The ability to direct attention in working memory is not impaired in adults with symptoms of attention deficit hyperactivity disorder (ADHD)

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Abstract

Neurotypical individuals can prioritize particularly valuable information in working memory. This is a well-replicated effect, demonstrated across a wide variety of task factors and age groups. However, it is not clear if individuals with symptoms of ADHD are able to do this effectively, as there is some evidence this group struggle to allocate attention in working memory tasks. Two experiments were conducted online to investigate this. Participants were presented with series of four coloured shapes, and asked to report the colour of each shape in a counterbalanced order following a brief delay. In some trials (equal value condition), all shapes were equally valuable with the correct recall of each shape gaining the participant 2 points. In other trials (differential value condition), the first item presented during the encoding phase was more valuable than the rest (5 point for the first item vs 1 point for the other items). Trial-by-trial feedback was either provided (Experiment 1) or omitted (Experiment 2). Across both experiments, there was a clear prioritization effect at the first (targeted) serial position, with higher accuracy in the differential value condition relative to the equal value condition. There were also clear costs at the less valuable serial positions. These effects did not differ as a function of ADHD symptoms. There were also no significant correlations between scores on the Adult ADHD Self-Report Screener and the prioritization effects. Taken together, this demonstrates that the ability to prioritize particularly valuable information in working memory is not impaired in individuals with symptoms of ADHD.

Keywords: working memory; attention; reward; value-based prioritization; attention deficit hyperactivity disorder; ADHD

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Working memory refers to an individual's ability to store and process a limited amount of information for a brief period of time (e.g. seconds; Baddeley et al., 2021; Cowan, 2017). It is considered crucial for a range of everyday activities, including following instructions (Gathercole et al., 2006; Waterman et al., 2017), reading comprehension (Cain et al., 2004; Cowan, 2014), and mental arithmetic (Fürst & Hitch, 2000). Working memory abilities are subject to large individual differences (Alloway, 2006), and working memory difficulties often co-occur with a range of neurodevelopmental diagnoses (Ramos et al., 2020; Smith-Spark & Fisk, 2007), including attention deficit hyperactivity disorder (ADHD; Ramos et al., 2020). ADHD is characterised by elevated levels of inattention and/or hyperactivity, and affects over 7% of children worldwide (Thomas et al., 2015), with symptoms often persisting into adulthood (Faraone et al., 2006; Kessler et al., 2006; Sibley et al., 2016; 2017).

It has been suggested that working memory may be a key area of cognitive dysfunction in ADHD (Ramos et al., 2020), and potentially a core deficit of the disorder (Kofler et al., 2008; Rapport et al., 2009). Supporting this, clear difficulties in working memory tasks have been observed, both during development and in adulthood (e.g. Alderson et al., 2013; Kasper et al. 2012; Martinussen et al., 2005; Ramos et al., 2020; Yang et al., 2017). For example, a metaanalysis investigating working memory difficulties in children with ADHD found moderatelarge impairments in both spatial and verbal working memory, with difficulties particularly prevalent with storage and processing of spatial information (Martinussen et al., 2005). Furthermore, a meta-analytic review examining working memory difficulties in adults with ADHD found impairments in both phonological and visuo-spatial working memory (Alderson et al., 2013). Taken together, these studies provide clear evidence of difficulties with working memory tasks in individuals with ADHD, which may persist into adulthood (Martinussen et al., 2005; Alderson et al., 2013). There is also evidence that difficulties in working memory in individuals with ADHD may be more associated with key outcomes than the inattentive and hyperactive symptoms themselves (e.g. Simone et al., 2018). For example, Simone et al. (2018) found that working memory difficulties in children with ADHD were more predictive of academic achievement than their symptoms of ADHD.

One key question of interest in neurotypical populations in recent years has been the extent to which individuals can direct their attention to particularly important or goal-relevant information in working memory (see Allen et al., 2024, Souza & Oberauer, 2016, for reviews). This question is of theoretical interest, as it reflects the extent to which individuals can use their limited working memory resources to prioritize particularly relevant information. It is also of practical interest, potentially revealing a novel approach to ensuring that important information is retained. One key approach used to examine this has been to manipulate an item's value, such that participants gain more notional points for recall of some items relative to others. In this paradigm, participants are typically presented with sets of coloured shapes to recall after a brief delay. Prior to encoding, participants are informed how many points an item is worth if they are asked about that item and they respond correctly. In some trials, they may be told that all items are equally valuable (e.g. worth 1 points). In other trials, participants may be told that correct recall of a particular item will gain them more points than the rest of the items (e.g. 4 points for the high value item, vs 1 point for the low value items). Using this approach, it has been demonstrated that individuals can prioritize particularly valuable information in working memory, as evidenced by enhanced recall for higher value items (e.g. Atkinson et al., 2018; Allen & Ueno, 2018; Hitch et al., 2018). This boost does, however, typically come at a cost to the less valuable items presented in the same trial (e.g. Allen et al. 2024; Allen & Atkinson, 2019; Atkinson et al., 2018; 2024). Thus, prioritization instructions do not increase overall working memory capacity, but instead result in a re-allocation of resources towards the high value information (e.g. Allen et al., 2024; Atkinson et al. 2024; Hitch et al. 2018).

The prioritization effect is considered to result from the high value item being stored in an active and accessible privileged state, termed the focus of attention (e.g. Allen et al., 2024; Hu et al., 2014). The effect is robust and has been demonstrated across a wide variety of task contexts, including across different modalities (e.g. visual, verbal, tactile, and cross-modal information; Atkinson et al., 2021; Roe et al., 2024), presentation contexts (sequential and simultaneous presentation; Allen & Ueno, 2018; Atkinson et al., 2018; Hitch et al., 2018; Hu et al., 2014; 2016), and retrieval methods (cued-recall, recognition, and color reproduction tasks; Atkinson et al., 2022; 2024; Sandry et al., 2014; 2020). Whilst the majority of research has focused on testing young adults, prioritization effects have also been observed in older adults (Allen et al., 2021) and children (provided they are sufficiently motivated; Atkinson et al., 2019).

Studies to date have focused on neurotypical individuals. As such, it is not clear whether individuals with symptoms of ADHD, who often experience working memory difficulties, would be able to utilise the focus of attention within working memory to prioritize particularly valuable information. One possibility is that individuals with symptoms of ADHD may be less able to prioritize particularly valuable information in working memory. Supporting this possibility, there is some evidence that individuals with ADHD may have difficulties in orienting attentional resources during encoding (e.g. Kim et al., 2014). Moreover, Castel et al. (2011) found that some individuals with ADHD show less selectivity for valuable information during a long-term memory task. In this study, children with ADHD-Combined type (exhibiting difficulties with inattention and hyperactivity), ADHD-Inattentive type (exhibiting difficulties with inattention), and controls took part in a memory task. They were shown lists of 12 words to be remembered. Each word from the list was shown sequentially and paired with a number ranging from 1-12, indicating the number of points that it was worth. After each list had been presented, participants were asked to recall as many words from the list as possible within 30 seconds. They were told that the points collected could be exchanged for prizes at the end of the experiment. All groups reported more high value words relative to low value words. However, there was some evidence that children in the ADHD-Combined group were less selective relative to the other two groups. As such, it is possible that individuals with symptoms of ADHD would be less able to prioritize particularly valuable information in working memory.

A second possibility is that individuals with symptoms of ADHD may be able to direct attention in working memory as effectively as controls. Evidence for this is provided by Superbia-Guimarães et al. (2022), who tested 10-16-year-olds with ADHD and controls on a working memory task where participants had to recognise the colour of animal drawings. Participants either received no location cues, were shown a pre-location cue (before the shapes appeared, an arrow pointed towards the location of the item that would later be tested), or a retro-cue (an arrow pointed at the relevant stimuli that would be tested after the animal shapes were shown). In this study, the ADHD and control groups benefitted equally from the pre-and retro-cues. Based on this it might be predicted that individuals with symptoms of ADHD will be able to direct their attention in working memory on the basis of reward as effectively as controls.

Finally, it is possible that individuals with symptoms of ADHD may show a larger prioritization boost relative to the controls. Indeed, there is some evidence that cognitive performance in individuals with ADHD may be more influenced by reward relative to matched controls (Dovis et al., 2012). Dovis et al. examined how motivational factors (such as monetary incentives) affect overall working memory performance of children with ADHD and controls. In their study, participants completed a visuo-spatial working memory task, under four conditions: feedback only, 1 euro, 10 euro, and 'game'. In the feedback only condition, participants received only feedback about their performance. In the 1 euro and 10euro conditions, participants were informed they could earn this amount of money if they performed well enough. The task was gamified in the 'game' condition, with participants completing the same working memory task to progress a robot through a storyline. Children with ADHD performed better in the 1 euro, 10 euros, and game conditions relative to the feedback only condition. In contrast, there were no significant differences between the incentive conditions in controls. The incentives did not "normalize" performance in the ADHD group, however, with performance in the ADHD group significantly worse than the control group across all conditions. Taken together, this demonstrates that the effect of rewards can be greater for individuals with ADHD, but that the effects are not large enough to remove the working memory difficulties experienced in this group relative to controls. Nevertheless, based on these findings, it might be suggested that individuals with symptoms of ADHD will show a larger prioritization effect relative to controls.

We conducted a pair of experiments to investigate this. Participants were presented with sets of four coloured shapes presented sequentially. After a short delay, the outline of each shape was presented (in a counterbalanced order), and participants were asked to report the colour of the shape using coloured buttons on screen. Participants completed two blocks of trials. In one block (equal value condition), they were told that all items were equally valuable, and that correct recall of each item would gain them two points. In the other block (differential value condition), participants were told that correct recall of the first item presented during encoding would gain them five points, whilst correct recall of the other items would gain them one point. In Experiment 1, feedback was presented on a trial-by-trial basis, informing participants which items they had responded correctly about, the number of points collected for each item, and the number of points collected in test trials overall so far. In Experiment 2, feedback was not presented, as is more typical in this paradigm (e.g. Allen & Ueno, 2018; Atkinson et al., 2018, 2021; 2022 Hu et al., 2014, 2016). Whilst participants completed the

experiments, they engaged in articulatory suppression to disrupt verbal recoding of the visual information (Baddeley, 1986). This was to ensure that the task was indeed measuring visual working memory as intended, instead of participants also utilising verbal working memory to retain the information. The experiments were conducted online, with participants recruited via Prolific. Participants in the ADHD-symptoms group considered themselves to have ADHD and reported symptoms consistent with a diagnosis on the Adult ADHD Self-Report Scale (ASRS; v1.1) Screener (e.g. Kessler et al., 2007). Meanwhile, those in the control group did not consider themselves to have ADHD and did not show symptoms consistent with a diagnosis on the ASRS Screener.

Experiment 1

Experiment 1 investigated whether the ability to prioritize particularly valuable information in working memory differs in individuals with ADHD relative to controls without ADHD. A 2 (Value: Differential vs Equal; within-subject) x 4 (Serial position (SP): 1-4; within-subject) x 2 (Group: ADHD-symptoms vs controls; between-subjects) mixed-design was employed. As in previous research, the value manipulation was targeted towards the first item (SP1; e.g. Atkinson et al., 2018; 2019). At SP1, it was expected that a significant effect of value would emerge, with participants exhibiting higher accuracy in the differential condition relative to the equal value condition. This would indicate that participants perform better at SP1 when this item is relatively more valuable than the other items in the trial, compared to a condition in which all items were as valuable as each other. Of particular interest was whether the value effect differed between groups. In particular, whether individuals with symptoms of ADHD would show increased, decreased, or similar sized prioritization effects relative to the control participants. At the less valuable SPs, it was expected that performance would be superior in the equal value condition relative to the differential value condition. This would indicate that prioritization of the particularly valuable item (SP1) comes at a cost to the other, less valuable serial positions (e.g. Allen et al. 2024; Allen & Atkinson, 2021; Atkinson et al., 2018; 2024). Again, of particular interest was whether the costs differed as a function of group.

Method

Participants

Power analysis was conducted using G*Power (Faul et al., 2007). This was calculated for an ANOVA including value (within-subjects: differential vs equal) and group (between-subjects: ADHD-symptoms vs control) at SP1, since this is where the value manipulation was targeted. Based on a medium effect size of $\eta_p^2 = .06$ and alpha = 0.05, it was estimated that a total sample size of 54 participants across groups would provide 95% power. This sample size would also provide 95% power for an equivalent ANOVA (e.g. 2 (value [within-subjects]: differential vs equal) x 2 (group [between-subjects]: ADHD-symptoms vs control) on data averaged across the less valuable SPs (SPs 2-4).

All participants were recruited via Prolific. Eighty participants completed the experiment in total: 40 who considered themselves to have ADHD (ADHD-symptoms group) and 40 controls who did not consider themselves to have ADHD (control group). As we were interested in participants with symptoms of ADHD vs controls, we administered the ASRS screening measure for ADHD and excluded participants at the analysis stage whose degree of inattentive and hyperactive symptoms was inconsistent with the category to which they were assigned (i.e. we excluded participants in the ADHD-symptoms group who did not exhibit symptoms consistent with ADHD and participants in the control group who exhibited symptoms consistent with ADHD). Therefore, the ADHD-symptoms group reflects participants who consider themselves to have ADD/ADHD and have symptoms consistent with this diagnosis, whilst the Control group reflects participants who do not consider themselves to have ADD/ADHD and do not have symptoms consistent with an ADHD diagnosis (for a more detailed description of the recruitment approach, please see the supplementary materials). On this basis, three participants in the ADHD-symptoms group were excluded for scoring below the cut-off for ADHD on the ASRS screening measure, suggesting they do not exhibit behaviours highly consistent with ADHD. Meanwhile, eight participants in the control group were excluded for scoring above the cut-off for ADHD on the ASRS, suggesting these individuals were exhibiting behaviours highly consistent with ADHD. One control participant was then excluded due to a technical error with the audio recording, which prevented the articulatory suppression check being completed. Further, one control participant and one participant in the ADHD-symptoms group were excluded for not completing the articulatory suppression task as instructed. Finally, one control participant was excluded as their average performance across the task was below the chance guessing rate. The final sample therefore comprised 65 participants overall (36 in the ADHD-

symptoms group: 17 female; 18 male; 1 non-binary; Mean (M) age = 27.92 years; standard deviation (SD) = 4.46; and 29 participants in the control group: 11 female; 18 male; M. age = 29.38, SD = 4.09). The groups did not significantly differ in terms of gender (p = .526) or age (t(63)=-1.36, p = .177). Sixteen participants in the ADHD-symptoms group had received a formal diagnosis of ADHD, whilst 19 had not and one preferred not to say. Seven participants in the ADHD-symptoms group reported taking medication for ADHD, whilst 29 responded that they were not taking medication for ADHD. All participants were recruited at a similar time of day and met the following criteria: 18-35 years of age, and had normal or corrected-to normal vision, no colour blindness, English as a first language, no formal diagnosis of Autism or Dyslexia, resided in the UK, and an approval rate of 95% or higher on Prolific. Participants were paid £5 for their participation (a rate equivalent to £10 per hour).

As expected, there was a significant difference between groups in the mean total ASRS Screener score (t(63) = 15.38, p < .001; range possible = 0-24), with the score higher in the ADHD-symptoms group (M = 18.36, SD = 2.54, Range = 14-24) than the control group (M = 8.21, SD = 2.77, Range = 1-13). The inattentive score (range possible: 0-16) was significantly higher in the ADHD-symptoms group (M = 12.39, SD = 2.03, Range = 9-16) than controls (M = 5.72, SD = 2.00 Range = 1-9; t(63) = 13.24, p < .001). The hyperactivity score (range possible: 0-8) was also significantly higher in the ADHD-symptoms group (M = 5.97, SD = 1.21, Range = 3-8) than the control group (M = 2.48, SD = 1.57, Range = 0-5; t(63) = 10.12, p < .001).

Materials

ASRS Screener

This comprises six items, designed to measure inattentive and hyperactive behaviours (e.g. delaying starting tasks, fidgeting). Participants report the extent to which they believe they have displayed each behaviour in the last six months, selecting from the following options: "Never", "Rarely", "Sometimes", "Often", "Very Often". In line with our ethical approval, we also included a "Prefer not to say" option (although this option was not selected by any participants). The ASRS Screener is a commonly used tool, with high diagnostic accuracy internal consistency, and good test-re-test reliability (Brevik et al., 2020; Kessler et al., 2005; 2007; Lewczuk et al., 2024; Matza et al., 2010). Moreover, the 6-item ASRS screener used

here is considered to perform at least as well as the original 18-item ASRS (Brevik et al., 2020; Kessler et al., 2005). In line with the ASRS Screener guidance, participants received a mark for each item if they responded: (i) "Sometimes", "Often" or "Very Often" for the first three items; and (ii) "Often" or "Very Often" to the final three items. Participants were recorded as scoring above the cut-off if they received four or more marks across the six items. This indicated the presence of attentional and hyperactive symptoms highly consistent with a diagnosis of ADHD. Meanwhile, participants were recorded as having not met the cut-off for ADHD if they received three or fewer marks. A total ASRS Screener score was also calculated per participant (Kessler et al., 2007). For each of the six statements, a numerical score was given (0 for "Never"; 1 for "Rarely"; "2" for "Sometimes"; 3 for "Often"; and 4 for "Very Often"). An overall score was calculated (minimum = 0, maximum = 24) which reflects the degree of inattentive and hyperactive symptoms (Kessler et al., 2007). Scores were also calculated for the total inattentive score (minimum = 0, maximum = 16) and the total hyperactivity score (minimum = 0, maximum = 8). For psychometric properties of the approaches used, we refer readers to Kessler et al. (2007).

Working memory task

In the working memory task, four items were presented sequentially in each trial. Stimuli were formed by pairing one of six shapes (circle, diamond, triangle, arrow, cross, flag) with one of six colours (red, blue, green, yellow, magenta, black). Pairings were formed with the constraint that no colour or shape could be repeated within the same trial. Shapes (sized to fit within an imaginary 2cm square) were presented on a white background. Locations were fixed, such that the items moved across the top of the screen from left to right (Allen et al., 2021). The four shapes were spaced equally about the horizontal centre of the screen, separated by 4.3cm horizontally, and located 1.8cm above the vertical centre of the screen. All items were tested on each trial in a counterbalanced order. The test probe was an outline of one of the shapes presented in the lower half of the screen with coloured buttons underneath. The response buttons always appeared in the same order in a horizontal line (from left to right): red, blue, green, yellow, magenta, black.

Design and procedure

The study employed a 2 (value: differential vs equal; within-subject) x 4 (SP: 1-4; within-subject) x 2 (Group: ADHD-symptoms vs control; between-subject) mixed design. The

dependent variable was accuracy (determined by the mean proportion of trials in which participants responded accurately). Participants completed two blocks of 24 trials: one for each 'value' condition. The order of blocks was randomised. All SPs were tested on every trial, with the order of probes being counterbalanced within the value blocks and presented to participants in a random order. Thus, each SP was tested first, second, third, and fourth six times within each block. At the start of each block, participants completed four practice trials.

At the start of the experiment, participants completed a short questionnaire to collect demographic information, including age, gender, whether they have a formal diagnosis of ADHD, and whether they take any medication for ADHD. As part of this questionnaire, participants also completed the ASRS Screener (e.g. Kessler et al., 2007). There was then a calibration phase to ensure that the participants' microphone was working correctly. In addition, participants were asked to re-size a rectangle presented on screen until it was the same size as a credit card. This enabled stimuli to be scaled based on the participants' screen size. This is a standard approach which ensures that stimuli are presented at the same size across participants when conducting online experiments (Gresch et al., 2021; 2022; Li et al., 2020).

Each condition began with written instructions accompanied by pictorial representations. Participants were first given general instructions about the task. They were then asked to watch a short video of an example trial, which included a recording of the articulatory suppression at the correct rate (once per second). They were then told the point values associated with each item. In the differential value trial, participants were told that the first shape presented was worth 5 points and all other shapes were worth 1 point. In the equal value trial, they were told that each shape was worth 2 points. This reflected the number of points participants would collect if they responded correctly about each item. In each condition, participants were presented with an example set of shapes, which indicated how many points they would collect for each item. Points were notional and did not equate to any physical reward or monetary bonus. To ensure that instructions had been read and understood, participants were then presented with a visual example of stimuli and asked how many points each item was worth (see Figure 1). They needed to respond correctly about this to proceed. An incorrect response led to repetition of the instructions and the test question. This repeated until participants responded correctly.

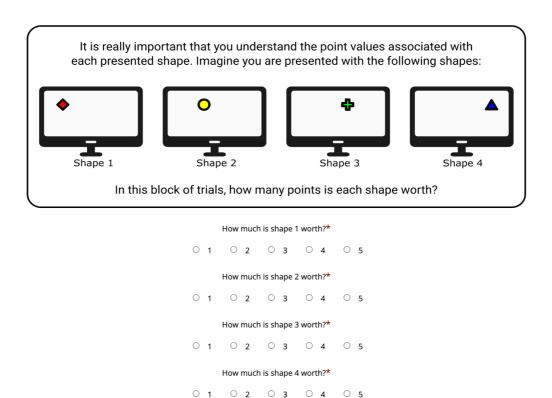


Figure 1. An example of how participants were required to demonstrate they had read and understood the instructions.

The mean number of attempts the ADHD-symptoms group took to respond correctly was 1.00 in the equal condition (SD = 0.00) and 1.08 (SD = 0.28) in the differential value condition. Meanwhile, the mean number of times participants in the control group took to respond correctly was 1.10 (SD = 0.31) in the equal value condition and 1.17 (SD = 0.47) in the differential value condition (see supplementary material for frequencies). This did not significantly differ in either value condition (equal condition p = .084; differential condition p = .666).

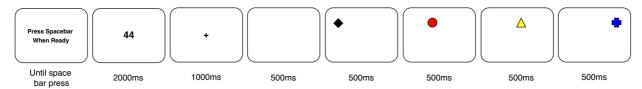
The experimental paradigm is displayed in Figure 2. Each trial began with the on-screen message "Press Spacebar When Ready", which remained on screen until participants pressed the spacebar. Next, a two-digit number pseudo-randomly selected between 20-99 was displayed on screen for 2000ms. Participants were asked to repeat this number aloud at a rate of one repetition a second to disrupt verbal rehearsal (Baddeley, 1986). This was followed by a fixation cross for 1000ms. After a blank interval of 500ms, each shape was presented individually for 500ms. There was then a 1000ms blank interval before all items were tested individually. At the test phase, each test probe remained on screen until the participant

responded. After the participant had responded about all shapes, there was a 500ms blank interval, before feedback was presented for 2000ms. This feedback screen informed participants which shapes they responded correctly about. The shapes were displayed in their correct colours in a horizontal line at the centre of the screen in the same order as they were presented during the encoding phase. When the colour of shapes had been correctly recalled, there was a green plus sign and number underneath (e.g. "+5"). This number corresponded to the number of points that participants had collected. A green tick was then presented underneath the number. Nothing was presented underneath shapes to which participants responded incorrectly. Underneath this was a running total, indicating the number of points participants had collected so far (e.g. "Your score so far: 25"). In practice blocks, this score started from zero and was self-contained, so did not contribute towards their overall score. Their overall score started from zero in the first test block, but carried over to the second test block, so that their performance in all test trials contributed towards their score at the end of the experiment.

To ensure compliance with the articulatory suppression instruction in the online testing environment, audio was recorded from the number being displayed until the end of the maintenance phase. Participants were told that their submission on Prolific would be rejected (i.e. they would not receive payment) if they did not engage in the articulatory suppression task as instructed. However, in reality, all participants were paid. Audio recordings were processed offline after the experiment to check compliance with the instruction.

Participants were given an untimed break halfway through each test block. The median time to complete the experiment was approximately 30 minutes.

Encoding



Retention interval: blank screen for 1000ms

Test and feedback

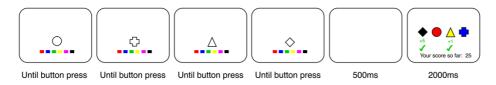


Figure 2. The experimental paradigm used in Experiment 1, with a differential value trial as an illustrative example. Figure not to scale.

Data analysis

Data for both experiments and the analysis scripts are available on the Open Science Framework (OSF): https://osf.io/y3zw9/

The main outcome measure was accuracy, determined as the proportion of correct responses. Frequentist analysis was conducted in R (R Core Team, 2023), using the afex (Singmann et al., 2023) and emmeans (Lenth, 2023) packages. Post-hoc comparisons for the frequentist ANOVAs were corrected using Bonferroni-Holm. Bayes Factor (BF) analysis was also conducted. This indicates the strength of evidence for the presence or absence of an effect. Bayesian ANOVAs were conducted using the 'BayesFactor' package (Morey & Rouder, 2022) in R (R Core Team, 2023). Default priors were used, with 500,000 iterations. All models were calculated, meaning that a model could contain an interaction without the constituent main effects. In addition to reporting the best model, we report BFs for the individual main effects and interactions. These Bayes factors were computed by re-running the model with which model set to 'top'. This compares a model that omits a main effect/interaction to the model containing all main effects and interactions. This produces BF_{01} values, which indicates evidence of no effect. BF_{10} values were derived by inverting the (BF_{01}) values $(1/BF_{01})$. When BF_{10} is below 1, BF_{01} are also reported for ease of interpretation. BF_{10} above 3 were taken as evidence for an effect, whereas a BF_{10} below 0.33 (i.e. BF_{0l} above 3) was taken as evidence of no effect. We primarily draw conclusions

based on p-values, but we draw readers' attention to any discrepancies that would result from interpreting p-values versus BFs.

The main analysis comprised a 2 (value; differential vs equal; within-subject) x 2 (group: ADHD-symptoms vs control; between-subject) mixed ANOVA at the targeted SP (SP1). A 2 (value; differential vs equal; within-subject) x 2 (group: ADHD-symptoms vs control; between-subject) mixed ANOVA was also conducted on data averaged across the less valuable SPs (2-4).

Subsidiary analysis was also conducted to examine: (i) effects after excluding participants who reported taking medication for ADHD; and (ii) whether effects differ in individuals with and without a formal diagnosis of ADHD (see supplementary materials). To summarise, the conclusions obtained in all of these analyses did not differ from those reported in the main body below. We also conducted a three-way ANOVA (a 2 (value; differential vs equal; within-subject) x 4 (SP: 1-4) x 2 (group: ADHD-symptoms vs control; between-subject) ANOVA), to examine performance across all SPs within the same analysis and where costs to the less valuable SPs lie (see supplementary materials). We note that these outcomes should be interpreted with caution, however, as our power calculation was based on the main 2 (value; differential vs equal; within-subject) x 2 (group: ADHD-symptoms vs control; between-subject) mixed ANOVAs.

Finally, exploratory analysis was conducted to investigate whether there was a correlation between the ASRS Screener score, the prioritization boost at SP1 (i.e. performance in the differential condition at SP1 minus performance in the equal value condition), and the cost at the less valuable SPs (i.e. performance in the differential condition minus performance in the equal value condition). This analysis used the total ASRS Screener score, and thus reflected both inattentive and hyperactive behaviours. Higher scores on the ASRS reflect more symptoms of ADHD.

Results

Effect at SP1 (targeted SP)

Mean proportion correct at SP1 is displayed in Figure 3A, as a function of value and group. Meanwhile, Figure 3B displays the difference between the differential and equal value conditions as a function of group. A 2 (Value: differential vs equal; within-subject) x 2 (Group: ADHD-symptoms vs control; between-subjects) mixed ANOVA revealed a main effect of value at SP1 (F(1, 63) = 93.63, MSE = 0.03, p < .001, $\eta_p^2 = .60$; $BF_{10} > 10,000$) whereby participants exhibited higher performance in the differential value condition (M = 0.77, SE = 0.02) relative to the equal value condition (M = 0.49, SE = 0.03). There was no effect of group (F(1, 63) = 2.40, MSE = 0.04, p = .127, $\eta_p^2 = .04$; $BF_{10} = 0.64$, $BF_{01} = 1.56$), and no interaction between value and group (F(1, 63) = 1.91, MSE = 0.03, p = .172, $\eta_p^2 = .03$; $BF_{10} = 0.60$, $BF_{01} = 1.66$). We note that although the BF evidence for the interaction is not strongly in favour of no effect, the pattern (see Figure 3A and 3B) indicates a numerically larger effect in the ADHD-symptoms group than the control group. There is thus no evidence that the ability to prioritize valuable information is impaired in the ADHD-symptoms group relative to the control group. The BF analysis revealed the best model included a main effect of value ($BF_{10} > 10,000$ relative to the model containing only participant).

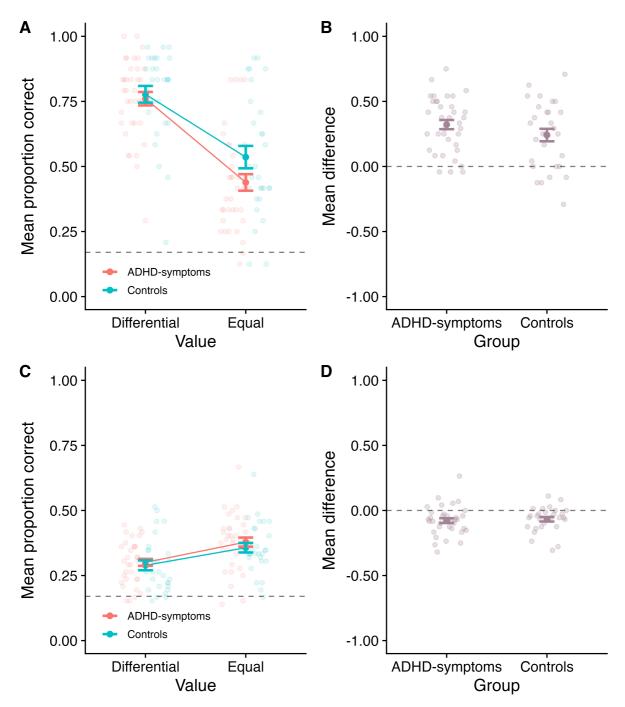


Figure 3. Outcomes for Experiment 1. Panel A: Mean proportion correct at SP1 as a function of value and group. Panel B: The mean difference between the differential and equal value conditions at SP1, as a function of group (calculating as Differential – Equal). Panel C: Mean proportion correct averaged across the less valuable SPs (SPs 2-4), as a function of value and group. Panel D: The mean difference between the differential and equal value conditions averaged across the less valuable SPs, as a function of group (calculating as Differential – Equal). In all panels, the bold point reflects the mean across participants, whilst error bars display standard error. The lighter points in the background reflect the means for individual participants. In Panels A and C, the dotted line at 0.17 reflects chance guessing rate, based on a stimulus set of six items. In Panels B and D, the dotted line at 0.00 reflects no difference.

Effects averaged across less valuable SPs (2-4)

Mean proportion correct averaged across the less valuable SPs (2-4) is displayed in Figure 3C, as a function of value and group. Figure 3D shows the mean difference between the Differential and Equal value conditions as a function of group. A 2 (Value: differential vs equal; within-subject) x 2 (Group: ADHD-symptoms vs control; between-subjects) mixed ANOVA was conducted. There was a main effect of value (F(1, 63) = 31.50, MSE = 0.01, p < .001, $\eta_p^2 = .33$; $BF_{10} > 10,000$), with participants displaying better performance in the equal value condition (M = 0.37, SE = 0.01) relative to the differential value condition (M = 0.30, SE = 0.01). There was no effect of group (F(1, 63) = 0.66, MSE = 0.01, p = .421, $\eta_p^2 = .01$; $BF_{10} = 0.37$, $BF_{01} = 2.73$). There was also no interaction between value and group (F(1, 63) = 0.16, MSE = 0.01, p = .690, $\eta_p^2 < .01$; $BF_{10} = 0.28$, $BF_{01} = 3.63$). The BF analysis revealed the best model included a main effect of value ($BF_{10} > 10,000$ relative to the model containing only participant).

Correlations between the degree of inattentive and hyperactive symptoms and prioritization boosts and costs

Exploratory analysis was then conducted to investigate whether the total ASRS Screener score (reflecting the degree of inattentive and hyperactive symptoms) correlates with the prioritization boost to SP1 and the cost to the other SPs. Figure 4 displays scatterplots of ASRS Screener scores, boosts to the high value item, and costs to the low value items.

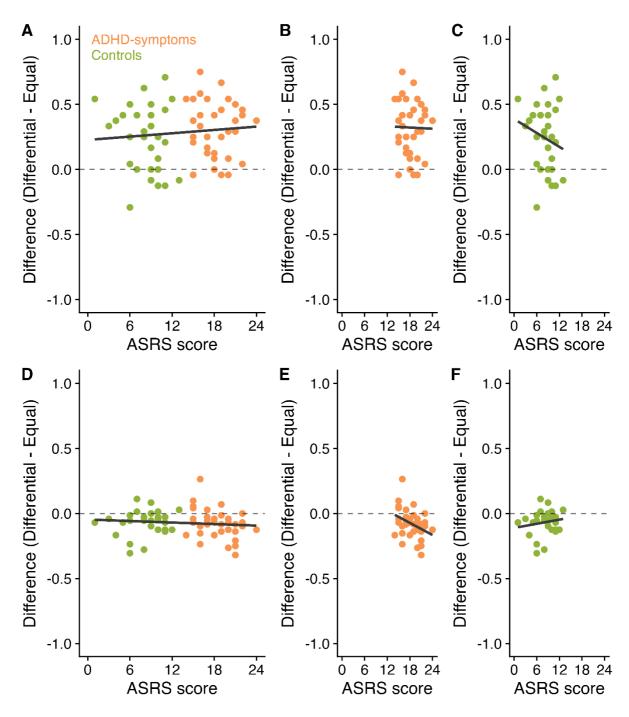


Figure 4. Correlations between ASRS Screener scores (reflecting the degree of inattentive and hyperactive behaviours), boosts to SP1 and costs to the less valuable items in Experiment 1. Panel A displays the correlation between ASRS Screener score and the difference in performance between the differential and equal value items at SP1 across participants. Panels B and C then present these correlations for the ADHD-symptoms and control groups, respectively. Panel D displays the correlation between ASRS Screener score and the difference in performance between the differential and equal value items at the less valuable SPs. Panels E and F then present these correlations for the ADHD-symptoms and control groups, respectively.

Pearson's correlation coefficients were conducted to investigate correlations between the ASRS Screener score (including all participants), the boost at SP1 and the costs at the less valuable SPs. Analysis was also conducted for each group separately. P-values were corrected using Bonferroni-Holm. Following correction, there was no significant correlation between any of the measures. In most cases, BF were at least anecdotally in favour of no correlations.¹

Table 1. Correlations between the ASRS score, the prioritization boost at SP1 and the prioritization costs at the less valuable SPs (2-4) in each experiment.

Variables	r	p	BF_{I0}	BF_{01}
Experiment 1				
ASRS score and prioritization boost at SP1 (all participants)	0.10	1.000	0.39	2.59
ASRS score and prioritization cost at less valuable SPs (all participants)	-0.11	1.000	0.39	2.53
ASRS score and prioritization boost at SP1 (ADHD-symptoms group)	-0.02	1.000	0.37	2.69
ASRS score and prioritization cost at less valuable SPs (ADHD-symptoms group)	-0.36	0.197	2.68	0.37
ASRS score and prioritization boost at SP1 (Control group)	-0.19	1.000	0.62	1.63
ASRS score and prioritization cost at less valuable SPs (Control group)	0.16	1.000	0.53	1.87
Experiment 2				
ASRS score and prioritization boost at SP1 (all participants)	0.05	1.000	0.30	3.39
ASRS score and prioritization cost at less valuable SPs (all participants)	< 0.01	1.000	0.28	3.62
ASRS score and prioritization boost at SP1 (ADHD-symptoms group)	0.21	1.000	0.70	1.43
ASRS score and prioritization cost at less valuable SPs (ADHD-symptoms group)	-0.26	0.813	0.99	1.01
ASRS score and prioritization boost at SP1 (Control group)	0.13	1.000	0.47	2.13
ASRS score and prioritization cost at less valuable SPs (Control group)	-0.08	1.000	0.42	2.40

Note: p-values corrected using Bonferroni-Holm within each experiment

Discussion

This experiment was the first to investigate the extent to which individuals with symptoms of ADHD can direct their attention in working memory on the basis of value, and to compare this to a neurotypical group. Across groups, a prioritization boost was observed, whereby

¹ There were also no significant correlations between the ASRS, boosts at SP1 and costs to the less valuable items following Bonferroni-Holm correction when measures of inattention and hyperactivity were analysed separately.

participants responded more accurately at the first SP when this item was associated with more notional points. Significant costs to the less valuable SPs were also observed. Importantly, there was no evidence that either the boost to SP1 or the costs to the less valuable SPs differed as a function of group. This provides the first evidence that individuals with symptoms of ADHD can prioritize information based on value in working memory, and that this ability is not impaired relative to controls.

These findings are broadly in line with those of Superbia-Guimarães et al (2022) who found that 10–16-year-olds with ADHD can direct their attention in working memory based on precues and retro-cues as effectively as controls. The current findings extend this, demonstrating similar findings in adults and using an alternative type of attentional direction, considered to be distinct to predictive cue-based prioritization (Allen et al., 2024; Atkinson et al., 2018; Hitch et al., 2018). These observations run counter to existing findings suggesting that individuals with ADHD have difficulty allocating attention (Kim et al., 2014). Moreover, although the prioritization boost at SP1 was numerically larger in the ADHD-symptoms group, this was not supported by the inferential analyses. Therefore, these findings also do not align with empirical research suggesting that reward affects working memory performance to a greater extent in individuals with ADHD than controls (Dovis et al., 2012).

Evidence of costs of prioritization to less valuable items were observed. This fits with a large body of literature in neurotypical populations (e.g. Allen et al. 2024; Allen & Atkinson, 2019; Atkinson et al., 2018; 2024), suggesting that directing attention based on reward in working memory does not increase overall performance. Instead, it appears to result in a re-allocation of limited resources, whereby participants direct attention away from less valuable items (e.g. Allen et al., 2024; Atkinson et al. 2024; Hitch et al. 2018). As indicated in Figure 3 and the inferential analyses, the costs observed to the lower value items did not appear to differ between groups. This provides evidence that in individuals with symptoms of ADHD, prioritization similarly does not enhance overall capacity but instead results in a re-allocation of resources.

In the current experiment, participants were given feedback after every trial, informing them which items they responded correctly about, and the number of points collected for each item. This contrasts with much of the value-based prioritization literature, in which participants do not typically receive feedback at all (e.g. Allen et al., 2021; Allen & Ueno,

2018; Atkinson et al., 2018; 2021; 2022; Hu et al., 2014; 2016). The value effect observed in both groups in this experiment was considerably larger than that typically observed in previous studies, which may have resulted from the regular reinforcement of the point differences induced by the provision of feedback.

However, it is possible that the combination of value and feedback may not have impacted the groups to the same extent. Indeed, previous research has suggested that a combination of reward and feedback may particularly enhance working memory performance in individuals with ADHD (Hammer et al., 2015). In this study, boys with ADHD (mean age = 10.5 years) and age-matched controls (mean age = 10.9 years) completed a visual 2-back task, in which they had to decide whether a letter presented on screen was in the same location as one presented two letters ago. The provision of trial-by-trial feedback (feedback vs no feedback) and reward (large/small) was manipulated. In the large reward condition, the reward for correct responses was 10x larger than in the small reward condition. In the ADHD group, participants performed significantly better in the large-reward feedback condition relative to all other conditions. In contrast, controls performed significantly better in the large-reward no feedback condition relative to all other conditions. This suggests that the combination of feedback and reward may have been particularly beneficial for individuals with ADHD. Further supporting this, controls performed significantly better than the ADHD group in all conditions except the large-reward feedback condition, and the degree of ADHD symptoms predicted performance in all conditions except this one. Although this study involved offering rewards for task completion, rather than encouraging participants to direct attention towards particular items, it is possible that the absence of any group effects observed in the current experiment reflects the combination of value and trial-by-trial feedback employed. Thus, one possibility is that prioritization within working memory may be impaired in individuals with symptoms of ADHD relative to controls, when feedback is not provided (as is more typical in the value-based prioritization literature; Allen et al., 2021; Allen & Ueno, 2018; Atkinson et al., 2018; 2021; 2022; Hu et al., 2014; 2016). This was examined in Experiment 2.

Experiment 2

Experiment 2 investigated whether adults with symptoms of ADHD would be able to direct attention in working memory based on reward as effectively as controls, when feedback was not provided. This is the paradigm more typically used when investigating this ability in

neurotypical populations (Allen et al., 2021; Allen & Ueno, 2018; Atkinson et al., 2018; 2021; 2022; Hu et al., 2014; 2016). Participants completed the same visual working memory task as in Experiment 1, except that feedback was not provided at the end of each trial. Based on the outcomes from Experiment 1, it may be predicted that a value-based prioritization effect would be observed in the ADHD-symptoms group, with the magnitude of the effect a similar size to the control group. However, based on Hammer et al. (2015) it may be predicted that the prioritization effect may be smaller when trial-by-trial feedback is not provided.

Method

Participants

The same recruitment approach was taken as in Experiment 1. The experiment was completed by 80 participants: 40 who consider themselves to have ADHD (ADHDsymptoms group) and 40 controls who do not consider themselves to have ADHD (control group). As in Experiment 1, participants where then excluded if their degree of inattentive and hyperactive symptoms (reported on the ASRS Screener) was inconsistent with the group to which they were assigned. On this basis, four participants in the control group were excluded for scoring above the cut-off for ADHD on the ASRS Screener, whilst two participants in the ADHD-symptoms group were excluded for scoring below the cut-off for ADHD on the ASRS Screener. A further four participants were excluded due to technical errors with the audio recording (three in the ADHD-symptoms group and one in the control group) which prevented the articulatory suppression from being checked. Finally, two participants (one in the ADHD-symptoms group and one in the control group) were excluded due to a failure to engage in the articulatory suppression task as instructed. No performancebased exclusions were applied, as all participants performed above chance guessing level. The final analysis was therefore performed on 34 individuals with symptoms of ADHD (19 female; 13 male; 2 non-binary, M. age = 26.35, SD = 4.17; M. years in education = 16.32; SD= 2.64) and 34 controls (14 female, 20 male; M. age = 29.97, SD = 3.97, M. years in education = 17.32; SD = 3.72). The groups did not significantly differ on gender (p = .132) or years in education (t(66)=-1.28, p=.206), but the ADHD-symptoms group were significantly younger than the control group (t(66)=-3.67, p<.001). In the ADHD-symptoms group, 11 had received a formal diagnosis of ADHD, 22 had no received a formal diagnosis, and one

preferred not to say. Moreover, five participants in the ADHD-symptoms group reported taking medication for ADHD. As in Experiment 1, participants were recruited through Prolific, paid £5 for their participation (equivalent to £10/hour), and met the following criteria: 18-35 years of age, had normal or corrected-to normal vision, no colour blindness, English as a first language, no formal diagnosis of Autism or Dyslexia, and resided in the UK. None of the participants had taken part in Experiment 1.

As expected, there was a significant difference between groups in the mean total ASRS Screener score (t(66) = 15.27, p < .001; range possible = 0-24), with the score higher in the ADHD-symptoms group (M = 18.41, SD = 2.75, Range = 13-24) than the control group (M = 7.12, SD = 3.32, Range = 0-12). The inattentive score (range possible: 0-16) was significantly higher in the ADHD-symptoms group (M = 12.21, SD = 2.06 Range = 9-16) than controls (M = 4.82, SD = 2.34, Range = 0-9, t(66) = 13.81, p < .001). The hyperactivity score (range possible: 0-8) was also significantly higher in the ADHD-symptoms group (M = 6.21, SD = 1.27, Range = 4-8) than the control group (M = 2.29, SD = 1.40, Range = 0-5; t(66) = 12.03, p < .001).

Materials, design and procedure

The materials, design and procedure were identical to Experiment 1, except that trial-by-trial feedback was not provided. To retain the same interval between retrieval and the initiation of the next trial the same as in Experiment 1, a blank screen was presented for 2500ms after participants responded about the last shape. A question was also included in the demographic information questionnaire to ask participants the number of years in education they had completed, in order to allow further demographic comparisons between the groups in this experiment.

Regarding the instructions check, the mean number of attempts the ADHD-symptoms group took to respond correctly was 1.09 in the equal condition (SD = 0.29) and 1.18 (SD = 0.39) in the differential value condition. Meanwhile, the mean number of times participants in the control group took to respond correctly was 1.12 (SD = 0.33) in the equal value condition and 1.03 (SD = 0.17) in the differential value condition (see supplementary material for frequencies). There were no significant differences between groups in either condition (equal value condition p = 1.00; differential value condition p = 1.05).

Results

Effect at SP1 (targeted SP)

Mean proportion correct at SP1 is displayed in Figure 5A, as a function of value and group. Figure 5B displays the difference between the differential and equal value conditions as a function of group. A 2 (Value: differential vs equal; within-subject) x 2 (Group: ADHD-symptoms vs control; between-subjects) mixed ANOVA revealed a significant main effect of value (F(1, 66) = 73.49, MSE = 0.04, p < .001, $\eta_p^2 = .53$; $BF_{10} > 10,000$) whereby participants exhibited higher performance in the differential value condition (M = 0.78, SE = 0.02). relative to the equal value condition (M = 0.50, SE = 0.02). There was no effect of group (F(1, 66) = 1.31, MSE = 0.05, p = .256, $\eta_p^2 = .02$; $BF_{10} = 0.37$, $BF_{01} = 2.68$), and no interaction between value and group (F(1, 66) = 0.07, MSE = 0.04, p = .794, $\eta_p^2 < .01$; $BF_{10} = 0.26$, $BF_{01} = 3.90$). BF analysis revealed that the best model included a main effect of value ($BF_{10} > 10,000$ relative to the model containing participant only).

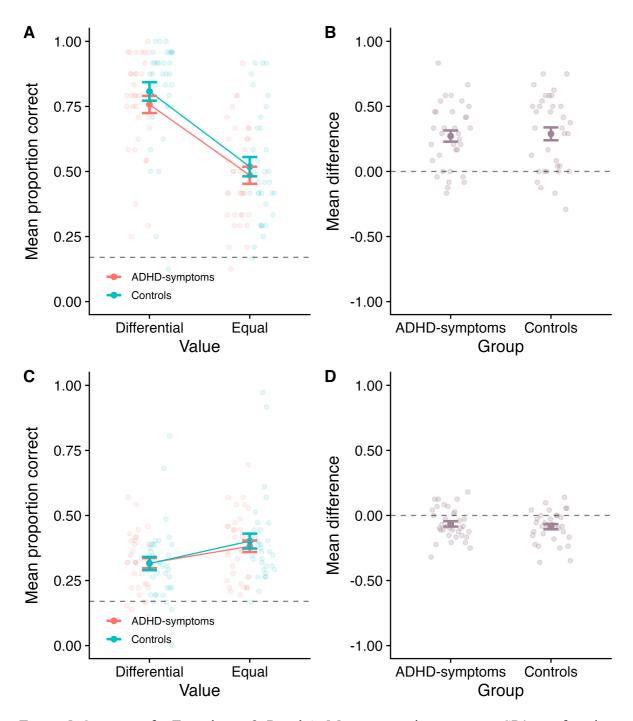


Figure 5. Outcomes for Experiment 2. Panel A: Mean proportion correct at SP1 as a function of value and group. Panel B: The mean difference between the differential and equal value conditions at SP1, as a function of group (calculating as Differential – Equal). Panel C: Mean proportion correct averaged across the less valuable SPs, as a function of value and group. Panel D: The mean difference between the differential and equal value conditions averaged across the less valuable SPs, as a function of group (calculating as Differential – Equal). In all panels, the bold point reflects the mean across participants, whilst error bars display standard error. The lighter points in the background reflect the means for individual participants. In Panels A and C, the dotted line at 0.17 reflects chance guessing rate, based on a stimulus set of six items. In Panels B and D, the dotted line at 0.00 reflects no difference.

Effects averaged across less valuable SPs (2-4)

Mean proportion correct averaged across the less valuable SPs (2-4) is displayed in Figure 5C, as a function of value and group. Meanwhile, Figure 5D displays the mean difference between the Differential and Equal value conditions as a function of group. A 2 (Value: differential vs equal; within-subject) x 2 (Group: ADHD-symptoms vs control; between-subjects) mixed ANOVA found a significant main effect of value (F(1, 66) = 27.33, MSE = 0.01, p < .001, $\eta_p^2 = .29$; $BF_{10} = 9387.86$) whereby participants performed better in the differential value condition (M = 0.32, SE = 0.02) relative to the equal value condition (M = 0.39, SE = 0.02). There was no effect of group (F(1, 66) = 0.09, MSE = 0.03, p = .767, $\eta_p^2 < .01$; $BF_{10} = 0.36$, $BF_{01} = 2.81$), and no interaction between value and group (F(1, 66) = 0.56, MSE = 0.01, p = .455, $\eta_p^2 = .01$; $BF_{10} = 0.31$, $BF_{01} = 3.19$). The BF analysis revealed that the best model included a main effect of value ($BF_{10} = 7734.59$ relative to the model containing participant only).

Correlations between the degree of inattentive and hyperactive symptoms and prioritization boosts and costs

As in Experiment 1, exploratory analysis was then conducted to investigate whether the degree of inattentive and hyperactive symptoms correlates with the prioritization boost (at SP1) and the cost (at the other SPs). Figure 6 displays scatterplots of ASRS Screener scores, boosts to the high value item, and costs to the low value items. Pearson's correlation coefficients were conducted to investigate correlations between scores on the ASRS Screener and the prioritization boosts at SP1 and the costs at the less valuable SPs. P-values were corrected using Bonferroni-Holm. There was no significant correlation between the ASRS Screener score and the prioritization boost, either when considering all participants or each group separately (see Table 1). In all cases, the BFs were at least anecdotally in favour of no correlations (although the BF for the correlation between the ASRS Screener score and the cost to less valuable SPs in the ADHD-symptoms group was close to 1).²

separately.

² There were also no significant correlations between the ASRS, boosts at SP1 and costs to the less valuable items following Bonferroni-Holm correction when measures of inattention and hyperactivity were analysed

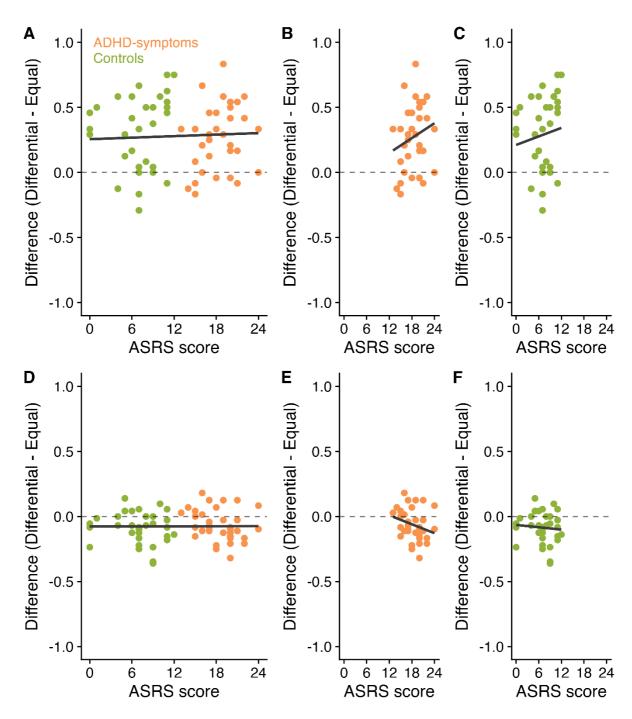


Figure 6. Correlations between ASRS Screener scores (reflecting the degree of inattentive and hyperactive behaviours), boosts to SP1 and costs to the less valuable items in Experiment 2. Panel A displays the correlation between ASRS Screener score and the difference in performance between the differential and equal value items at SP1 across participants. Panels B and C then present these correlations for the ADHD-symptoms and control groups, respectively. Panel D displays the correlation between ASRS Screener score and the difference in performance between the differential and equal value items at the less valuable serial positions. Panels E and F then present these correlations for the ADHD-symptoms and control groups, respectively.

Discussion

Previous research has suggested that the combination of reward and feedback may be particularly beneficial for working memory performance in individuals with ADHD (Hammer et al., 2015). Experiment 2 therefore investigated whether prioritization boosts could be observed in individuals with symptoms of ADHD in the absence of trial-by-trial feedback. The results demonstrated that individuals with ADHD were able to prioritize a particularly high value item in working memory in the absence of trial-by-trial feedback. Further, there were no interactions including group, demonstrating that the effects observed did not significantly differ between the ADHD-symptoms group and a control group. These findings are in line with Experiment 1, providing further evidence that individuals with symptoms of ADHD do not exhibit difficulties in value-based prioritization within working memory. Further, these findings indicate that the ability to prioritize particularly valuable information in individuals with symptoms of ADHD is not contingent on trial-by-trial reinforcement regarding the points system. As in Experiment 1, there were costs to the less valuable SPs, with no evidence that these differed across groups. This provides further support for the notion that individuals with symptoms of ADHD experience costs of prioritization in a similar way to controls.

Cross-experimental analyses

Given the similarities between the experiments, cross-experimental analyses were conducted to investigate whether effects differed as a function of a group when the data for both experiments was combined, thus providing greater statistical power. This analysis therefore ignores experiment (i.e. whether trial-by-trial feedback was provided or not).

Performance at SP1

Mean performance at SP1 is displayed in Figure 7A, as a function of value and group. Figure 7B displays the difference between the differential and equal value conditions as a function of group. A 2 (Value: differential vs equal; within-subject) x 2 (Group: ADHD-symptoms vs control; between-subjects) mixed ANOVA was conducted. This revealed a significant main effect of value (F(1, 131) = 166.64, MSE = 0.03, p < .001, $\eta_p^2 = .56$; $BF_{10} > 10,000$) with higher accuracy observed in the differential value condition (M = 0.78, SE = 0.02). relative to

the equal value condition (M = 0.49, SE = 0.02). There was no effect of group (F(1, 131) = 3.77, MSE = 0.04, p = .054 $\eta_p^2 = .03$; $BF_{I0} = 0.85$, $BF_{0I} = 1.17$), and no interaction between value and group (F(1, 131) = 0.48, MSE = 0.03, p = .488, $\eta_p^2 < .01$; $BF_{I0} = 0.23$, $BF_{0I} = 4.27$). BF analysis revealed that the best model included a main effect of value ($BF_{I0} > 10,000$ relative to the model containing participant only).

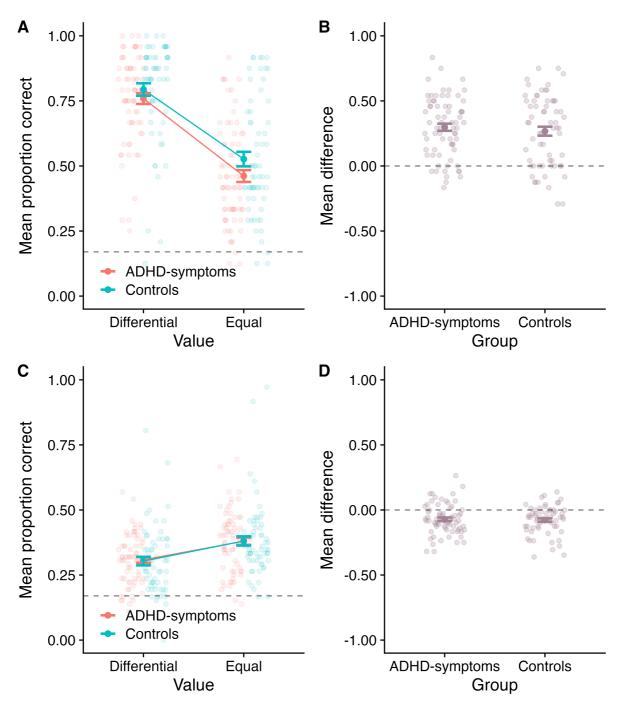


Figure 7. Outcomes when data for both experiments were combined. Panel A: Mean proportion correct at SP1 as a function of value and group. Panel B: The mean difference between the differential and equal value conditions at SP1, as a function of group (calculating

as Differential – Equal). Panel C: Mean proportion correct averaged across the less valuable SPs, as a function of value and group. Panel D: The mean difference between the differential and equal value conditions averaged across the less valuable SPs, as a function of group (calculating as Differential – Equal). In all panels, the bold point reflects the mean across participants, whilst error bars display standard error. The lighter points in the background reflect the means for individual participants. In Panels A and C, the dotted line at 0.17 reflects chance guessing rate, based on a stimulus set of six items. In Panels B and D, the dotted line at 0.00 reflects no difference.

Effects averaged across less valuable SPs (2-4)

Mean performance the less valuable SPs is presented in Figure 7C, as a function of value and group. Figure 7D presents the difference between the differential and equal value conditions as a function of group. A 2 (Value: differential vs equal; within-subject) x 2 (Group: ADHD-symptoms vs control; between-subjects) mixed ANOVA revealed a significant main effect of value $(F(1, 131) = 59.51, MSE = 0.01, p < .001, \eta_p^2 = .31; BF_{10} > 10,000)$ with participants performing more accurately in the equal value condition (M = 0.38, SE = 0.01) relative to the differential value condition (M = 0.31, SE = 0.01). No effect of group emerged $(F(1, 131) = 0.01, MSE = 0.02, p = .911, \eta_p^2 < .01; BF_{10} = 0.25, BF_{01} = 4.00)$, and there was also no interaction between value and group $(F(1, 131) = 0.10, MSE = 0.01, p = .750, \eta_p^2 < .01; BF_{10} = 0.20, BF_{01} = 5.10)$. BF analysis revealed that the best model included a main effect of value $(BF_{10} > 10,000 \text{ relative to the model containing participant only})$.

Other cross-experimental analyses

Analysis was also performed to investigate whether correlations emerged between ASRS Screener scores, the prioritization boost at SP1, and the prioritization costs across all participants (ignoring experiment). Following correction for multiple comparisons, there were no significant correlations (although the BF analysis did provide some support for a correlation between the ASRS Screener and costs to the less valuable SPs in the ADHD-symptoms group only; see supplementary materials for full reporting).

Furthermore, analysis was performed across all data in both experiments to investigate whether the effects observed differ depending on whether individuals had a formal diagnosis of ADHD or not. The effects were similar across all groups (Control, ADHD-symptoms

group (with formal diagnosis), ADHD-symptoms group (without formal diagnosis; see supplementary material for full reporting).

Finally, analysis was conducted to investigate whether overall performance and the size of the prioritization effects differed between experiments, where trial-by-trial feedback was present (Experiment 1) and absent (Experiment 2). Given that the above analyses found no effect of group and that the prioritization effects did not differ as a function of group, the effects were examined across group to maximise statistical power. The outcomes are presented in the supplementary materials. To summarise, there was no effect of Experiment, and no interactions including Experiment. This provides evidence that the provision of trial-by-trial feedback did not impact overall performance, or the size of the prioritization effects.

General Discussion

The current pair of experiments investigated the extent to which adults with symptoms of ADHD can direct their attention in working memory to particularly valuable information, and whether this ability was impaired (or enhanced) relative to individuals without symptoms of ADHD. The provision of trial-by-trial feedback was varied across experiments, with feedback provided in Experiment 1 and absent in Experiment 2. Across both experiments, performance for the high value item (SP1) was superior in the differential value condition, relative to the equal value condition. As in previous research, there were also costs to the less valuable items. Crucially, there was no evidence that the boost to the particularly valuable item or the costs to the less valuable items differed as a function of group. This provides evidence that adults with symptoms of ADHD were as able to direct attention in working memory based on value as control participants. Moreover, exploratory analyses revealed no consistent evidence that the degree of inattentive and hyperactive symptoms (measured by the ASRS Screener) correlated with prioritization effects, either across groups or within the groups. Similar outcomes were also observed across a number of subsidiary and crossexperimental analyses (see supplementary materials), indicating that the findings are highly robust to alternative, plausible analytical decisions. Taken together, these experiments provide the first evidence that individuals with symptoms of ADHD can prioritize particularly valuable information in working memory, with no impairment observed relative to control participants.

Evidence that individuals with ADHD symptoms are as able to direct attention in working memory is in line with previous research which found similar findings using predictive cueing paradigms with children and adolescents (Superbia-Guimarães et al., 2022). Taken together, these findings suggest that the ability to use the focus of attention within working memory to focus on particularly valuable or task-relevant information is not impaired in individuals with symptoms of ADHD. In the current pair of experiments, the boost at the first (targeted) serial position was accompanied by costs to less valuable items. Within nonclinical groups, this has been taken as evidence that value instructions do not enhance participants' working memory capacity, but instead results from a re-distribution of attention, whereby participants focus more on the targeted item at the expense of others in the sequence (e.g. Allen et al., 2024; Atkinson et al. 2024; Hitch et al. 2018). Here, a similar pattern was observed across both groups, suggesting a similar pattern is present in individuals with symptoms of ADHD and controls. Thus, increasing the value of a particular item did not appear to motivate participants with symptoms of ADHD to simply try harder or perform better on the task in general. Instead, as with controls, it resulted in them focusing their existing working memory resources towards the particularly valuable information.

Previous studies with non-clinical groups suggested that prioritization effects are likely to emerge, in part, through participants refreshing the high value item more during the retention interval (Atkinson et al., 2022). Given that the values associated with items are known at encoding, these effects may also at least partially reflect participants directing their attention towards the item during this phase (Allen et al., 2024). Whilst the current pair of experiments suggest that individuals with symptoms of ADHD are as able to direct attention in working memory as controls, underlying mechanisms were not explored. As such, it is possible that the mechanisms underlying the prioritization effects observed differ in individuals with symptoms of ADHD and controls. Some existing research may support this possibility. For example, previous studies have found that individuals with ADHD have difficulties allocating attention during the encoding phase (e.g. Kim et al., 2014). As such, it may be that individuals with symptoms of ADHD achieve the same size prioritization boost as controls, but through different strategic approaches. It would be beneficial for further research to investigate this.

Previous research has found that overall working memory accuracy is enhanced in children and adolescents with ADHD if participants have the potential to earn financial rewards. For

example, Dovis et al. (2012) found that individuals with ADHD performed better on a working memory task if they had the potential to earn 1 euro or 10 euros, relative to a condition which only feedback was provided. In contrast, controls performed no differently in these conditions. The current findings are somewhat inconsistent with this, finding that increasing the value of particular items did not result in a greater boost for individuals with symptoms of ADHD. These differences may reflect differences between the studies. For example, it is possible that the effect of value was not larger in the ADHD-symptoms group relative to the control group due to the rewards offered. In the current pair of experiments, participants could earn notional points. However, in previous research finding larger effects in the ADHD-symptoms group, participants were able to earn financial reward. Studies in neurotypical adults have found that the size of prioritization effects in working memory do not substantially differ when using notional rewards or financial incentives (Zheng et al., 2022). However, it is possible that the potential to earn financial incentives have a larger impact in the ADHD-symptoms group. A second possibility is that the difference may reflect the age groups of the participants tested. Adults were tested in the current pair of experiments, compared to children and adolescents in Dovis et al. (2012). Therefore, one possibility is that the effect of value may become smaller over time in individuals with ADHD.

There was no overall effect of group in either experiment. This suggests that individuals with symptoms of ADHD performed no worse than controls in these experiments. Given that working memory difficulties are commonly observed in individuals with ADHD, it is possible that there could have been a selection bias whereby participants in the ADHD-symptoms group may have less impaired working memory relative to the ADHD population as a whole. However, we believe this is unlikely to be the case for several reasons. Firstly, all participants included in the ADHD-symptoms group screened positively for ADHD on the ASRS Screener, indicating that they exhibited symptoms that were highly consistent with ADHD. Secondly, the outcomes did not differ in individuals with and without a formal diagnosis of ADHD (see supplementary materials). Finally, there was no significant correlation between the ASRS Screener score, the prioritization boost at the first serial position, and the prioritization cost at the less valuable serial positions. Therefore, there is no evidence to suggest that the effects observed differ depending on the degree of inattentive and hyperactive symptoms. An alternative possibility for the absence of a group effect is that the ADHD-symptoms group may have been more motivated when completing the task in

general relative to the control group, possibly due to the potential to earn points. If so, this may mask any overall performance differences between the groups. However, this is unlikely to explain the absence of an effect, as previous research has found that even financial rewards do not "normalise" working memory performance in individuals with ADHD (Dovis et al., 2012). As such, it is unlikely that the potential to earn notional points (with no real-world value) would eliminate any differences between groups.

A related possibility is that the ADHD-symptoms group may have impairments in the ability to direct attention in working memory, but that these were masked because the ADHD-symptoms group tried harder to prioritize the high value item. However, this is unlikely due to several reasons. Firstly, if participants in the ADHD-symptoms group were trying harder to direct attention to SP1, this should have been accompanied by greater costs to the less valuable items. This was not observed in either experiment or in the cross-experimental comparisons. Secondly, if this was the case, one might have expected that the ADHD-symptoms group would show a smaller prioritization effect in Experiment 2 when they received no feedback regarding the points system. However, the results of both experiments were consistent in showing no impairment in the ADHD-symptoms group's ability to direct attention in working memory. Thirdly, evidence that the ADHD-symptoms group showed no impairment in their ability to direct attention in working memory is consistent with existing research using other manipulations that do not involve the manipulation of points or rewards (e.g. cueing; Superbia-Guimarães et al., 2022). As such, this explanation is unlikely to explain the pattern of results observed.

The cross-experimental analyses revealed no difference in overall performance between the experiments, and that the prioritization effects observed did not differ as a function of experiment. This suggests that prioritization effects in working memory are not significantly affected by whether trial-by-trial feedback is provided. This adds to a growing body of research identifying factors that do, and do not, appear to affect prioritization effects in working memory. For example, whilst the number of items presented does appear to influence the size of the effect (Atkinson et al., 2019; Zheng et al., 2022), neither the type of reward (e.g. notional points vs monetary rewards; Zheng et al., 2022) or the provision of trial-by-trial feedback appears to.

In both experiments, the value manipulation was targeted towards SP1. This was based on previous research, which have often targeted the first SP when implementing this paradigm (e.g. Atkinson et al., 2018; 2019). As we directly compared data between the differential and an equal value condition at SP1, our comparison would not be confounded by any primacy effect which is typically observed in working memory tasks (higher accuracy often observed at the first SP). Within non-clinical adult samples, there is evidence that individuals can prioritize valuable information presented at early, middle, or late positions in a sequence (Atkinson et al., 2021; Hitch et al., 2018). It is, however, possible that prioritization at different SPs may require different skills or abilities. For example, if the first item is associated with a higher value, individuals may need to protect the privileged status associated with this item when encoding subsequent items. The need to do this may be greatly reduced if participants are asked to prioritize later positions. However, prioritizing these later positions may instead require other skills (e.g. tracking the position of items within the sequence). It would therefore be beneficial for further research to investigate prioritization in individuals with symptoms of ADHD at other positions within the sequence. It would also be beneficial for further research to examine prioritization in individuals with ADHD in other contexts. For example, research with non-clinical adult samples has begun to explore the extent to which individuals can direct attention in working memory when value information is provided only during the retention interval (Allen & Atkinson, 2021; Jeanneret et al., 2023; 2024; Hautekiet et al., 2024). This may be more challenging, as individuals cannot direct attention during the encoding phase in this paradigm (Allen & Atkinson, 2021; see Allen et al., 2024). Indeed, prioritization effects are markedly smaller in such tasks than when individuals can prioritize information during encoding (Allen & Atkinson, 2021). One possibility is that individuals with symptoms of ADHD would show impairments in the ability to direct attention in working memory under these more challenging circumstances.

Although our primary conclusions were drawn based on the outcomes from the ANOVAs, we also conducted exploratory analyses examining correlations between the ASRS, boosts to SP1, and costs to the less valuable SPs. However, it is worth highlighting that the ASRS was primarily implemented in these experiments as a screening tool to aid group assignment and the range was somewhat limited (i.e. on a 0-24 scale). It is therefore possible that correlations between the degree of inattentive and hyperactive symptoms and the prioritization effects may be observed if a more fine-grained measure of ADHD symptoms was used. It would be beneficial for further research to investigate this.

A further limitation is that the groups were not age-matched in Experiment 2, with the ADHD-symptoms group significantly younger than the control group. However, it is unlikely this affected overall performance between the groups, as working memory abilities are relatively stable between the ages of 18-35 years (the possible age range of participants; Brockmole & Logie, 2013). This is also unlikely to have affected the size of the prioritization effects observed, as the ability to prioritize valuable information in working memory is preserved even in older adulthood (Allen et al., 2021). Moreover, it is possible that the groups could have differed in an unmeasured variable that could explain the pattern of results observed (e.g. education and/or IQ in Experiment 1, or IQ in Experiment 2). Although there is no evidence that the ability to direct attention in working memory is linked to factors such as educational level or IQ, the absence of such measures remains a limitation of this study. It may therefore be beneficial for future research to further examine whether the ability to direct attention in working memory is impaired in individuals with ADHD who are matched on a wider range of variables (e.g. age, gender, education level, and IQ).

To summarise, across all analyses, there was clear evidence that individuals with symptoms of ADHD are able to prioritize particularly valuable information in working memory. The ability in this group did not significantly differ relative to controls, suggesting no impairment in the ability to prioritize particularly valuable information in working memory. Converging evidence of this was also demonstrated from correlational analyses, which found no significant associations between scores on the ASRS Screener (indicating the degree of inattentive and hyperactive symptoms) and the prioritization boost at the targeted SP (SP1) and costs at the less valuable serial positions. Taken together, this study provides clear and consistent evidence that the ability to prioritize particularly valuable information in working memory is not impaired in individuals with symptoms of ADHD.

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