measures associated with later reading comprehension is
157 important because it informs hypotheses about which constructs warrant attention in early
158 educational contexts. In seeking to make accurate and generalizable predictions of later reading
159 comprehension, our analyses targeted the accuracy of response to comprehension test items at
160 Grade 6 so that prediction models could take into account both participant- and item-level
161 sources of variance.

162 Our research questions were:

- 163 1. Which language, cognitive and code-related skills in PK or K best predict reading
164 comprehension in Grade 6 for the monolingual and bilingual groups?
- 165 2. Of these predictors, which PK and K measures could flag children who may benefit from
166 more intense instruction to promote future reading comprehension?

167 **Method**

168 This study was approved by the Arizona State University Internal Review Board under
169 study numbers 1010005588 and 00006696.

170 **Participants**

171 Participants were initially enrolled in a 5-year longitudinal study conducted by LARRC.
172 The children who began PK speaking English (monolingual group; $N = 397$) were from Arizona,
173 Kansas, Missouri, Nebraska, and Ohio. The children who began PK speaking Spanish (bilingual
174 group; $N = 280$) were from Arizona. In Arizona by state law, instruction must be provided in
175 English.

176 Participants were recruited through partnerships with multiple school districts and
177 individual schools selected based on student population size, demographic diversity, and
178 willingness to participate in a longitudinal study. Within participating schools, all PK teachers
179 were invited to participate, and children were recruited through those classrooms. Families of all
180 children in participating classrooms received study information and consent materials. Children
181 were not randomly sampled; instead, eligibility was determined through parent and teacher
182 screening to confirm that participants met inclusionary criteria and were able to complete the
183 study assessments.

184 Because children entered the study in PK and then matriculated into many different
185 kindergartens, the number of schools represented increased. Additional school changes occurred
186 as families relocated during grades K–5, resulting in a large number of schools represented in
187 Grade 6. In PK monolingual participants were enrolled in 25 different schools, in K 58 different
188 schools, and in Grade 6 in 57 different schools. In PK bilingual participants were enrolled in 33
189 different schools, in K in 41 different schools, and in Grade 6 in 54 different schools. Nine or
190 fewer participants were enrolled in the same school at any grade level for either group.

191 The monolingual English-speaking sample was drawn from multiple states, increasing
192 geographic and educational variability and strengthening the generalizability of findings within
193 that group. In contrast, the Spanish–English bilingual sample was recruited from a single state,
194 which limited geographic variability but resulted in a more linguistically cohesive group of
195 children who primarily spoke Mexican Spanish. This sampling structure aligns with our analytic
196 approach of examining predictors within each language group rather than making direct
197 comparisons across groups.

198 The purpose of both the LARRC and MBRC studies was to identify and model oral
199 language, cognitive, environmental, and code-related skills important for reading
200 comprehension. In the original LARRC study, data were collected annually from PK through
201 Grade 3. Subsequently, children were re consented and enrolled in a follow-up longitudinal study
202 conducted by the Monolingual and Bilingual Reading Comprehension (MBRC) group, which
203 collected additional data in Grade 6. Because not all original participants were retained through
204 Grade 6, we conducted attrition analyses to determine whether children who remained in the
205 study differed systematically from those who were lost to follow-up. Among monolingual
206 participants, attrition was not associated with gender, ethnicity, race, income, or eligibility for
207 free or reduced-price lunch, although attrition varied by site, with participants from Nebraska
208 more likely to be lost to follow-up than those from Arizona. Among bilingual participants, no
209 significant demographic differences were observed between retained and attrited participants.
210 Accordingly, the analytic sample used in the present study reflects children with Grade 6
211 outcome data, and results should be interpreted as characterizing associations within the retained
212 sample.

213 We also assessed whether the Ohio sample, which was lost during the transition from
214 LARRC to MBRC, differed from other sites at PK. Results indicated that Ohio participants were
215 more likely to be from low- and middle-income households and to qualify for free/reduced
216 lunch, although no differences were found in other demographics.

217 Previous published studies involving the same LARRC and MBRC samples include:
218 Currie et al., 2021; LARRC, 2015a; 2015b; 2015c; 2017a; 2017b; 2017c; Murphy et al., 2016;
219 LARRC & McIlraith, 2018; LARRC et al., 2018; LARRC & Muijselaar, 2018; LARRC & Chiu,
220 2018; LARRC et al., 2018a; LARRC et al. 2018b; LARRC et al., 2019a; LARRC et al., 2019b;
221 LARRC et al. 2021). These papers report analyses from grades PK through Grade 3 but have no
222 overlap in research questions with the current study.

223 Inclusionary criteria for participants in their PK year included: (a) child spoke English as
224 their native language for the monolingual group or Spanish as their native language for the
225 bilingual group based on parent report (some researchers refer to these children as dual-language
226 learners); (b) child had no significant speech, language, cognitive, sensory, or motor disabilities
227 that would preclude participation in assessments according to parent and teacher reports; (c)
228 child was attending PK, and (d) child was eligible to enter K the following year. Language
229 dominance was not used to form subgroups within the bilingual sample. Instead, variability in
230 Spanish and English skills was treated as continuous and reflected directly in the predictor
231 measures used in the analyses.

232 For this study we used longitudinal data from PK, K, and Grade 6. Demographic
233 information about the Grade 6 samples is presented in Table 1 for the monolingual group and
234 Table 2 for the bilingual group. We intentionally separated the groups descriptively and in all
235 analyses (LARRC, 2015a).

236 Procedures

237 Children were assessed in the spring of each school year in a quiet room at their school or
238 at a community center or library. Data collection followed the same procedure each year. Trained
239 research assistants (RAs) administered the measures 1:1. For the bilingual group, Spanish-
240 English bilingual RAs provided the assessment in the target language with Spanish measures
241 administered on a different day than English measures. Children completed about six hours of
242 assessments divided into multiple sessions of 1 – 2 hours each. The order of test/task
243 administration was counterbalanced across participants, although the measures were arranged in
244 blocks based on the language of the day, the need for recording, and the age of the child. In PK
245 and K children received \$40 in literacy materials and parents received a \$15 gift card as an
246 incentive. In Grade 6 children received a \$125 incentive.

247 Measures

248 Our reading comprehension outcome measure in Grade 6 was the Woodcock Reading
249 Mastery Test - Passage Comprehension subtest (Woodcock, 1998), administered in English for
250 both groups. This subtest required students to read a sentence or short passage and identify a
251 missing word. To answer correctly, students must use a variety of comprehension and
252 vocabulary skills to understand the passage.

253 The LARRC and MBRC studies included multiple measures of key constructs associated
254 with reading comprehension. A subset of these measures (Table 3 for the monolingual group;
255 Table 4 for the bilingual group) was selected based on their availability at both PK and
256 kindergarten. This approach ensured consistency in construct representation across early grades
257 and allowed us to examine whether the same skills, measured with comparable instruments, were
258 associated with later reading comprehension across developmental stages. It also aligns with our

259 goal of identifying screening measures that could be used flexibly in either PK or kindergarten
260 settings.

261 We used measures administered in English for the monolingual group and in Spanish for
262 the bilingual group. In PK the children came to school speaking primarily Spanish, thus it would
263 not be appropriate to assess their skills in English. The exceptions were that we administered the
264 English language Peabody Picture Vocabulary Test-4 (PPVT-4; Dunn & Dunn, 2007) and the
265 English language Woodcock Reading Mastery Test (WRMT) Letter Identification (LI) task to
266 children in the bilingual group because they were receiving school instruction in English. More
267 details on measure selection are provided in the Analytic Approach section. Reliability
268 calculations for measures administered to the monolingual group are reported in Supplementary
269 Table 1 and to the bilingual group in Supplementary Table 2. Monolingual alphas in PK ranged
270 from .71-.96 and in kindergarten from .57-.96. Bilingual group alphas ranged from .60-.97 in PK
271 and .64-.96 in kindergarten.

272 ***Code-Related Skills in PK and K***

273 In PK and K, the Woodcock Reading Mastery Test - *Letter ID* subtest (Woodcock, 1998)
274 assessed the child's ability to identify alphabet letters presented in isolation. For the present
275 analyses, we included scores from the English version at PK and K for both groups.

276 ***Oral Language***

277 As shown in Tables 3 and 4, we assessed four oral language constructs: vocabulary,
278 grammar, listening comprehension, and higher-level language. Although the CELF-4 and CELF-
279 4 Spanish are designed for children ages 5-21, we administered them to four-year-olds in this
280 study but used only raw scores for analyses.

281 We assessed vocabulary breadth in the monolingual English group using the Peabody
282 Picture Vocabulary Test – 4th Edition (PPVT-4; Dunn & Dunn, 2007) and vocabulary depth
283 using the Clinical Evaluation of Language Fundamentals – 4th Edition Word Classes subtest
284 (CELF-4; Semel, Wiig, & Secord, 2003). Both are receptive measures. The bilingual group
285 completed the Clinical Evaluation of Language Fundamentals – 4th Edition Spanish Edition
286 Word Classes subtest (Semel, Wiig, & Secord, 2003) in Spanish in PK and K and the PPVT-4 in
287 English in K.

288 For grammar we administered the CELF-4 Word Structure subtest to the monolingual
289 group and the CELF-4 Spanish Word Structure subtest to the bilingual group. Both assess syntax
290 and morphology in their respective languages.

291 In PK and K, we assessed listening comprehension using modifications of the CELF-4
292 Understanding Spoken Paragraphs subtest for the monolingual group and the CELF-4 Spanish
293 Understanding Spoken Paragraphs for the bilingual group. The bilingual group also completed
294 the English CELF-4 Understanding Spoken paragraphs in K. Because the CELF-4 and CELF-4
295 Spanish do not have passages for PK, we created new paragraphs and questions using lexile and
296 readability measures appropriate for PK. In K the existing paragraphs were divided so that only 1
297 trial paragraph and two additional paragraphs were administered.

298 To assess higher-level language, we administered the Inference-Making Stories Task, an
299 experimenter-made measure based on the work of Cain and Oakhill (1999), Oakhill and Cain,
300 (2012), and Currie and Cain (2015). Stories were presented in English to the monolingual group
301 and Spanish to the bilingual group in both PK and K. Children listened to an audio recording of
302 two short narratives and answered eight open-ended questions after each passage. In K one of the
303 stories was new and one story had been administered previously in PK. For each passage, four

304 questions assessed the ability to make local coherence inferences and four assessed the ability to
305 make global coherence inferences.

306 We also administered the Picture Arrangement Task (PAT), adapted from the Picture
307 Arrangement Test of the Wechsler Intelligence Scale for Children-3 (WISC-III UK edition;
308 Wechsler, 1992) that assessed children's ability to sequence a series of picture cards into a
309 coherent story. The total number of correct stories was tallied for the raw score, which was used
310 in analyses (possible range: 0–12). Internal consistency for our sample was $\alpha = .85$.

311 *Working Memory*

312 Working memory was assessed using the Auditory Working Memory subtest of the
313 Woodcock-Johnson Test of Cognitive Ability III normative update (WJ-III, Woodcock, 1998)
314 for the monolingual group and the Auditory Working Memory subtest of the Bateria III
315 Woodcock-Muñoz (Muñoz-Sandoval et al., 2005) for the bilingual group. This assessment tests
316 children's ability to listen to a list of words and numbers then to name the words in order
317 followed by the numbers in order. The Spanish version does not include monosyllabic words.

318 We also administered an experimenter-made working memory updating measure based
319 on the work of Belacchi et al. (2010) that evaluated the child's ability to modify the contents of
320 working memory using comparison of objects. The child saw pictures of an increasing number of
321 objects and was asked a question like, "Tell me the names of the one/two/three/four/five smallest
322 things." Instructions and prompts were in Spanish for the bilingual group.

323 *Nonverbal IQ*

324 To assess nonverbal IQ, we administered the Kaufman Brief Intelligence Test – 2nd
325 Edition Matrices subtest (Kaufman & Kaufman, 1997) in PK, but not in K. The Matrices subtest
326 assesses children's ability to solve new problems, perceive relationships and complete visual

327 analogies without testing vocabulary or language skills. Spanish prompts were provided for the
328 bilingual sample. This is in line with the test manual instructions, which state, “When assessing
329 individuals whose comprehension of spoken English is limited, such as examinees who speak
330 little English or those with hearing impairments, the examiner may present directions for
331 Matrices in any foreign language or alternative communication method (such as signing) shared
332 by the examinee and the examiner. For the convenience of Spanish-speaking examiners, each
333 Matrices easel page presents administration instructions in both English and Spanish.”

334 **Results**

335 **Analysis Method**

336 We analyzed Grade 6 reading comprehension using item-level response accuracy on the
337 WRMT Passage Comprehension subtest. Building on prior work using item-level approaches to
338 reading comprehension (e.g., Kulesz et al., 2016; Miller et al., 2014; Spencer et al., 2019), we
339 used Bayesian generalized linear mixed-effects models (GLMMs) to estimate how
340 prekindergarten and kindergarten language, cognitive, and code-related measures predict the log-
341 odds of a correct response. In these models, the fixed effects represent the effects of early
342 predictors on comprehension accuracy, while the random effects account for variability
343 associated with differences between children and differences between items. The fixed effects in
344 these models represent the effects of early language, cognitive, and code-related predictors on
345 comprehension accuracy, while the random effects account for variability associated with
346 differences between children and differences between items. Modeling at the item level allows
347 simultaneous estimation of between-child differences in overall comprehension ability and
348 between-item differences in difficulty, preserving variability that would be obscured by summed

349 scores and aligning with prior longitudinal prediction work (e.g., Kulesz et al., 2016; Miller et
350 al., 2014; Spencer et al., 2019).

351 We assumed a Bernoulli likelihood function for item-level response accuracy and
352 specified weakly informative prior distributions for population- and group-level parameters to
353 stabilize estimation and regularize extreme estimates while allowing the data to primarily
354 determine parameter values. These priors reflect the assumption that fixed effects are likely to
355 fall within a moderate range on the log-odds scale and that random effects variances are positive
356 and typically small (McElreath, 2020). For the primary models, we specified Normal (mean = 0,
357 sd = 1) priors for fixed effects, Half-Student's t (location = 0, scale = 10, shape = 3) priors for
358 random effects variances, and LKJ (shape = 2) priors for random effects covariance structures.
359 The models included child-level and item-level random intercepts, along with random slopes
360 where appropriate, to reflect the hierarchical structure of the data and to allow the effects of early
361 predictors to vary across children and across items. This specification reflects the expectation
362 that both overall comprehension ability and the influence of early skills may differ across items
363 of varying difficulty. Residual item-level variance remained substantively non-zero after
364 accounting for child-level differences, indicating that responses varied across items beyond what
365 would be expected if Rasch assumptions fully held in this sample. Accordingly, modeling item-
366 level accuracy with item random effects provides a more appropriate framework than total-score
367 approaches for estimating associations between early predictors and later reading
368 comprehension.

369 To address our research questions, we fitted a series of models. To evaluate which
370 language, cognitive, or code-related skills influenced reading comprehension in Grade 6 we
371 fitted: (1) models including (as predictors or fixed effects) PK measurements of grammar,

372 vocabulary, higher-level language, listening comprehension, nonverbal IQ, and working
373 memory; and (2) models including K measurements of the same constructs. PK and K predictors
374 were modeled separately to (a) avoid collinearity introduced by including multiple timepoints in
375 a single model and (b) evaluate the unique predictive contribution of skills assessed at each
376 developmental stage. Analyses were conducted separately within each language group given
377 prior evidence of measurement non-invariance across monolingual and bilingual samples
378 (LARRC, Yeomans-Maldonado, Bengochea, & Mesa, 2018).

379 **Data Analysis Results**

380 The initial LARRC dataset included measures administered in PK (N = 397 monolingual,
381 280 bilingual) and K (N = 373 monolingual, 263 bilingual). Analyses were conducted using
382 participants with Grade 6 WRMT Passage Comprehension outcome data (N = 88 monolingual,
383 95 bilinguals; see Participants for attrition details). Results are presented separately for the
384 monolingual and bilingual groups, with PK and K predictor models reported within each group.

385 Using Bayesian mixed-effects models allowed us to include all available item-level
386 responses without listwise deletion; however, this analytic approach was selected primarily to
387 appropriately represent child- and item-level sources of variance rather than to address missing
388 data. The presence of missing item-level data did not drive the choice of analytic approach;
389 rather, item-level modeling was selected to appropriately represent child- and item-level sources
390 of variance and to address limitations of total-score approaches. Given some attrition, results
391 should be interpreted as estimating associations within the retained analytic samples, rather than
392 as definitive population-level effects.

393 ***Model Estimates for the Monolingual Group***

394 Table 5 presents estimates from the models used to analyze monolingual group
395 responses, including (1) PK predictor models and (2) K predictor models. These analyses address
396 Research Question 1 by identifying which PK and K predictors are most strongly associated with
397 later reading comprehension within the monolingual group. For each model we report coefficient
398 estimates on the log-odds scale, along with standard errors and 95% credible interval bounds.
399 Posterior probabilities are interpreted descriptively, with values $\geq .95$ indicating strong evidence
400 for an effect and values between .66 and .95 indicating moderate evidence. Posterior
401 distributions for each predictor are also displayed in Figures 2–3.

402 For the monolingual group, PK letter identification scores were associated with a higher
403 probability of correct responses in Grade 6 reading comprehension. The posterior distribution
404 indicates strong evidence for a positive effect ($\beta = .39$), with approximately 95% of the
405 distribution lying above zero. The credible interval suggests that the effect of PK code-related
406 skills ranges from small ($\beta \approx .03$) to relatively large ($\beta \approx .77$), while the probability of a negative
407 effect is very small ($< 2.5\%$). There was also moderate evidence that higher PK scores on
408 auditory memory, nonverbal intelligence (KBIT-2 Matrices), and grammatical knowledge
409 (CELF-4 Word Structure) were associated with a higher probability of correct responses.

410 For the kindergarten models, higher vocabulary (PPVT-4) and grammatical knowledge
411 (CELF-4 Word Structure) scores were associated with a higher probability of correct responses
412 in Grade 6 reading comprehension. The posterior distributions provide strong evidence for
413 positive effects of these predictors. There was also moderate evidence that listening
414 comprehension (CELF-4 Understanding Spoken Paragraphs) was positively associated with
415 response accuracy.

416 Figures 2 and 3 display the posterior probability distributions for the PK and K
417 predictors. Density plots illustrate the distribution of estimated effects for each predictor, while
418 the dot-and-whisker plots summarize these distributions using posterior medians and credible
419 intervals.

420 *Model Estimates for the Bilingual Group*

421 We present estimates for each model used to analyze bilingual group responses in Table 6
422 including: (1) PK predictor model estimates and (2) K predictor model estimates. We also
423 present these estimates as posterior probability distributions (Figures 5-7). These analyses
424 address Research Question 1 within the bilingual group by identifying early predictors
425 probabilistically associated with later English reading comprehension.

426 Together, the results summarized across PK and K models for each group address
427 Research Question 2 by identifying early measures that show the strongest probabilistic
428 associations with later reading comprehension, while explicitly acknowledging uncertainty and
429 sample-specific limitations.

430 For the bilingual group, higher PK scores on memory updating in Spanish, vocabulary in
431 English (PPVT-4), and listening comprehension in Spanish (CELF-4 Understanding Spoken
432 Paragraphs) were associated with a higher probability of correct responses in Grade 6 reading
433 comprehension. The data suggest that higher PK scores on the Spanish Inference Making Stories
434 test of higher-level language are also associated with response accuracy, although the posterior
435 distributions are compatible with small, null, or negative as well as positive effects.

436 In comparison, only higher kindergarten vocabulary scores (PPVT-4) showed strong
437 evidence (posterior probability ≥ 0.95) of predicting a higher probability of correct responses in
438 Grade 6 reading comprehension. There was also moderate posterior probability (0.66–0.95) that

439 higher scores on grammar (CELF-4 Word Structure in Spanish), working memory (Memory
440 Updating in Spanish), higher-level language ability (Inference Making Stories in Spanish), and
441 English code-related skills (Letter Identification) were associated with a higher probability of
442 correct responses. However, for each of these predictors, the posterior distributions also remain
443 compatible with small, null, or negative effects. Figures 4 and 7 illustrate how model-derived
444 effects translate into predicted probabilities of comprehension accuracy.

445 *Robustness Checks*

446 When multiple candidate measures were available for the same construct (e.g., PPVT-4
447 vs. CELF-4 Word Classes for vocabulary; Inference Making Stories vs. Picture Arrangement
448 Task for higher-level language), we refit robustness check models substituting alternate measures
449 to assess sensitivity to measurement choice (see Tables 3–4 and Supplemental Figures 3–4).
450 Overall, effect estimates were similar across alternate measurement selections, indicating
451 stability of the primary results. One notable pattern was that, for bilingual children, the predictive
452 influence of Kindergarten vocabulary on Grade 6 English comprehension was stronger when
453 vocabulary was measured with the English PPVT-4 than with the Spanish CELF-4 Word Classes
454 measure — suggesting that predicting an English outcome benefits from including English
455 vocabulary assessments.

456 We also checked whether, in analyzing Grade 6 reading comprehension for bilingual
457 children, entering both Spanish and English measures would help to inform predictions of
458 outcome comprehension accuracy. We did this only for the K predictors model because in PK we
459 administered measures in Spanish, thus language comparisons were not possible. We found that
460 the predictive effect of vocabulary or listening comprehension scores was much stronger for the
461 English than for the Spanish measures of these constructs (see Supplemental Figure 5), even

462 when measures in both languages were included in models. It is the case that English vocabulary
463 and higher-level language measures in PK and K are most clearly associated with predicted
464 change in reading comprehension outcome scores in English in Grade 6.

465 Because Bayesian inference depends on prior choices, we assessed prior sensitivity by
466 refitting models under several reasonable prior specifications: Normal (0,1), Normal (0,10), and
467 Student's t (3, 0, 10) for population- and group-level parameters, and LKJ(1) and LKJ(2) for
468 correlation/covariance structures. Estimates and substantive conclusions remained stable across
469 these specifications (see Supplemental Figures 6–7). Supplemental Figure 6 displays prior
470 sensitivity results for the monolingual group and Supplemental Figure 7 for the bilingual group;
471 Figure 8 provides an integrated visualization of these robustness checks.

472 Discussion

473 We identified prekindergarten and kindergarten language, cognitive, and code-related
474 predictors associated with Grade 6 reading comprehension within monolingual and bilingual
475 groups, characterizing group-specific patterns of association and informing targeted prevention
476 strategies (LARRC, Yeomans-Maldonado, Bengochea, & Mesa, 2018).

477 For each group, early predictors showed a developmental shift from code-related skills in
478 PK to broader language skills in kindergarten. Vocabulary emerged as the most consistent
479 predictor, with additional contributions from grammar, listening comprehension, and, for
480 bilingual children, cross-linguistic and cognitive factors.

481 Based on the literature, there are several reasons we may see different patterns of
482 predictors between language groups. For vocabulary, bilingual children showed additional
483 benefits from vocabulary in Spanish, which aligns with previous findings showing that bilingual
484 students can benefit from cross-linguistic transfer to leverage their Spanish vocabulary to

485 scaffold English reading comprehension (Grimm et al., 2018; LARRC et al., 2019; Proctor et al.,
486 2006). Because bilingual children must process and integrate information across languages, it is
487 not surprising that Spanish working memory, as assessed by memory updating, was a predictor
488 of Grade 6 reading comprehension in the bilingual group. These group-specific patterns are
489 consistent with theoretical and empirical work suggesting that reading comprehension develops
490 through interacting linguistic and cognitive processes that may differ in timing or salience across
491 language contexts.

492 **The Centrality of Vocabulary**

493 Our results illustrate the central role that vocabulary plays in subsequent reading
494 comprehension beginning in the PK years. Similar results were reported by Grimm et al. (2017)
495 in their longitudinal study of Spanish-speaking children and by Psyridou et al. (2021) and Manu
496 et al. (2021) in their longitudinal studies of Finnish speaking children. In those studies, K
497 vocabulary scores predicted significant variance in reading comprehension in 9th and 10th grades.

498 The importance of vocabulary for reading comprehension is well established - it was one
499 of the “five pillars” essential to reading instruction identified by the National Reading Panel in
500 2000 (NICHD, NIH). What is concerning, given the centrality of vocabulary in academic
501 achievement, is the lack of a consistent, systematic educational approach to increasing children’s
502 vocabulary breadth and depth (Black & Wright, 2023). Challenges include identifying which
503 words to teach (Wasik et al., 2016), how to teach them (Hadley & Mendez, 2021), a lack of
504 clarity on effective instructional approaches (e.g., Cervetti et al., 2023; Dai et al., 2024)
505 challenges with time availability in the school day (e.g. Wright & Neuman, 2014) and evidence
506 that many bilingual children begin school knowing fewer words than their peers (e.g., Morgan et
507 al., 2015; Rowe & Goldin-Meadow 2009).

508 Grammar and Morphology

509 In K, grammar and morphology were significant predictors of Grade 6 reading
510 comprehension for the monolingual and bilingual groups, each in their primary language. In an
511 earlier LARRC study examining the concurrent relationship between grammar and reading
512 comprehension in the monolingual group, grammar was most closely related to reading
513 comprehension in third grade (LARRC & Logan, 2017). Our results show that this relationship
514 holds longitudinally at least through Grade 6.

515 Grammatical knowledge is thought to support a reader's ability to monitor syntactic
516 information in the text for meaning and to generate meaning coherence during the reading
517 process (Fender, 2001). Morphological awareness, including understanding inflectional and
518 derivational morphemes, prefixes and suffixes, helps the reader decode multisyllabic words and
519 understand their meaning (Levesque et al., 2019; Sparks & Metsala, 2023).

520 A recent meta-analysis of 86 studies investigating the relationship between grammatical
521 knowledge and reading comprehension from elementary school through university in students
522 who spoke Chinese, Dutch, English, and French showed a large effect size (Fisher's $z = 0.54$)
523 with a significant interaction by grade (Zheng et al., 2023). The secondary school group had a
524 significantly larger effect size than the primary grade group. These results are consistent with our
525 finding that grammar and morphology in K predicted Grade 6 reading comprehension and are
526 also consistent with longitudinal studies by Lervag et al. (2018), who reported that grammar
527 assessed in Grade 2 Norwegian speaking children predicted reading comprehension in Grade 7
528 and Farnia and Geva (2013), who found that syntax in Grade 1 English language learners
529 predicted English reading comprehension in Grade 6.

530 Just as systematic vocabulary instruction could benefit students who may be at risk for
531 poor reading comprehension, systematic instruction for grammar and morphology can benefit
532 young learners (Buggy & Dillon, 2025). For example, a 2010 systematic review on the effects of
533 morphological instruction on literacy skills showed that this instruction was particularly
534 beneficial for less able readers and was as effective for young students (e.g. PK, K) as it was for
535 older students (Bowers et al, 2010). Morphological instruction was most effective when
536 integrated with phonological, orthographic, grammatical, and semantic information.

537 **Listening Comprehension**

538 The importance of listening comprehension as a predictor of reading comprehension is
539 emphasized in models such as the Simple View of Reading (Gough & Tunmer, 1986; Hoover &
540 Gough, 1990). In reading research, listening comprehension is considered a higher-level
541 language skill that integrates words, phrases, and sentences to form a mental model of the text.
542 Research shows that the importance of listening comprehension, relative to code-related,
543 increases over time as children become fluent readers (Garcia & Cain, 2014) and that
544 vocabulary, inferencing, background knowledge, working memory, and attention all impact
545 listening comprehension ability (Hogan et al., 2014). Kim and Pilcher (2016) suggested that
546 “listening comprehension instruction should be an integral part of reading and writing
547 instruction...” (p. 2) but also acknowledge that “...listening comprehension is a large problem
548 space” (p.19). By this they mean that, like reading comprehension, listening comprehension is
549 not a single skill that can be taught. It requires a systematic approach to building oral language
550 skills that support oral language understanding embedded in language-rich environments. To this
551 end, schools can evaluate whether their language arts, science, and math curricula promote oral
552 language development at the discourse level (e.g., conversations, stories, discussions, debates)

553 every day of the school year. Along with the lower level language skills of vocabulary, grammar,
554 and morphology, this is crucial for skilled reading comprehension.

555 **Letter Knowledge**

556 In our study PK letter knowledge was an important predictor of Grade 6 reading
557 comprehension for the monolingual group when all other predictors were in the model. Similar
558 results were reported by Leppanen et al. (2008), Manu et al (2021), and Psyridou et al, (2021) in
559 Finnish-speaking children. This builds on our work demonstrating the important role of letter
560 knowledge for reading comprehension in Grade 3 via word reading (LARRC & Chiu, 2018). The
561 indirect influence of letter knowledge on reading comprehension via word reading is supported
562 by a systematic review of PK predictors of later reading comprehension which found that code-
563 related skills, such as phoneme awareness and letter knowledge in PK, were indirectly related to
564 reading comprehension via word reading (Hjetland et al., 2017). Together this research
565 demonstrates that a strong foundation in code-related skills supports future word reading and
566 reading comprehension.

567 **Working Memory**

568 Working memory plays an important role in reading comprehension by holding and
569 updating the mental representation of information in the text referred to as the text base. Good
570 comprehenders keep explicit information from the text and inferencing results activated to
571 construct and update their own representation of text meaning, known as the situation model
572 (Kintsch, 2009).

573 It is noteworthy that performance on a Spanish memory updating task by bilingual PK
574 students was a significant predictor of reading comprehension seven years later in Grade 6.
575 Previous longitudinal studies of reading comprehension have not typically assessed working

576 memory and results for cross-sectional studies are mixed. Some have not found a direct relation
577 between working memory and reading comprehension (e.g., LARRC & Logan, 2017; Oakhill &
578 Cain, 2012) but some report a significant relationship. For example, Linares and Pelegrina
579 (2023) assessed fourth graders in Spain and found that after controlling for fluid intelligence,
580 reading and math fluency, vocabulary and reading ability, the retrieval component of working
581 memory updating was directly related to reading comprehension. A recent meta-analysis by Shin
582 (2020) examined the relationship between performance on variations of the reading span task
583 commonly used to assess working memory where participants listen to or read a series of
584 sentences and make some decision about the sentences such as whether each is grammatically
585 correct, while holding the final word of the sentence in memory. Shin reported a moderate
586 relationship between reading comprehension in L2 and working memory ($r = .30$).

587 **Limitations**

588 Several limitations should be considered when interpreting these findings. First, the
589 monolingual and bilingual samples differed in their geographic recruitment. Monolingual
590 English-speaking children were recruited from multiple states, whereas the Spanish–English
591 bilingual group was recruited solely from Arizona. This difference may limit the generalizability
592 of findings for bilingual learners to other Spanish-speaking populations in the United States. At
593 the same time, recruiting the bilingual sample from a single state resulted in a relatively cohesive
594 linguistic group, with most children speaking Mexican Spanish and experiencing similar
595 educational language policies. This reduced dialectal and sociolinguistic variability may have
596 increased internal validity for analyses conducted within the bilingual group. Importantly,
597 because prior work with these samples has demonstrated that oral language and reading
598 comprehension constructs are not measurement-invariant across bilingual and monolingual

599 children, we did not conduct direct cross-group comparisons. Accordingly, differences in
600 geographic recruitment should not be interpreted as a confound affecting between-group
601 contrasts but rather as contextual factors relevant to understanding group-specific developmental
602 patterns.

603 **Educational Implications**

604 The present findings do not provide evidence for the effectiveness of specific
605 instructional approaches; however, they inform the selection of early screening measures that
606 may help identify children who could benefit from additional instructional support. Based on our
607 results, we recommend different assessment approaches by population and grade. For children
608 who enter PK speaking monolingual English, we suggest the WRMT-R/NU Letter ID subtest;
609 for children who enter PK speaking primarily Spanish, we suggest the PPVT-4 or the CELF-4
610 Spanish Understanding Spoken Paragraphs subtest. In kindergarten, we suggest the PPVT-4 or
611 the CELF-4 Word Structure subtest for monolingual English-speaking children, and the PPVT-4
612 and the CELF-4 Spanish Word Structure subtest for Spanish–English bilingual children. The
613 predictive utility of these measures may vary across samples, outcome measures, and test
614 editions; therefore, they should be used as screening tools to guide instruction rather than for
615 diagnostic classification.

616 These assessments are often used in psychoeducational evaluations, but they may also
617 serve as screening tools to identify children who could benefit from additional instructional
618 support. Importantly, this use is distinct from identifying risk for a language or literacy disorder.
619 Measures such as the PPVT may yield lower scores in linguistically and culturally diverse
620 populations (e.g., Restrepo et al., 2006; Washington & Craig, 1992) and do not accurately
621 identify language impairment (Gray et al., 1999). Accordingly, these tools should be interpreted

622 within culturally and linguistically responsive frameworks. When used appropriately, they can
623 support educators in identifying areas for targeted instruction, including vocabulary, grammar,
624 and code-related skills, and in promoting equitable literacy outcomes.

625 **Artificial Intelligence Statement:** ChatGPT was used during the revision process for this
626 manuscript to reduce the number of words in the manuscript and to review the manuscript for
627 clarity, redundancy, and formatting. The resulting text was reviewed and revised by the authors
628 prior to inclusion in the manuscript

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Declaration of Interest Statement:

The authors report there are no competing interests to declare.

Supplemental Material

Supplemental material is available to accompany this manuscript.

Data Availability Statement

Portions of the data from this study may be available upon reasonable request from the study authors.

643

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Figure Captions

Figure 1

Mean Reading Comprehension Accuracy on the WRMT-R/NU Update Passage Comprehension Subtest by Child and by Test Item for the Monolingual Group

Note. Each point on the plot represents the mean response accuracy, by child or by item, in the reading comprehension measure. Means are calculated over all responses made by a child (for all items, to get the mean by child) or over all responses made to an item (made by all children, to get the mean by item). We have jittered the position of the points (adding a little random noise to the coordinates) because there may be multiple (by-child, or by-item) scores for some mean accuracy values.

Figure 2

Mean Reading Comprehension Accuracy on the WRMT-R/NU Update Passage Comprehension Subtest by Child and by Test Item for the Bilingual Group

Note. Each point on the plot represents the mean response accuracy, by child or by item, in the reading comprehension measure. Means are calculated over all responses made by a child (for all items, to get the mean by child) or over all responses made to an item (made by all children, to get the mean by item). We have jittered the position of the points (adding a little random noise to the coordinates) because there may be multiple (by-child, or by-item) scores for some mean accuracy values.

Figure 3

Posterior Probability Distribution for Prekindergarten Predictors of Reading Comprehension in the Monolingual Group

Note: Dots represent the posterior median. Lines represent the 66% (thicker line) and 95% (thinner line) credible intervals (CrI).

Figure 4

Posterior Probability Distribution for Kindergarten Predictors of Reading Comprehension in the Monolingual Group

Note. Dots represent the posterior median. Lines represent the 66% (thicker line) and 95% (thinner line) credible intervals.

Figure 5

Model Predictions for the Monolingual Group

Note. The plots represent the central (median) predicted change in the probability of outcome response accuracy, together with the 95% (light shaded ellipse) and the 66% (darker shaded ellipse) credible intervals. Plots indicate the predicted probability of a correct response given variation in code-related test scores (panel 1); and for the kindergarten predictors model, given

variation in grammar (panel 2), vocabulary (panel 3) and listening comprehension test scores (panel 4).

Figure 6

Posterior Probability Distribution for Prekindergarten Predictors of Reading Comprehension in the Bilingual Group

Note. Dots represent the posterior median. Lines represent the 66% (thicker line) and 95% (thinner line) credible intervals (CrI).

Figure 7

Posterior Probability Distribution for Kindergarten Predictors of Reading Comprehension in the Bilingual Group

Note. Dots represent the posterior median. Lines represent the 66% (thicker line) and 95% (thinner line) credible intervals.

Figure 8

Model Predictions for the Bilingual Group

Note. The plots represent the central (median) predicted change in the probability of outcome response accuracy, together with the 95% (light shaded ellipse) and the 66% (darker shaded ellipse) credible intervals. Plots indicate the predicted probability of a correct response for the Prekindergarten predictors model, given variation in code-related test scores (panel 1); and for the kindergarten predictors model, given variation in grammar (panel 2), vocabulary (panel 3) and listening comprehension test scores (panel 4).

Supplemental Figure 1

Predictor Variable Density Plots for the Monolingual Group at Each Measurement Time Point

Note. Scores range from 0 to the sample maximum for each variable. The height of the density curve reflects the number of children with a given score. The one-dimensional rug plot beneath each curve indicates observed values, with darker ticks representing multiple children sharing the same score. Plot titles indicate the construct and corresponding measurement.

Supplemental Figure 2

Predictor Variable Density Plots for the Bilingual Group at Each Measurement Time Point

Note. Scores range from 0 to the sample maximum for each variable. The height of the density curve reflects the number of children with a given score. The one-dimensional rug plot beneath each curve indicates observed values, with darker ticks representing multiple children sharing the same score. Plot titles indicate the construct and corresponding measurement.

Supplemental Figure 3

Posterior Probability Distributions for Prekindergarten and Kindergarten Models Including Different Measures of Vocabulary and Higher-Level Language Constructs

Note. Light green represents estimates from the primary model (reported in the main text), and dark green represents estimates from alternate models including different measures of vocabulary and higher-level language. Dots indicate posterior medians, and lines represent 95% credible intervals (CrI).

Supplemental Figure 4

Posterior Probability Distributions for Prekindergarten and Kindergarten Models Including Alternate Measures of Vocabulary and Higher-Level Language Constructs

Note. Light blue represents estimates from the primary model (reported in the main text), and dark blue represents estimates from alternate models. Dots represent the posterior median. Lines represent the 95% credible intervals (CrI).

Supplemental Figure 5

Posterior Probability Distributions for Kindergarten Models Including Both English and Spanish Measures

Note. In light blue we show the primary model estimates (reported in the main text) versus in dark blue alternate model estimates, for models incorporating both English and Spanish measures of vocabulary and listening comprehension. Dots represent the posterior median. Lines represent the 95% quantile credible intervals.

Supplemental Figure 6

Posterior Probability Distributions for Models with Prekindergarten or Kindergarten Measures as Predictors for the Monolingual Group

Note. Different shades of green represent estimates from models fitted under alternative prior specifications. Dots indicate posterior medians, and lines represent 95% credible intervals (CrI).

Supplemental Figure 7

Posterior Probability Distributions for Models with Prekindergarten or Kindergarten Measures as Predictors for the Bilingual Group

Note. Different shades of blue represent estimates from models fitted under different alternative prior specifications. Dots represent the posterior median. Lines represent the 95% credible intervals (CrI).