2

Assessing the growth and climate sensitivity of secondary forests in highly deforested
Amazonian landscapes

5

Fernando Elias^{1*}, Joice Ferreira^{1,2}, Gareth D. Lennox³, Erika Berenguer^{3,8}, Socorro Ferreira²,
Gustavo Schwartz², Lia de Oliveira Melo⁴, Denilson do Nascimento Reis Júnior⁵, Rodrigo Oliveira
do Nascimento⁶, Fabrício Nascimento Ferreira², Fernando Espirito-Santo⁷, Charlotte C. Smith³,
Jos Barlow^{3,9}

10

11 ^{1*}Universidade Federal do Pará/Embrapa Amazônia Oriental, Instituto de Ciências Biológicas,

12 Programa de Pós-Graduação em Ecologia, Belém, Pará, 66075-110, Brazil.

¹³ ²Embrapa Amazônia Oriental, Belém, Pará, 66095-903, Brazil.

¹⁴ ³Lancaster Environment Centre, Lancaster University, Lancaster, LA1 4YQ, UK.

⁴Universidade Federal do Oeste do Pará, Santarém, Pará, 68040-255, Brazil.

⁵Universidade Federal Rural da Amazônia, Departamento de Engenharia Florestal, Belém, Pará,
66077-830, Brazil.

18 ⁶Universidade Federal do Pará, Instituto de Geociências, Programa de Pós-Graduação em Ciências

19	Ambientais,	Belém,	Pará,	66075-1	l 10, Brazil.
----	-------------	--------	-------	---------	---------------

⁷Centre for Landscape and Climate Research, Leicester Institute of Space and Earth Observation,

- 21 School of Geography, Geology and Environment, University of Leicester, University Road,
- 22 Leicester LE1 7RH, UK.

⁸Environmental Change Institute, School of Geography and the Environment, University of

24 Oxford, Oxford, OX1 3QY, UK.

⁹Universidade Federal de Lavras, Lavras, Minas Gerais, 37200-000, Brazil.

26 *Corresponding author. E-mail: <u>fernandoeliasbio@gmail.com</u>

27

28 Abstract

29 Tropical forests hold 30% of Earth's terrestrial carbon and at least 60% of its terrestrial 30 biodiversity, but forest loss and degradation are jeopardizing these ecosystems. Although the 31 regrowth of secondary forests has the potential to offset some of the losses of carbon and 32 biodiversity, it remains unclear if secondary regeneration will be affected by climate changes such 33 as higher temperatures and more frequent extreme droughts. We used a dataset of 10 repeated 34 forest inventories spanning two decades (1999-2017) to investigate carbon and tree species 35 recovery and how climate and landscape context influence carbon dynamics in an older secondary 36 forest located in one of the oldest post-Columbian agricultural frontiers in the Brazilian Amazon. Carbon accumulation averaged 1.08 Mg ha⁻¹ yr⁻¹, while species richness was effectively constant 37 38 over the studied period. Moreover, we provide evidence that secondary forests are vulnerable to 39 drought stress: carbon balance and growth rates were lower in drier periods. This contrasts with 40 drought responses in primary forests, where changes in carbon dynamics are driven by increased 41 stem mortality. These results highlight an important climate change-vegetation feedback, whereby 42 the increasing dry-season lengths being observed across parts of Amazonia may reduce the 43 effectiveness of secondary forests in sequestering carbon and mitigating climate change. In 44 addition, the current rate of forest regrowth in this region was low compared with previous pan-45 tropical and Amazonian assessments – our secondary forests reached just 41.1% of the average 46 carbon and 56% of the tree diversity in the nearest primary forests — suggesting that these areas

47 are unlikely to return to their original levels on politically meaningful timescales.

48 Keywords

49 Amazon, Bragantina region, biodiversity, carbon, climate change, secondary forests

50

51 Introduction

52 Secondary forests regenerating after land clearance represent > 50% of all tropical forests 53 (Chazdon et al. 2009), and there is hope that they can cost-effectively mitigate climate change and 54 biodiversity loss (Chazdon 2014, Lewis et al. 2019) - two of the defining crises of the 55 Anthropocene (Malhi et al. 2014). For example, secondary forest carbon uptake is estimated to be 56 11 times higher than that of Neotropical primary forests (Poorter et al. 2016), providing an estimated global carbon sink of 130 Pg between 2016 and 2100 (Houghton and Nassikas 2018). 57 58 Secondary forests can also provide important biodiversity co-benefits in landscapes otherwise 59 devoid of primary forests or with high levels of deforestation (Vieira and Gardner 2012, Lennox 60 et al. 2018, Matos et al. 2019). However, to realize their potential to mitigate climate change and 61 biodiversity loss, tropical secondary forests must be able to recover under novel climatic 62 conditions.

Global climate change is affecting the humid tropics through higher temperatures and levels of atmospheric CO_2 (Malhi et al. 2014), increases in dry-season length (Fu et al. 2013) and the frequency and intensity of climate extremes (Brando et al. 2019). These climate changes may be exacerbated by declines in precipitation and increases in temperature linked to regional deforestation (Spracklen et al. 2018, Baker and Spracklen 2019). Three types of evidence suggest tropical secondary forests will be sensitive to these changes. First, primary forest estimates show a decrease in the carbon sink during extreme droughts (Brienen et al. 2015). Second, large-scale studies of secondary forests indicate a strong effect of climate on the recovery rates of carbon and biodiversity (Anderson-Teixeira et al. 2013, Poorter et al. 2016, Rozendaal et al. 2019). Third, longitudinal studies have revealed how droughts modulate recovery speed, increase mortality and reduce recruitment and growth (Chazdon et al. 2005, Maza-Villalobos et al. 2013, Martínez-Ramos et al. 2018), due to the negative effects on the water balance and photosynthetic rates of trees, especially in the earlier stages of succession (Bretfeld et al. 2018).

76 While these studies provide a useful knowledge base, there are some important limitations. 77 First, secondary forests are likely to be even more sensitive than primary forests to changes in 78 precipitation (Uriarte et al. 2016) because pioneer species are more vulnerable to embolism 79 (Markesteijn et al. 2011). Second, the large-scale studies that have inferred temporal trends from 80 spatial data (i.e. the chronosequence approach) are complicated by factors such as species turnover 81 and natural variation between samples (Johnson and Miyanishi 2008, Norden et al. 2015, Franca 82 et al. 2016) and idiosyncratic determinants of recovery (Chazdon et al. 2007, Arroyo-Rodríguez et 83 al. 2017). Third, most longitudinal studies focus on short-term assessments restricted to a single 84 drought event (Chazdon et al. 2005, Maza-Villalobos et al. 2013, Bretfeld et al. 2018, Martínez-85 Ramos et al. 2018), which cannot detect responses to longer-term increases in temperature or post-86 drought growth compensation (Berenguer et al. 2018). Furthermore, the few longer-term studies 87 relating secondary forest carbon dynamics to climate focus on tropical dry forests (e.g., Álvarez-88 Yépiz et al. 2018, Martínez-Ramos et al. 2018), meaning there is a lack of research in humid 89 tropical regions where species may be even more sensitive to drought (Esquivel-Muelbert et al. 90 2019).

Initiatives such as the Bonn Challenge highlight two additional knowledge gaps linked to
 management. It seems likely that the increase in landscape-scale restoration will also increase

93 secondary forest permanence, moving beyond the current situation where secondary forests are 94 often cleared again within 5-20 years (Aguiar et al. 2016, Reid et al. 2018). Although our current 95 understanding of younger forests is good (Poorter et al. 2015, 2016, 2019, Martínez-Ramos et al. 96 2018, Villa et al. 2018, Rozendaal et al. 2019), and the non-linear response of forest recovery over 97 time is well established (Poorter et al. 2016, Ferreira et al. 2018, Lennox et al. 2018, Requena 98 Suarez et al. 2019, Rozendaal et al. 2019), there is far too much variation in the relationship to use 99 young forests to accurately predict recovery rates in older forests. Moreover, the relative recovery 100 rates of carbon and biodiversity are unclear. While some regional assessments report a tight 101 coupling at all stages of succession (Lennox et al. 2018), continent-scale assessments suggest a 102 much faster recovery of biodiversity: plots recovered 80% of tree species richness 20 years after 103 abandonment (Rozendaal et al. 2019) and only 27% of the carbon stock (Poorter et al. 2016). As 104 recovery of both carbon and biodiversity are mediated by factors such as stand age and landscape 105 and local conditions (Jakovac et al. 2015, Magnago et al. 2017, Villa et al. 2018, Ferreira et al. 106 2018, Lennox et al. 2018, Matos et al. 2019), it is not clear how these rates compare in heavily 107 deforested landscapes.

108 Here, we address these knowledge gaps by investigating secondary forest recovery in the 109 Bragantina region, the oldest deforestation frontier in the Brazilian Amazon (Almeida et al. 2010). 110 We undertook 10 repeated censuses of 3 ha of older secondary forests (~60 years) over 18 years and addressed four questions: 1) What are the recovery rates of carbon stocks and tree species 111 112 richness, and what is the relationship between these measures? 2) Does variation in seasonal 113 temperature, water stress and landscape context influence carbon dynamics? 3) How does carbon 114 recovery in the Bragantina region compare to estimates from other tropical regions? 4) What is the 115 timeframe for carbon to return to typical primary forest levels?

116

117 Methods

118 Study area and landscape context

Our study focused on the eastern-Amazonian municipality of Bragança (Appendix S1: Fig. S1). This municipality has lost 90.2% of its native forests and mangrove areas. Secondary forest covers 28.8% of the landscape, accounting for 67.5% of the total forest cover (including mangrove areas) (Appendix S1: Fig. S1). Primary and secondary forests (established post-1985) are situated in small, isolated, and selectively logged fragments (~30 to 60 hectares; MapBiomas 2019).

We established three study-plot classes. In 1999, we began inventories in 12 secondary forest plots (50 x 50 m = 0.25 ha; Appendix S1: Fig. S1, Table S1; hereafter, "*long-term plots*"). These plots were separated by a mean distance of 265 m (range 70-590 m). In 2017, we established an additional four secondary forest plots (250 x 10 m = 0.25 ha; hereafter, "*extra plots*") in different fragments of the Bragança municipality. In 2017, we also established three plots in primary forests (250 x 10 m 0.25 ha; Appendix S1: Fig. S1; hereafter "*primary forest plots*").

130 Both long-term and extra secondary forest plots were abandoned after successive cycles of 131 slash-and-burn agriculture (for manioc, maize, and rice cultivation). There was no record of 132 wildfires after agricultural abandonment at our plots. Given that undisturbed primary forests are 133 extremely rare in the region, to select primary forest plots we conducted interviews with local 134 people to identify sites that retain the structural characteristics of Amazonian old-growth forests; 135 nonetheless, it is likely that selective logging and edge effects have already altered our primary 136 forest plots. All plots were located in *terra-firme* areas. The predominant soils across all plots are 137 oxisols with low fertility and 15-35% clay in the superficial layers (Da Silva Castro et al. 2013). 138 Plots were flat, and average elevation was 35 m (range: 30-66 m).

139

140 Age of secondary forests

141 We used two approaches to evaluate secondary forest age. The long-term plots are older 142 than the existing satellite record and were provisionally aged by interviews conducted by the 143 researchers who established them in 1999. The estimated age in 2017 was c. 48-58 years old. 144 Unfortunately, Landsat 1 images from the region (1974) were also too cloudy to support these 145 interviews. We therefore estimated ages by back casting the non-linear growth trajectories 146 (Question 4). For the extra plots recovering post 1985, we used the MapBiomas 3.1 dataset to 147 estimate their ages (details in Appendix S1; MapBiomas 2019). 148 149 Tree censuses 150 The long-term secondary forest plots were established with a full tree census in 1999, with 151 repeated surveys taking place annually between 2000-2004 then subsequently in 2006, 2011, 2014 152 and 2017, while a full tree census was undertaken in the extra and primary forest plots when they 153 were established in 2017. In all plots, we measured all trees ≥ 10 cm diameter at breast height 154 (DBH). In the last census, we also estimated tree height by visual assessment. Plant identification 155 was conducted in the field and when necessary botanical samples were collected for comparisons 156 from the Herbário IAN (Embrapa Amazônia Oriental) collection. Tree census data are available at 157 ForestPlots.net (Lopez-Gonzalez et al. 2011). 158

159 Estimation of carbon stocks and tree species richness

160 We used three approaches to estimate the above-ground biomass (AGB) of individual 161 stems. First, we calculated AGB using the allometric equation: $AGB = 0.637 \times (\rho D^2 H)^{0.976}$ 162 where ρ represents tree wood density (g cm⁻³), *D* represents stem DBH (cm) and *H* represents stem 163 height (Chave et al. 2014). Stem wood densities were taken from the Global Wood Density 164 Database (Chave et al. 2009). We subset the data to entries for South America and used the mean 165 wood density across entries at the lowest available taxonomic level (e.g. where no data was 166 available for a species, we used the genus average).

167 To estimate the height of stems for censuses prior to 2017, we used non-linear least-squares 168 to determine height-DBH relationships at each study plot. Using the 2017 height and DBH values 169 as training data, we assumed a functional relationship described by the Michaelis-Menten model: $H = \frac{a \times D}{b \times D}$, where a and b are estimated from the training data. Second, due to potential errors in 170 the tree height sample caused by our visual assessment approach, we applied correction functions 171 172 to the 2017 stem heights then determined pre-2017 tree heights as described above. The height-173 correction functions were derived from large primary and second secondary forest tree samples (c. 174 500 individuals in both forest types) from the eastern Brazilian Amazon for which visual and laser-175 based heights were estimated. These data suggest that visual assessments tend to underestimate 176 stem height, especially stems < 15 m tall in secondary forests (Appendix S1: Fig. S2). Finally, we 177 also estimated stem biomass using the following allometric equation that does not include a height parameter: $AGB = exp \left[-1.803 - 0.976E + 0.976 \ln (\rho) + 2.673 \ln (D) - 0.299 \ln (D)^2 \right]$ 178 179 where E is a measure of environmental stress, defined by cumulative water deficit and temperature 180 and precipitation seasonality (Chave et al. 2014).

181 Results across these approaches suggest that use of the uncorrected values of stem height 182 values underestimates plot-level carbon by around 10-20%, so we do not report results from this 183 approach. In the main text, we focus on results from the height-corrected estimation technique 184 because including height is known to improve AGB estimates (Sullivan et al. 2018). However, the height-corrected approach returned highly similar carbon estimates to those obtained using the allometric equation that does not include height (mean root square difference < 1.5 Mg C ha⁻¹). The results for this latter approach can be found in Appendix S1: Fig. S6. We assumed that the carbon content of the individuals represents 50% of the AGB. Tree species richness (of individuals > 10 cm DBH) in each plot was assessed by rarefying richness to 100 individuals, the minimum abundance in the sampled plots.

191

192 Climatological and landscape predictors of carbon dynamics

Monthly precipitation and temperature for each census year were determined using data collected at a local meteorological station (INMET 2018; Appendix S1: Fig. S1). These data show that the study region has annual precipitation of 1,850 mm, with a rainy season from December to July (222 mm month⁻¹) and a dry season from August to November (19.5 mm month⁻¹).

197 We calculated three climatological predictors of carbon dynamics: the Standardized 198 Precipitation Evapotranspiration Index (SPEI; Vicente-Serrano et al. 2010), the Maximum 199 Climatological Water Deficit (MCWD; Malhi et al. 2009), and the maximum dry season 200 temperature (MDST). SPEI is a measure of water stress based on the difference between monthly 201 precipitation and potential evapotranspiration expressed as a standardized index, with negative 202 values representing drier periods over the timescale considered relative to median values for a 203 long-term reference time-series. To account for medium-term water deficit, we estimated monthly 204 SPEI values for a 12-month moving window from May to April. MCWD is the most negative 205 accumulated value of the climatological water deficit (CWD) over each one-year period (between 206 May and April; 1973-2017). CWD is defined as monthly precipitation, minus evapotranspiration, 207 minus the CWD of the previous month (Malhi et al. 2009). MDST was measured as the highest

208 monthly temperature value during the dry season (i.e., between August and November). Finally, 209 we used edge distance as our predictor of landscape context effects on carbon dynamics. For each 210 secondary forest plot, edge distance was calculated as the distance from the plot centroid to the 211 nearest forest edge using images from Google Earth 7.1.7.2600 (earth.google.com).

212

213 Carbon anomalies

214 We calculated the amount of carbon incorporated by growth (G) and recruitment (R), as 215 well as that lost by mortality (M) for each census interval, thus obtaining periodic carbon balance 216 (B = G + R - M). We used anomalies to subtract the expected natural carbon accumulation that 217 occurs through time. Censuses occurred at varying intervals (1, 2, 3 and 5 years). Let $i \in [1, 2, ..., n]$ 9 be one of the intervals, let T_{-i} be the total length of the time series minus the length of interval 218 *i*, let C_{-i} be the change in carbon balance or one of its components over all intervals not including 219 220 *i*, let n_i be the number of years in interval *i*, and let a_i be the observed change C in interval *i*. Therefore, the expected change in C in interval i is $E(C_i) = \frac{n_i \times C_{-i}}{T_{-i}}$, and the anomaly is $a_t - E(C_t)$. 221

222

223 Statistical analysis

We used Bayesian piecewise linear splines to estimate i) carbon recovery with time; ii) tree species richness recovery with time; and iii) the relationship between carbon and tree species richness recovery (Question 1). Bayesian model comparison with reversible jump Markov Chain Monte Carlo simulations (Lunn et al. 2009) was used to average multiple piecewise-linear models to produce smoothed curvilinear relationships (Thomson et al. 2010). The number and location of change-points were assigned hierarchical prior distributions that included zero changepoints as a possibility. In each simulation, the spline function was comprised of up to k = 6 linear coefficients (i.e., between zero and six) and k - 1 corresponding changepoints. The resulting posterior distributions yielded model-averaged parameter estimates and 95% credible intervals that account for uncertainties about model structure. The model for long-term secondary forest recovery included plot-level random effects to account for expected similarities between plots.

235 To explore whether variation in maximum dry season temperature, water stress (SPEI and 236 MCWD) and landscape context (edge distance) influence tree growth, recruitment, mortality and 237 carbon dynamics (Question 2), we used Bayesian variable selection with nonlinear covariate 238 effects (Thomson et al. 2010). We used the lowest recorded value of MCWD in a given period as 239 our measure of one-off drought severity (see also Esquivel-Muelbert et al. 2019). We used SPEI 240 values (Vicente-Serrano et al. 2010) as a measure of the longer-term (here 12-month) water 241 balance, which is known to be significant for tree growth (Greenwood et al. 2017). The relative 242 importance of an explanatory variable was assessed by the posterior probability of a non-zero 243 effect. We considered posterior probabilities > 0.75 to be indicative of a statistical association.

244 We compared the average carbon accumulation rates of our secondary forest plots (both 245 long-term and extra plots) to estimates of carbon accumulation from leading tropical and global 246 assessments (Bonner et al. 2013, Poorter et al. 2016, Lennox et al. 2018, Requena Suarez et al. 247 2019) and pan-Amazonian estimates in old-growth and logged primary forests (Brienen et al. 2015, 248 Rutishauser et al. 2015) (Question 3). To investigate secondary forest recovery timeframes 249 (Question 4), we fit monotonically increasing and marginally decreasing rational functions, based 250 on third-order polynomials as the numerator and second-order polynomials as the denominator, to 251 the median and lower and upper bounds of the 95% credible intervals for carbon levels in the long-252 term plots (Fig. 1A).

254 **Results**

255 *Climate variation over the time-series (1972-2017)*

The maximum dry season temperature increased by 0.1 °C per decade across our time series (Appendix S1: Fig. S5). Water stress (MCWD and SPEI values) was highly variable, declined weakly, and was highest (lowest values of MCWD and SPEI) in strong El-Niño years. MCWD and SPEI were only weakly related (r = 0.44 across the 1973-2017 time series; r = 0.19 during census years – Appendix S1: Fig. S3).

261

262 Carbon and tree species richness recovery

263 The average carbon stock in 2017 was 141.7 (±16.3 SE), 58.3 (±2.7) and 4.27 (±1.19) Mg 264 ha⁻¹ in, respectively, the primary forest, long-term, and extra plots, whereas rarefied richness was 265 $60.05 (\pm 4.07), 41.2 (\pm 1.4)$ and $33.9 (\pm 3.5)$ (Appendix S1: Table S1). In the long-term secondary 266 forest plots, the recovery of carbon and rarefied richness showed different trajectories among 267 census (Fig. 1A-B). The mean accumulation of carbon in our long-term plots was 1.08 Mg ha⁻¹ yr⁻ 268 ¹ between 1999 and 2017. Carbon and time were associated at the 0.07 level (i.e., 93% of the R^2 269 posterior mass was greater than zero). Rarefied richness displayed a much weaker relationship 270 with time, being associated only at the 0.33 level and with a mean trend of 0.21 species yr⁻¹. Carbon 271 stock in the long-term secondary forest plots reached only 41.1% of the average primary forest 272 level in 2017, while rarefied richness recovered to 56%. Above-ground carbon stocks and plant 273 diversity were not related to each other over time (Fig. 1).

274

275 Influence of climate and edge distance on carbon anomalies

276 Carbon balance was positively related to SPEI in the long-term plots. These changes were

driven by variation in stem growth, with lower than average carbon accumulation in drier periods (Fig. 2). Although the maximum dry season temperature did not affect the overall carbon balance, it had countervailing effects on carbon growth and recruitment: growth was highest and recruitment lowest in the warmest years and vice versa. Carbon loss from mortality was not associated with any of the predictors (Fig 2), and edge distance and MCWD were not associated with any of the carbon anomalies (Appendix S1: Fig. S7).

283

284 *Comparison between carbon recovery rates*

The carbon recovery rate in the long-term plots was lower than most estimates of secondary forest recovery from the Amazon or across the tropics, and our younger extra plots had even lower recovery rates than the older long-term plots (Fig. 3). Recovery rates at our study plots were also below Amazon-wide estimates of primary forest recovery after selective logging and were only slightly higher than the average carbon uptake rate of tropical primary forests (Fig. 3).

290

291 Extrapolating carbon recovery timeframes

Our backwards extrapolation (to zero) gives approximate ages for the long-term secondary forests of 39 and 48 years in 2017 (Fig. 4), which is marginally lower but generally consistent with the ages reported during interviews (c. 48-58 years old in 2017). As a consequence of the low carbon recovery rates of the long-term plots, our extrapolation of future carbon levels — aimed at providing a rough estimate of a plausible recovery window under present-day conditions suggests that it will take at least another century (c. 150 years since abandonment) until the site regains carbon levels similar to local primary forests (Fig. 4).

300 Discussion

301 Influence of climate on carbon recovery in secondary forests

302 Our results show for the first time that secondary forest carbon accumulation in the Amazon 303 can be influenced by periods of water deficit. This suggests that the role of water availability in 304 regulating carbon uptake observed in humid primary forests (e.g., Phillips et al. 2009, Anderson 305 et al. 2018), dry primary (Mendivelso et al. 2014, Álvarez-Yépiz et al. 2018) and dry secondary 306 forests (Maza-Villalobos et al. 2013, Martínez-Ramos et al. 2018) also extends to humid secondary 307 forests. The generality of this finding is consistent with the strong link between the 308 evapotranspiration index SPEI and spatial variation in carbon recovery in tropical forests 309 (Schwalm et al. 2017). Indeed, secondary forests in the humid tropics may be even more sensitive 310 than those in tropical dry regions given that the former show longer recovery times in the face of 311 seasonal water availability (Poorter et al. 2019). Despite the similarity of the findings across forest 312 types, our results also suggest that different mechanisms may underpin the sensitivity of the carbon 313 balance to water deficit. Drought-mediated changes in primary forest carbon balance are driven by 314 increased mortality (Phillips et al. 2009, Anderson et al. 2018), while the changes in the carbon 315 balance observed in this study were primarily driven by reduced growth (Fig. 2).

Although the maximum dry season temperature (MDST) had no influence on the carbon balance of the studied secondary forests, increasing MDST increased the growth of established stems while reducing recruitment. The mechanistic explanation for such contrasting results requires further investigation. However, it might indicate a differential temperature effect between stems size classes. For example, higher temperatures exacerbate the physiological consequences of acute water stress (Lloyd and Farquhar 2008, Markesteijn et al. 2011), and shallow-rooted smaller stems (<10cm DBH) may be more sensitive to drought effects than the medium-sized 323 stems (10-30 cm DBH) (Elias et al. 2018). The increase in growth also suggests that established 324 stems that can cope with higher temperatures may also benefit from factors associated with hotter 325 years, such as better light conditions (Bentos et al. 2017) or faster litter decomposition rates and 326 nutrient cycling (Eichenberg et al. 2013).

The positive relationship between temperature and growth in the assessed secondary forest contrasts with findings in tropical primary forests, where negative relationships (e.g. Feeley et al 2007) and a lack of any relationship (e.g. Rowland et al. 2014) have been detected. This difference may be explained by the dominance in secondary forests of fast growing, heliophilic species (Vieira et al. 1994, Chazdon 2014), which tend to show relatively high optimum temperature points (Slot et al. 2014). A more complete understanding of secondary forest temperature responses is crucial given their importance as a climate change mitigation technology.

334

335 *Carbon recovery*

336 Carbon accumulation rates in our long-term study plots (48-58 years old) were low 337 compared to most other studies (Fig. 3), and only around twice that recorded in Amazonia's 338 primary forests. Although this rate of carbon accumulation is comparable to recent continental-339 scale estimates of biomass recovery in older (20-80 years old) tropical secondary rainforests in the 340 Americas (Requena Suarez et al. 2019; Fig. 3), this similarity masks one important difference: our 341 long-term plots had a much lower lifetime recovery rate than the same continent-scale estimates 342 (Fig. 4B). Combining the mean < 20- and 20-80-year-old accumulation rates of Requena Suarez 343 et al. (2019) suggests that a typical American tropical secondary rainforest will accumulate the 344 median carbon seen at our study site (Fig. 1) in just 20-21 years, less than half the site's estimated 345 age (Fig 4B). Moreover, even the lower bound of the Requena Suarez et al. (2019) 95%

accumulation rate confidence intervals suggests that American tropical secondary rainforestsshould attain our site's median carbon in 28 years.

348 The low lifetime recovery rate and equivalent rates of recovery over the last two decades 349 can only be reconciled if forests in the Bragantina region have a very slow rate of early forest 350 succession (Fig 4B). This is supported by the four younger extra secondary forest plots, which 351 displayed even lower recovery rates than the long-term plots and are far below the predictions of 352 Requena Suarez et al. (2019) for age-equivalent plots (Fig. 3). Furthermore, our backward 353 extrapolation of growth tended to underestimate the secondary forest age when compared to 354 interview data – if we assume the interview-based dates are correct, this would be consistent with 355 slower than predicted growth rates when the forests were younger.

356 Such slow early-successional recovery rates could reduce the effectiveness of climate 357 mitigation strategies in regions that have a long history of human occupation, have lost most of 358 the original forest cover, and have suffered the widespread defaunation of large-bodied vertebrates 359 through hunting (Almeida et al. 2010, Moura et al. 2014). These factors reduce the dispersal and 360 colonization of forest tree species (Hooper et al. 2005), reduce carbon stocks (Bello et al. 2015) 361 and increase recovery times (Jakovac et al. 2015, Chazdon 2014, Villa et al. 2018). Moreover, it 362 is important to note that these measures of carbon recovery were observed under the climatic and 363 landscape conditions over the lifetimes of the secondary forest plots. Future recovery could be 364 slower if the last forest remnants are lost, further decreasing forest connectivity (Aguiar et al. 2016, 365 Reid et al. 2018, Matos et al. 2019), if the frequency of farm-fallow cycles increases (Jakovac et 366 al. 2015, Villa et al. 2018), or if deforestation and climate change further increase water deficits 367 (Fu et al. 2013, Spracklen et al. 2018).

368

Interesting insights can be drawn from the two predictors – the Maximum Climatological

369 Water Deficit (MCWD) and edge distance – that had no discernible impact on carbon balance. 370 First, MCWD is one of the most frequently used measures of water deficit in studies of tropical 371 forests (e.g., Malhi et al. 2009, Anderson et al. 2018, Berenguer et al. 2018, Esquivel-Muelbert et 372 al. 2017, 2019), but it appears that longer-term measures such as the 12-month SPEI index can 373 reveal ecological processes that are not influenced by inter-period water deficit maximums. 374 Second, edge effects have been detected in a wide range of contexts, including in secondary forests 375 (e.g. Magnago et al. 2017); the lack of any effect in our study could be due to the limited range of 376 distance-to-edge (62-266 m) or the possibility that all long-term study plots were under some form 377 of edge influence in this highly deforested landscape.

378

379 The recovery of biodiversity

380 Although rarefied richness exhibited greater convergence to primary forest levels (56%) 381 than carbon (41.1%), there was a near-zero increase in secondary forest rarefied richness between 382 1999 and 2017. Richness therefore responds differently to carbon recovery, as tree diversity 383 appears to have increased quickly in the initial stages of succession (see also Lennox et al. 2018) 384 before reaching a state of impeded or arrested succession (Arroyo-Rodríguez et al. 2017). This 385 slowdown could be due to the absence of diverse seed inputs and the slow generation time of trees 386 limiting the recruitment of older-growth species. It also suggests that the strong carbon-387 biodiversity relationships observed in recovering forests elsewhere (Gilroy et al. 2014, Ferreira et 388 al. 2018, Lennox et al. 2018) may not occur in older secondary forests or in highly deforested 389 landscapes. Finally, it is also likely that we are overestimating the relative recovery of biodiversity: 390 although we rarefied richness by 100 stems, which is two to five times higher than previous studies 391 assessing biodiversity recovery (Rozendaal et al. 2019), the absolute difference between primary

and secondary forest would likely be much greater if we considered larger plots that capture moreof the high beta diversity of primary forest (e.g., Solar et al. 2015).

394

395 Conclusion

396 The capacity of regenerating tropical secondary forests to sequester carbon and provide 397 habitat for tropical species has profound implications for global climate change and biodiversity 398 conservation. Our study is among the first to utilize data collected over two decades through 399 periodic resampling and finds that the ability of secondary forests to mitigate climate change and 400 limit biodiversity loss are likely to be negatively affected by increases in the rate of tropical 401 deforestation and ongoing climate change. Understanding the generality and future climate 402 sensitivity of these responses will require further investment in long term studies in human 403 modified tropical forests.

404

405 Acknowledgments

We are grateful to the projects PRODETAB, ProManejo, INOVAGRI, Rede Biomassa
Florestal, BioRed (Nerc Foundation/FAPESP, NE/N01250X/1 - BIORED), PELD-RAS (CNPq,
Proc. No. 441659/2016-0) and RESFLORA (CNPq, Proc. No. 420254/2018-8) for financial
support during the field stages. We are also grateful to CAPES (Process, no. 1661300) for the PhD
scholarship granted to the first author. We thank two anonymous reviewers who provided valuable
comments on an earlier version of the manuscript.

412

413 Literature cited

414 Aguiar, A. P. D., et al. 2016. Land use change emission scenarios: Anticipating a forest transition

- 415 process in the Brazilian Amazon. Global Change Biology 22:1821–1840.
- Almeida, A. S., et al. 2010. Nonfrontier Deforestation in the Eastern Amazon. Earth Interactions
 14:1–15.
- 418 Álvarez-Yépiz, J. C., et al. 2018. Special Issue: Resilience of tropical dry forests to extreme
- 419 disturbance events. Forest Ecology and Management 426:1–6.
- 420 Anderson, L. O., et al. 2018. Vulnerability of Amazonian forests to repeated droughts.
 421 Philosophical Transactions of the Royal Society B: Biological Sciences 373:20170411.
- 422 Anderson-Teixeira, K. J., et al. 2013. Altered dynamics of forest recovery under a changing
 423 climate. Global Change Biology 19:2001–2021.
- 424 Arroyo-Rodríguez, V., et al. 2017. Multiple successional pathways in human-modified tropical
- 425 landscapes: new insights from forest succession, forest fragmentation and landscape ecology
 426 research. Biological Reviews 92:326–340.
- Baker, J. C. A., and D. V Spracklen. 2019. Climate Benefits of Intact Amazon Forests and the
 Biophysical Consequences of Disturbance. Frontiers in Forests and Global Change 2:1–13.
- Bello, C., et al. 2015. Defaunation affects carbon storage in tropical forests. Science Advances
 1:e1501105.
- Bentos, T. V., et al. 2017. Effects of lightgaps and topography on Amazon secondary forest:
 Changes in species richness and community composition. Forest Ecology and Management
 396:124–131.
- 434 Berenguer, E., et al. 2018. Tree growth and stem carbon accumulation in human-modified
 435 Amazonian forests following drought and fire. Philosophical Transactions of the Royal Society
- 436 B: Biological Sciences 373:20170308.
- 437 Bonner, M. T. L., et al. 2013. A meta-analytical global comparison of aboveground biomass

- 438 accumulation between tropical secondary forests and monoculture plantations. Forest Ecology439 and Management 291:73–86.
- Brando, P. M., et al. 2019. Droughts, Wildfires, and Forest Carbon Cycling: A Pantropical
 Synthesis. Annual Review of Earth and Planetary Sciences 47:555–581.
- 442 Bretfeld, M., et al. 2018. Plant water use responses along secondary forest succession during the
- 443 2015-2016 El Niño drought in Panama. New Phytologist 219:885–899.
- Brienen, R. J. W., et al. 2015. Long-term decline of the Amazon carbon sink. Nature 519:344–
 348.
- Chave, J., et al. 2009. Towards a worldwide wood economics spectrum. Ecology Letters 12:351–
 366.
- Chave, J., et al. 2014. Improved allometric models to estimate the aboveground biomass of tropical
 trees. Global Change Biology 20:3177–3190.
- 450 Chazdon, R. L. 2014. Second growth: the promise of tropical forest regeneration in an age of451 deforestation. University of Chicago Press.
- 452 Chazdon, R. L., et al. 2005. Effects of climate and stand age on annual tree dynamics in tropical
 453 second-growth rain forests. Ecology 86:1808–1815.
- 454 Chazdon, R. L., et al. 2007. Rates of change in tree communities of secondary Neotropical forests
- following major disturbances. Philosophical Transactions of the Royal Society B: Biological
 Sciences 362:273–289.
- 457 Chazdon, R. L., et al. 2009. The potential for species conservation in tropical secondary forests.
 458 Conservation Biology 23:1406–1417.
- 459 Da Silva Castro, R. M., et al. 2013. Influence of the rainfall in the content of nutrients in litter in
- 460 agroforestry systems managed with burning and without burning in Amazon. Agricultural

461 Sciences 04:26–36.

- 462 Eichenberg, D., et al. 2013. Shifts in community leaf functional traits are related to litter
 463 decomposition along a secondary forest succession series in subtropical China. Journal of Plant
 464 Ecology 8:401–410.
- 465 Elias, F., et al. 2018. Idiosyncratic soil-tree species associations and their relationships with
 466 drought in a monodominant Amazon forest. Acta Oecologica 91:127–136.
- Esquivel-Muelbert, A., et al. 2017. Seasonal drought limits tree species across the Neotropics.
 Ecography 40:618–629.
- 469 Esquivel-Muelbert, A., et al. 2019. Compositional response of Amazon forests to climate change.
- 470 Global Change Biology:1–19.
- 471 Feeley, K. J., et al. 2007. Decelerating growth in tropical forest trees. Ecology Letters 10:461–469.
- 472 Ferreira, J., et al. 2018. Carbon-focused conservation may fail to protect the most biodiverse
 473 tropical forests. Nature Climate Change 8:744–749.
- 474 França, F., et al. 2016. Do space-for-time assessments underestimate the impacts of logging on
 475 tropical biodiversity? An Amazonian case study using dung beetles. Journal of Applied Ecology
 476 53:1098–1105.
- Fu, R., et al. 2013. Increased dry-season length over southern Amazonia in recent decades and its
 implication for future climate projection. Proceedings of the National Academy of Sciences
 110:18110–18115.
- 480 Gilroy, J. J., et al. 2014. Cheap carbon and biodiversity co-benefits from forest regeneration in a
 481 hotspot of endemism. Nature Climate Change 4:503–507.
- 482 Greenwood, S., et al. 2017. Tree mortality across biomes is promoted by drought intensity, lower
- 483 wood density and higher specific leaf area. Ecology Letters 20:539–553.

- 484 Hooper, E., et al. 2005. Barriers to forest regeneration of deforested and abandoned land in
 485 Panama. Journal of Applied Ecology 42:1165–1174.
- Houghton, R. A., and A. A. Nassikas. 2018. Negative emissions from stopping deforestation and
 forest degradation, globally. Global Change Biology 24:350–359.
- 488 INMET. 2018. BDMEP Banco de Dados Meteorológicos para Ensino e Pesquisa.
 489 http://www.inmet.gov.br/portal/index.php?r=bdmep/bdmep.
- Jakovac, C. C., et al. 2015. Loss of secondary-forest resilience by land-use intensification in the
 Amazon. Journal of Ecology 103:67–77.
- Johnson, E. A., and K. Miyanishi. 2008. Testing the assumptions of chronosequences in
 succession. Ecology Letters 11:419–431.
- Lennox, G. D., et al. 2018. Second rate or a second chance? Assessing biomass and biodiversity
 recovery in regenerating Amazonian forests. Global Change Biology 24:5680–5694.
- 496 Lewis, S. L., et al. 2019. Restoring natural forests is the best way to remove atmospheric carbon.
 497 Nature 568:25–28.
- 498 Lloyd, J., and G. D. Farquhar. 2008. Effects of rising temperatures and [CO₂] on the physiology
- of tropical forest trees. Philosophical Transactions of the Royal Society B: Biological Sciences
 363:1811–1817.
- Lopez-Gonzalez, G., et al. 2011. ForestPlots.net: A web application and research tool to manage
 and analyse tropical forest plot data. Journal of Vegetation Science 22:610–613.
- Lunn, D. J., et al. 2009. Generic reversible jump MCMC using graphical models. Statistics and
 Computing 19:395–408.
- Magnago, L. F. S., et al. 2017. Do fragment size and edge effects predict carbon stocks in trees
 and lianas in tropical forests? Functional Ecology 31:542–552.

- Malhi, Y., et al. 2009. Exploring the likelihood and mechanism of a climate-change-induced
 dieback of the Amazon rainforest. Proceedings of the National Academy of Sciences
 106:20610-20615.
- 510 Malhi, Y., et al. 2014. Tropical Forests in the Anthropocene. Annual Review of Environment and

511 Resources 39:125–159.

- 512 MapBiomas. 2019. Projeto de Mapeamento Anual da Cobertura e Uso do Solo no Brasil 513 MapBiomas v.3.1. http://mapbiomas.org/
- Markesteijn, L., et al. 2011. Hydraulics and life history of tropical dry forest tree species:
 coordination of species' drought and shade tolerance. New Phytologist 191:480–495.
- 516 Martínez-Ramos, M., et al. 2018. Effects of long-term inter-annual rainfall variation on the 517 dynamics of regenerative communities during the old-field succession of a neotropical dry
- 518 forest. Forest Ecology and Management 426:91–100.
- Matos, F. A. R., et al. 2019. Secondary forest fragments offer important carbon-biodiversity cobenefits. Global Change Biology:gcb.14824.
- 521 Maza-Villalobos, S., et al. 2013. Effects of ENSO and temporal rainfall variation on the dynamics
- of successional communities in old-field succession of a tropical dry forest. PLoS ONE8:e82040.
- Mendivelso, H. A., et al. 2014. Time-dependent effects of climate and drought on tree growth in a
 Neotropical dry forest: Short-term tolerance vs. long-term sensitivity. Agricultural and Forest
 Meteorology 188:13–23.
- Moura, N. G., et al. 2014. Two hundred years of local avian extinctions in Eastern Amazonia.
 Conservation Biology 28:1271–1281.
- 529 Norden, N., et al. 2015. Successional dynamics in Neotropical forests are as uncertain as they are

- 530 predictable. Proceedings of the National Academy of Sciences 112:8013–8018.
- 531 Phillips, O. L., et al. 2009. Drought Sensitivity of the Amazon Rainforest. Science 323:1344–1347.
- 532 Poorter, L., et al. 2015. Diversity enhances carbon storage in tropical forests. Global Ecology and
- 533 Biogeography 24:1314–1328.
- 534 Poorter, L., et al. 2016. Biomass resilience of Neotropical secondary forests. Nature 530:211–214.
- Poorter, L., et al. 2019. Wet and dry tropical forests show opposite successional pathways in wood
 density but converge over time. Nature Ecology & Evolution 3:928–934.
- Reid, J. L., et al. 2018. The ephemerality of secondary forests in southern Costa Rica. Conservation
 Letters:e12607.
- 539 Requena Suarez, D., et al. 2019. Estimating aboveground net biomass change for tropical and
- subtropical forests: Refinement of IPCC default rates using forest plot data. Global ChangeBiology:gcb.14767.
- Rowland, L., et al. 2014. The sensitivity of wood production to seasonal and interannual variations
 in climate in a lowland Amazonian rainforest. Oecologia 174:295–306.
- Rozendaal, D. M. A., et al. 2019. Biodiversity recovery of Neotropical secondary forests. Science
 Advances 5:eaau3114.
- 546 Rutishauser, E., et al. 2015. Rapid tree carbon stock recovery in managed Amazonian forests.
 547 Current Biology 25:R775–R792.
- 548 Schwalm, C. R., et al. 2017. Global patterns of drought recovery. Nature 548:202–205.
- 549 Slot, M., et al. 2016. Temperature response of CO₂ exchange in three tropical tree species.
 550 Functional Plant Biology 43:468.
- 551 Solar, R. R. de C., et al. 2015. How pervasive is biotic homogenization in human-modified tropical
- forest landscapes? Ecology Letters 18:1108–1118.

- Spracklen, D. V., et al. 2018. The Effects of Tropical Vegetation on Rainfall. Annual Review of
 Environment and Resources 43:193–218.
- Sullivan, M. J. P., et al. 2018. Field methods for sampling tree height for tropical forest biomass
 estimation. Methods in Ecology and Evolution 9:1179–1189.
- Thomson, J. R., et al. 2010. Bayesian change point analysis of abundance trends for pelagic fishes
 in the upper San Francisco Estuary. Ecological Applications 20:1431–1448.
- 559 Uriarte, M., et al. 2016. Impacts of climate variability on tree demography in second growth
- 560 tropical forests: the importance of regional context for predicting successional trajectories.
- 561 Biotropica 48:780–797.
- Vicente-Serrano, S. M., et al. 2010. A Multiscalar Drought Index Sensitive to Global Warming:
 The Standardized Precipitation Evapotranspiration Index. Journal of Climate 23:1696–1718.
- Vieira, I. C. G., and T. A. Gardner. 2012. Florestas secundárias tropicais: ecologia e importância
 em paisagens antrópicas. Bol. Mus. Para. Emilio Goeldi. Ciências Naturais 7:191–194.
- 566 Vieira, I. C. G., C. Uhl, and D. Nepstad. 1994. The role of the shrub Cordia multispicata Cham. as
- a "succession facilitator" in an abandoned pasture, Paragominas, Amazônia. Vegetatio 115:91–
 99.
- Villa, P. M., et al. 2018. Intensification of shifting cultivation reduces forest resilience in the
 northern Amazon. Forest Ecology and Management 430:312–320.
- 571

572 Figure Legends

Figure 1: The recovery of secondary forests in Bragança. Carbon (A) and rarefied species richness
(B) and the relationship between carbon and rarefied richness (C) between 1999 and 2017 in the
twelve long-term study plots. Points show results at each census. Points and lines are color-coded

576 by plot. The black line shows the median trend. The grey band shows the 95% credible interval.

Figure 2: Climate and secondary forest carbon dynamics. The marginal effect of the standardized precipitation-evapotranspiration index (SPEI; A, C, E, G) and the maximum dry season temperature (Max temperature; B, D, F, H) on carbon balance (A-B), growth (C-D), mortality (E-F) and recruitment (G-H) in the twelve long-term secondary forest study plots. Lines show the median relationship. Bands show the 95% credible interval. A statistical association was assumed for posterior probabilities of a non-zero effect (Pr(>0)) > 0.75.

Figure 3: Estimates of carbon accumulation rates from previous studies compared to those from the twelve long-term and four extra secondary forest study plots. Bars show median values (or the mean where the median was unavailable). Error bars show 95% confidence/credible intervals. The Poorter et al. (2016) tropical secondary forest estimate uses only Neotropical plots. For the Requena Suarez et al. (2016) study, we used their estimates from tropical rainforests in the Americas. The Rutishauser et al. (2015) and Brienen et al. (2015) studies are from the Amazon.

589 Figure 4. (A) Extrapolation of past and future carbon levels for the 12 long-term study plots. The 590 white line shows median carbon; the grey band shows a possible range. These estimates were 591 derived by fitting rational functions to the median and 95% credible lower and upper bounds of 592 carbon recovery shown in Fig. 1A. The inset shows the data. Points shows carbon levels at each 593 census. Points and lines are color coded by plot. The black dashed line shows mean primary forest 594 carbon. (B) Comparison of carbon recovery trajectories from the present study and Requena 595 Suarez et al. (2019). For this comparison, we used the Requena Suarez et al. (2019) 95% 596 confidence interval carbon accumulation rates for <20- and >20-year-old tropical American 597 secondary rainforests. For the present study, we assumed the 12 long-term study plots were 50

years old, meaning that the first census occurred (in 1999) at age 32 (shown by the vertical dashed
line). Carbon recovery rates post this age were taken from the 95% credible lower and upper
bounds (Fig. 1A). Rates prior to this age were extrapolated linearly from the 32-year-old 95%
credible lower and upper bound values.