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Risk Minimizing Strategies for Revenue Management Problems with Target Values

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Abstract

Consider a risk-averse decision maker in the setting of a single-leg dynamic revenue management problem with revenue controlled by limiting capacity for a fixed set of prices. Instead of focussing on maximizing the expected revenue, the decision maker has the main objective of minimizing the risk of failing to achieve a given target revenue.

Interpreting the revenue management problem in the framework of finite Markov decision processes, we augment the state space of the risk-neutral problem definition and change the objective function to the probability of failing a certain specified target revenue. This enables us to obtain a dynamic programming solution which generates the policy minimizing the risk of not attaining this target revenue. We compare this solution with recently proposed risk-sensitive policies in a numerical study and discuss advantages and limitations.

Keywords: capacity control, revenue management, multi-period, risk, target level criterion

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1 Introduction

Revenue management systems have become a standard tool in various industries beyond the original airline industry. These newer industries range from cruise lines, rental cars, and media advertising to medical services and event management, see e.g. Talluri and van Ryzin (2005) and Chiang et al. (2007).

We consider a typical revenue management model: a firm operating in a monopolistic setting offering multiple products. These products consume a fixed resource of a limited capacity. The firm sells the products over a finite time horizon. At the end of this time, the salvage value of the resource is assumed to be zero.

The firm can influence its revenue stream by allocating capacity to different classes of demand. Its objective is to find a policy which optimizes an objective function. Normally, this objective function is risk-neutral, and the policy is chosen to maximize expected revenue. Such a risk-neutral objective can be motivated by the law of large numbers if the revenue process repeats itself very often, e.g. a daily operating airline flight connection.

However, a risk-neutral policy might not be requested under all scenarios and a riskaverse policy might be advantageous for the decision maker.

Lancaster (2003) remarks that a risk-neutral model is often not sufficient, even in the airline industry, as a stable revenue might be preferable due to financial constraints.

In practice, decision makers present some level of risk aversion in revenue management, as mentioned by Bitran and Caldentey (2003). Weatherford (2004) reports the same experience. He observed that airline analysts feel uncomfortable with recommendations of their (risk-neutral) revenue management systems, in particular while waiting for the high-fare passengers a few days before flight departure.

Levin et al. (2008) provide arguments to employ risk-averse policies in scenarios with only a small number of reiterations. They consider an event promoter who has high fixed costs which must be recovered in order to prevent a possible loss. Minimum targets can not be ignored in such a case. They present further scenarios which require strategical and financial circumstances which necessitate risk considerations.

Taking into account costs of price changes, Koenig and Meissner (2008) illustrate that risk considerations can make a difference when choosing a revenue management strategy.

In recent papers by Barz and Waldmann (2007), Huang and Chang (2009) and Koenig and Meissner (2009), risk-neutral and risk-sensitive policies are analyzed. The results show that an appropriate risk-averse policy can be selected if the decision maker knows the parameters representing his level of risk aversion. Such parameters have to be determined, which is something that is not straightforward in either of the published approaches, whether the underlying concept is a exponential utility function or a discount factor relaxing an optimality condition. Usually, the parameters have to be estimated by running numerical experiments and evaluating risk measures, such as mean-variance or conditional-value-at-risk, on the results.

Thus, we propose using the target percentile risk measure, discussed by Boda and Filar (2006), as the object function. The target percentile risk measure computes the probability of the return failing to achieve a previous given fixed target. There are several advantages of using this measure.

First, one important structural property is its time consistency. It says that optimality of decisions should only consider the future. Time consistency is a desirable property for multi-period risk measures as it allows its use in dynamic programming, as shown by the works of Boda and Filar (2006), Artzner et al. (2007), Shapiro (2009). Second, it does not assume a special kind of revenue distribution, as it measures the percentile of the given target. Third, it is easily interpreted by practitioners and does not require a risk sensitivity parameter which might be difficult to assess. Fourth, numerical computation schemes are available as described by Wu and Lin (1999). Fifth, Boda and Filar (2006) show that multi-stage versions for the well-established risk measures value-at-risk and conditional-value-at-risk can be developed using the target percentile.

The structure of the paper is as follows. We look at related literature in Section 2. In Section 3, we describe our model as a Markov decision process and its extension to apply the target percentile risk measure. This section also contains some implementation details. Section 4 shows numerical results of our approach and provides a comparison with results of other approaches. Finally, we conclude the paper in Section 5.

2 Related Work

Related work can be divided into three classes" risk and its measurement, in particular multi-period risk measures; Markov decision processes (MDPs) and dynamic programming considering risk in general; revenue management and risk. As literature about risk and its measures is vast, we will only point out some recent publications which we classify to be of interest in context of this paper. We then proceed to discuss links to MDP literature, which provides the basis of our approach.

Considering risk and quantifying it inevitably leads to expected utility theory as developed by von Neumann and Morgenstern (1947) and the mean-variance measure introduced by Markowitz (1952). However, both approaches have drawbacks, e.g. utility theory requires knowledge about one's utility function, and mean-variance theory is only useful with bell-shaped probability distributions and cannot properly used with every probability distribution. Thus, other risk measures were investigated and have become more attractive, such as value-at-risk (V@R) and conditional-value-at-risk (CV@R), see Artzner et al. (1999) or Rockafellar and Uryasev (2000). Both measures compute a value for the downside risk with respect to a given confidence value and can be interpreted as kinds of probabilistic constraints. The latter measure has mathematical properties defining its coherence which provide advantages when compared to the former. However, the disadvantage of these measures is they are only suitable for a single-stage decision. They are missing the timeconsistency property and, thus, are not suitable for multi-period decision making. This has been recognized, and several authors, such as Boda and Filar (2006), Artzner et al. (2007), and Shapiro (2009), investigate time-consistent risk measures. Such measures are feasible as objective functions for dynamic programming. Hence, they are appropriate for use in revenue management models employing the Bellman equation.

As many problems can be modeled as MDPs (cf. White, 1993; Puterman, 2005), many authors have investigated risk issues in this context. A good overview about risk in MDPs is provided by White (1988). Thus, we refer to this publication for a more complete overview and point out only briefly the general concepts here. One popular stream of implementing risk in MDPs, which goes back to Howard and Matheson (1972), is using utility functions for computing risk-sensitive policies. Another stream focuses on the variance and probabilistic

constraints as risk measures, beginning with the early works of White (1974), Sobel (1982), and ongoing with recent works such as of Sladky and Sitar (2008) and Defourny et al. (2008). We are interested in the target percentile risk measure which is analyzed by Bouakiz and Kebir (1995), Wu and Lin (1999). Their works builds the basis for the time-consistent dynamic risk measures proposed by Boda and Filar (2006). The main idea is the extension of the state space in order to use one, or possibly more, variables to keep track of the history.

Most revenue management models use a risk-neutral objective function. We refer to the work of Talluri and van Ryzin (2005) for an overview of these kinds of models. The risk-neutral model of dynamic capacity control, which we consider here, was introduced by Lee and Hersh (1993). The corresponding Markov decision process is described by Lautenbacher and Stidham (1999).

The approaches for incorporating risk in revenue management models are analogous to the general decision making under risk: expected utility theory, mean-variance considerations, probabilistic constraints such as V@R.

Expected utility theory as an element for reflecting risk in revenue management is recommended by Weatherford (2004). He states that the assumption of risk neutrality is not given for many practical scenarios and proposes expected utility theory as a risk-averse solution. Instead the well-adopted (risk-neutral) expected marginal seat revenue model (EMSR), standard algorithms introduced by Beloba (1989), and the expected marginal seat utility (EMSU) heuristic can reflect risk-sensitivity for decision making.

Recent works of Barz and Waldmann (2007) and Feng and Xiao (2008) are employing expected utility theory, too. Both papers support the application of an exponential utility function to account for risk aversion. Barz and Waldmann (2007) use the Markov decision process formulation of static and dynamic capacity control models, whereas Feng and Xiao (2008) provides closed form solutions from a more general point of view.

Finally, Lim and Shanthikumar (2007) apply the equivalence of robust and risk-sensitive control with an exponential utility function to dynamic pricing.

As the first revenue management model with risk considerations, the model of Feng and Xiao (1999) uses variance as risk measure, in particular, the variance of sales due to price changes. In order to integrate risk into their objective function, they combine expected revenue with a weighted penalty function for the sales variance. The risk sensitivity of the decision maker can be adjusted by the weighting.

With a revenue management application in traffic and networks, Mitra and Wang (2005) compare mean-variance, mean-standard-deviation and mean-CV@R for building an objective function. Their final choice is a mean-standard-deviation formulation. They demonstrate the influence of risk-sensitivity by the efficient frontier for truncated Gaussian demand distribution.

In an application for hotel revenue management, Lai and Ng (2005) formulate a robust optimization model for mean versus average deviation.

Recently, Huang and Chang (2009) presented a risk-sensitive modification of the optimality condition for the dynamic capacity control model and investigated their method by measuring mean versus standard deviation in simulation runs. They offer a ranking of their risk-sensitive policies using as the Sharpe ratio revenue per unit of risk divided by standard deviation.

Illustrating the vulnerability of risk-neutral revenue management due to demand forecast inaccuracy, Lancaster (2003) recommends a relative revenue per available seat mile at risk metric, which integrates risk measurement with the V@R metric. This metric is the expected maximum of underperformance over a time horizon at a choice confidence level.

Risk sensitivity is incorporated by Levin et al. (2008) into a dynamic pricing model of perishable products. Their objective function consists of maximum expected revenue constrained by a desired minimum level of revenue with minimum acceptable probability. This constraint is similar to a V@R formulation. Principally, their model extends a dynamic pricing risk-neutral MDP with a further state for already gained revenue.

Using both risk measures standard deviation and CV@R, Koenig and Meissner (2008) compare the suitability of two different pricing strategies considering the cost of price changes. In a further paper, Koenig and Meissner (2009) evaluate a range of risk-sensitive policies for the dynamic capacity control model.

3 Description of Model

In the following, we describe the dynamic capacity control problem as a Markov decision process in a similar way as previously done by Lautenbacher and Stidham (1999) and Barz and Waldmann (2007). This model is then expanded in the state space in order to become a model which allows the application of a risk-minimizing policy. We follow the approach of Wu and Lin (1999) here. Our objective function is the target percentile dynamic risk measure. Finally, we point out some aspects for implementation of this approach.

3.1 Markov Decision Process for Dynamic Capacity Control Model

We consider the capacity control model stated by Lee and Hersh (1993), which is often referred to as dynamic capacity control. Although originally developed for airline revenue management, it can be transferred to other industries. We describe the model in terms of its original airline revenue management context in order to be more intuitive.

We assume that the booking requests follow a Poisson arrival process. Thus, the booking period for a single-leg flight is separated into N decision periods, in such a way that the probability of more than one request can be ignored. The decision periods are denoted by $n \ in\{0,...,N\}$. Further, there are k booking classes with fares F_i , $F_1 > F_2 > ... > F_k$ and $F = \{F_1,...,F_k\}$. The probability of a request for fare class i in decision period n is given by $p_{n,i}^r$. Further, we set the probabilities for n = 0 to zero for all fare classes: $p_{0,i}^r = 0$; this step just supports our model setting as the last decision will be made at time n = 0. The probability of no request at all is $p_{n,0}^r = 1 - \sum_{i=1}^k p_{n,i}^r$. The initial capacity of seats is given by C. The remaining seats are given by $c \le C$ in a time period.

We have a finite-state, discrete-time, Markov decision process $\Gamma = (S, A, R, P)$ with countable state space *S* and action space *A*. Further, *R* denotes the reward set and *P*, the set of transition probabilities. Time runs in discrete steps from $n \in \{0, 1, ..., N\}$ and represents the remaining time before flight departure.

The state space *S* contains all possible configurations of remaining capacity *c* and request for a fare class *i*. Thus $S = \{0, 1, ..., C\} \times \{0, 1, ..., k\}$ and a state $(c, i) \in S$ says that we have *c* seats left and a request for fare class *i*. We set the fare class 0 with fare $F_0 = 0$, as is often common.

Our action space A(c, i) corresponds to the "reject" and "accept" decisions for a given state. We have $A(c, i) = \{0, 1\} \forall (c, i) \in S | i > 0$ and $A(c, 0) = \{0\}$ to only allow the accepting and rejecting of seats at the valid fare prices and not for the artificial class i = 0. For each $s \in S$, the action space A(s) is finite.

Let *R* be the set of rewards (fares) when accepting one booking. Rewards are denoted by $r_n(s, a) \in R$ with $s \in S$, $a \in A$ and $r_n((c, i), a) = aF_i$ for n, c > 0 and zero otherwise.

The transition probabilities $p \in P$ are defined for states $(c, i), (c - a, j) \in S$ with $a \in A$ by $p_n((c - a, j)|(c, i), a) = p_{n,i}^r$ for n = N, N - 1, ..., 0, and 0 otherwise.

A decision maker decides on a sequence of rules $a_n = d_n(c_n, i_n)$, which determine a policy $\pi = \{d_n, d_{n-1}, \dots, d_1\}$. Thus, one of the possible actions is chosen by accepting or rejecting a booking request for each state (c_n, i_n) .

Now let $\rho_N^{\pi}(c, i) = \sum_{n=0}^N r_n$ denote the random variable of the gained revenue for a particular policy π beginning with capacity c and request i at N remaining time steps. Its expected revenue is given by

$$v_N^{\pi}(c,i) = \mathcal{E}_{\pi}\left[\rho_N^{\pi}(c,i)\right] = \mathcal{E}_{\pi}\left[\sum_{n=1}^N r_n((c_n,i_n),d_n(c_n,i_n)) + r_0(c_0,i_0)\right].$$

The maximal expected revenue and an associated policy can be computed by the Bellman equation for this problem. However, we are interested in a policy which minimizes the time-consistent dynamic risk measure of not achieving a target revenue x in the accumulated return.

3.2 Expanded Markov Decision Process for Minimizing Risk of Failing Target

We are interested in minimizing the risk of not attaining a specified target revenue x for the dynamic capacity control model. Thus, we want to find a policy π which minimizes the objective function representing the probability of not achieving a previous specified target value x. In order to derive this objective function, we follow the approaches mentioned by White (1988), Wu and Lin (1999) and Boda and Filar (2006) and expand the Markov decision process Γ by a larger state space. The extended Markov decision process $\tilde{\Gamma}$ is similar to Γ . It consists of $\tilde{\Gamma} = (\tilde{S}, \tilde{A}, \tilde{R}, \tilde{P}) = (\tilde{S}, A, R, P)$, as described below. The state space *S* is replaced by the new state space $\tilde{S} = S \times R$ with elements ((c, i), x). It consists of states of the configurations of remaining capacity *c* and a request for fare class *i*, and additionally, a revenue target *x*. All state variables are updated over time, e.g. the revenue target *x* decreases by the realized fare price in accordance with decrementing *c* by selling a seat.

The action space \tilde{A} is generated from action state A by $\tilde{A}((c, i), x) = A(c, i), \forall ((c, i), x)$ and, thus, $\tilde{A} = \bigcup_{(c,i),x) \in \tilde{S}} \tilde{A}((c,i), x) = \bigcup_{(c,i) \in S} A(i) = A$.

In a similar way, the reward set \tilde{R} is build from R. For $\tilde{s} \in \tilde{S}$, $a \in A$, the reward $\tilde{r}_n(\tilde{s}, a) \in \tilde{R}$ is $\tilde{r}_n((c, i), x, a) = aF_i$ for c, i > 0 and zero otherwise. Thus, $\tilde{R} = R$.

As well, $\tilde{P} = P$, as the transition probabilities \tilde{P} are determined by P. We have $\tilde{p} \in \tilde{P}$ and, with states $((c, i), x), ((c - a, j), x - aF_i) \in \tilde{S}$ and $a \in A$, the transition probability is given by $\tilde{p}_n((c - a, j), x - aF_i)|((c, i), x), a) = p_{n,j}^r$ for n = N, N - 1, ..., 0 and else 0.

We are interested in the probability that our obtained total revenue does not attain a target x. Let the set of deterministic Markovian policies be $\tilde{\Pi}$ and let the random variable for the cumulative gained reward, applying policy $\tilde{\pi} \in \tilde{\Pi}$ beginning with capacity c, request i, remaining time steps N, and target x, be $\tilde{\rho}_N^{\pi}((c, i), x) = \sum_{n=0}^N \tilde{r}_n$. For the policy $\tilde{\pi}$, the target percentile risk measure is defined as

$$V_N^{\hat{\pi}}((c,i),x) := \mathcal{P}\left(\tilde{\rho}_N^{\pi}((c,i),x) \le x\right),\tag{1}$$

where \mathcal{P} denotes a probability. The time consistency property of the target percentile risk measure can be shown as demonstrated by Boda and Filar (2006).¹

Thus, we are looking now for an optimal policy $\tilde{\pi}^*$ for each objective function $V_N^{\tilde{\pi}}((c, i), x)$ that minimizes the risk of failing target x:

$$\tilde{\pi}^* = \underset{\tilde{\pi} \in \tilde{\Pi}}{\arg\min} \left\{ V_N^{\tilde{\pi}}((c,i),x) | ((c,i),x) \in \tilde{S}, n \ge 1 \right\}.$$
(2)

The associated target percentile (minimum risk level for *x*) is denoted $V_N^{\tilde{\pi}^*}$.

Following Wu and Lin (1999) and Boda and Filar (2006), we can derive the following dynamic programming equations for computation of the minimum target percentile $V_N^{\tilde{\pi}^*}$

¹We have non-stationary transition probabilities in our model, as opposed to the assumption of Boda and Filar (2006). Nevertheless, their approach can be used. Our non-stationary case could be transformed to a stationary one by a state augmentation with n which would resolve the dependence of the probabilities on n.

for $x \in \mathcal{R}$, $\forall (c, i) \in S$:

$$V_{0}^{\tilde{\pi}^{*}}((c,i),x) = \begin{cases} 0 & x > 0 \\ 1 & \text{otherwise,} \end{cases}$$
$$V_{n}^{\tilde{\pi}^{*}}((c,i),x) = \min_{a \in A} \left\{ \sum_{j \in S} p_{n-1,j}^{r} V_{n-1}^{\tilde{\pi}^{*}}((c-a,j),x-aF_{i}) \right\}, n \in [1...N].$$
(3)

Note that the target percentile for the whole process is given in $V_{N+1}^{\tilde{\pi}^*}(c, i, x)$ as the process must be entered correctly; in $V_N^{\tilde{\pi}^*}(c, i, x)$ we already know the requested class *i* at time *N*.

The optimal policy $\tilde{\pi}^*$ can be computed from the minimum target percentile $V_N^{\tilde{\pi}^*}$ by Equation 2. It should be pointed out that the optimal policy describes the best way to only obtain the target percentile. This means that if we have in some state achieved the target, the following states are arbitrarily chosen by the objective function. In practice, the policy for the ongoing states should be optimized under another criterion, such as the expected revenue. Furthermore, if the target can never be obtained in the given setting, all policies are equally improper and no optimal target percentile policy exists.

Example

In order to illustrate the method, we can give a very simple example. Consider only two classes with fares $F_1 = 200$; $F_2 = 100$, two remaining time periods N = 2, only one seat left C = 1, and the probabilities for arrivals $p_{1,1} = 0.10$, $p_{1,2} = 0.15$, $p_{2,1} = p_{2,2} = 0.20$. Thus, for example, a request of fare 2 in period 1 before departure is 15 percent. We have only a few scenarios in this setting: if a request for a distinct fare class comes in period 2 before departure, we can accept it or reject this fare class and then wait for possible arrivals in the last period and, if they appear, accept. It is easy to see, that the policy which always accepts (expected revenue of 81) is better off when compared with others. However, consider that now we want the best policy for a target value of 200. The expected revenue maximizing policy fails that target with probability of 0.74. A better choice for this target would be only acceptance of the highest fare class, a policy which fails only with a likelihood of 0.72. The computation using the proposed method is shown in Figure 1. Note that the transition

probabilities of the last period n = 0 are zero for all classes except for the artificial class 0, where they are one.

$V_0(c, i, x)$	=1, x > 0,
$V_0(c, i, x)$	$=0, x \le 0$
$V_1(1, 0, 200)$	$=1 \cdot V_1(1,0,200) + 0 \cdot V_1(1,1,200) + 0 \cdot V_1(1,2,200)$
$V_1(1, 1, 200)$	=1 = min{1 · $V_0(1, 0, 200)$ +0 · $V_0(1, 1, 200)$ +0 · $V_0(1, 2, 200)$, 1 · $V_0(0, 0, 0)$ +0 · $V_0(0, 1, 0)$ +0 · $V_0(0, 2, 0)$ }
$V_1(1, 2, 200)$	=0 = min{1 · V ₀ (1,0,200) +0 · V ₀ (1,1,200) +0 · V ₀ (1,2,200), 1 · V ₀ (0,0,100) +0 · V ₀ (0,1,100) +0 · V ₀ (0,2,100)} =1
$V_2(1, 0, 200)$	$= p_{1,0}V_1(1,0,200) + p_{1,1}V_1(1,1,200) + p_{1,2}V_1(1,2,200)$ = 0.90
$V_2(1, 1, 200)$	$=\min\{p_{1,0}V_1(1,0,200) + p_{1,1}V_1(1,1,200) + p_{1,2}V_1(1,2,200), \\ p_{1,0}V_1(0,0,0) + p_{1,1}V_1(0,1,0) + p_{1,2}V_1(0,2,0)\}$
$V_2(1, 1, 200)$	$=\min\{p_{1,0}V_1(1,0,200) + p_{1,1}V_1(1,1,200) + p_{1,2}V_1(1,2,200), \\ p_{1,0}V_1(0,0,100) + p_{1,1}V_1(0,1,100) + p_{1,2}V_1(0,2,100)\} \\ = 0.90$
$V_3(1,0,200)$	$=p_{2,0}V_2(1,0,200) + p_{2,1}V_2(1,1,200) + p_{2,2}V_2(1,2,200)$ $=0.72$

Figure 1: Exemplary computation of $F_N^{\tilde{\pi}^*}(c, i, x)$, superscript omitted.

Implementation Details

The dynamic programming formulation given in Equation 3 is inefficient for implementation. We can point out two remarks in order to obtain a better implementable approach.

First, we can apply an often used transformation of the dynamic programming formulation, allowing us to scale down the state space. Thus, introducing the operator $T_n(c, x) := \sum_{i=0}^{i_k} p_{n,i}^r V_n(c, i, x)$ helps by reducing the state space by variables representing the fare class of an arrival. Defining $W_n(c, x) := T_n(c, x)V_n(c, i, x)$, we transform Equation 3, as follows, for $x \in \mathcal{R}, c \in \{0, ..., C\}$:

$$W_{0}^{\tilde{\pi}^{*}}(c,x) = \begin{cases} 0 & x > 0 \\ 1 & \text{otherwise,} \end{cases}$$
$$W_{n}^{\tilde{\pi}^{*}}(c,x) = T_{n}(c,x)V_{n}^{\tilde{\pi}^{*}}(c,i,x)$$
$$= \sum_{i=0}^{i_{k}} p_{n,i}^{r} \min_{a \in A} \left\{ W_{n}^{\tilde{\pi}^{*}}(c-a,x-aF_{i}) \right\}.$$
(4)

Second, the computation of all possible cumulative rewards given by the variable x is very impractical and should be done on a suitable grid for larger problems, as described in the works of Wu and Lin (1999) and Boda et al. (2004). We refer to the latter publication for a deeper analysis of this approximative dynamic programming solution. Instead of taking into account every value which x can take while solving the problem, we can use an interval between 0 and the initial x. E.g., if we have m + 1 values for this interval, the interval could be $[y_0, y_1, y_2, \ldots, y_m] = [0, \frac{x}{m}, \frac{2x}{m}, \ldots x]$. The computation of $W_n^{\tilde{\pi}^*}$ proceeds now only on these samples, and we obtain a grid of values $\{(y_0, W_n^{\tilde{\pi}^*}(c, y_0), \ldots, (y_m, W_n^{\tilde{\pi}^*}(c, y_m))\}$ for $c \in \{0, \ldots, C\}$. The program fills the values on the grid and rounds occurring values to the upper value: $W_n^{\tilde{\pi}^*}(c, y) = W_n^{\tilde{\pi}^*}(c, y_{j+1}) \forall y \in (y_j, y_{j+1}]$.

4 Numerical Simulation and Results

In their introductory paper about dynamic capacity control, Lee and Hersh (1993) used an example which also served for illustration in the recent papers of Barz and Waldmann (2007), Huang and Chang (2009) and Koenig and Meissner (2009). Thus, we can also demonstrate the proposed target percentile policy in the same exemplary setup.

Simulation Setup

There are N = 30 number of time periods before departure, and the initial number of seats is C = 10. The four fare classes are $F_1 = 200$, $F_2 = 150$, $F_3 = 120$, $F_4 = 80$. The probabilities for a request of a certain fare class in a certain time period are shown in Table 1.

In order to see how the target percentile policy works, we conducted an experiment with a 10,000 sample run. Random arrivals were simulated in a Monte Carlo manner using the

i	F_i	$1 \le n \le 4$	$5 \le n \le 11$	$12 \le n \le 18$	$19 \le n \le 25$	$26 \le n \le 30$
1	200	0.15	0.14	0.10	0.06	0.08
2	150	0.15	0.14	0.10	0.06	0.08
3	120	0	0.16	0.10	0.14	0.14
4	80	0	0.16	0.10	0.14	0.14

Table 1: Fares and request probabilities for fare class *i* and time period *n*.

values of Table 1. When compared with other proposed policies, the random values in the setting were, of course, always the same.

A single simulation run is initialized with values for remaining seats, time periods before departure, and a policy. The policy contains for each state the acceptable fare classes. The state is described by remaining time periods, remaining seats, and remaining target value. Then, the simulation continues with loop over the time periods until the departure time zero is reached. Inside the loop, a random generator simulates requests for fare classes which are accepted if the current policy allows acceptance of the class or else rejected. An update of the state is follows: time periods are always decremented by one, seats are decremented only if fare is accepted, and target value is decremented by the gained fare price. Our policies have one more dimension than the policies of the above mentioned references. Hence, we can abandon the target percentile dimension when simulating the policies used for comparison.

Policy Illustration

Figure 2 visualizes the policy $\tilde{\pi}^*$ for the described example. We see slices through a threedimensional matrix codifying protection levels in color. Basically, this three-dimensional matrix displays the protection levels — the maximum allowed fare class — for each state (c, x) in time n with initial target of 1200. In order to use the policy, we starts in the state (30, 1200) at 30 time periods to go. This is the top corner on the right hand side of the presented box. The state at this position in the matrix has a protection level which lets one decide how to act at this point in time before departure. Only fare classes with higher prices than the associated protection level are accepted. As time marches on, one moves always one step further in direction of the time dimension to departure time zero; this is parallel to the south-west direction in the figure. The policy decides now which way to move in both other dimensions. An acceptance of a request causes a move downwards the dimension of the capacity, orthogonal downwards in the matrix. Finally, the price of an accepted fare means where to move in the target direction, the parallel the north-west direction in the figure. Thus, considering the figure, the simulation will generate random trajectories from the top corner on the right hand side to the bottom corner on the left hand side. Of course, the end of each trajectory will often be different due to the random realizations but it has to end with coordinate n = 0.



Figure 2: Target percentile policy : Protection levels of policy $\tilde{\pi}^*$ are shown for each state (c, x) with remaining capacity $0 \le c \le 10$, revenue target $0 \le x \le 1200$ in time $n, 1 \le n \le 30$. The protection levels for the four fare classes are visualized by color: class 1 (blue), class 2 (green), class 3 (orange), and class 4 (red). Only fare classes of lower price than indicated by the color are accepted in a state. Note, states which offer no optimal solution because the revenue is not obtainable, given the remaining capacity and time, are displayed in transparency.

Evaluation

As the proposed policy optimizes the target percentile, we start our evaluation with different (obtainable) target revenues, comparing the theoretical and the simulation results. As mentioned in Section 3.2, there are scenarios when target revenue is achieved but time is remaining and one or more seats are left. We present the average of remaining time and seats for such cases as well. Further, the averaged revenue is computed by switching to the risk-neutral policy when the target has been achieved. Table 2 shows the results for five different targets. The average of failed cases in the simulation is very close to the theoretical target percentile, validating that the policy does as expected.

Target value	1600	1500	1400	1300	1200	
$V_n^{\tilde{\pi}^*}$	0.734	0.528	0.336	0.183	0.088	
Simulation of target value policies						
Failed target (%)	0.745	0.537	0.347	0.184	0.087	
Revenue	1231.7	1319.7	1347.4	1342.6	1325.3	
Std. deviation	329.6	273.8	217.7	165.1	152.0	
Remaining time	0.15	1.54	3.18	5.30	7.35	
Remaining seats	0.01	0.03	0.10	0.22	0.49	
Simulation of risk-neutral policy						
Failed target (%)	0.840	0.632	0.430	0.260	0.146	
Revenue	← 1405.6 →					
Std. deviation $\leftarrow 195.5 \rightarrow$						

Table 2: Results of policy simulation for different target values. The probabilities and averages for failing to achieve a target are given (lower means better) and also, averages of remaining time and seats if target could be achieved. An expected revenue optimizing risk-neutral policies yields theoretically 1407.2. For comparison, the results of simulating the risk-neutral policy are given.

The expected revenue for the analyzed problem are 1407.2. Looking at the results of Table 2, we see that a policy which aims towards a lower target revenue than the expected value accepts an upcoming request early in time. Decisions are made soon and not postponed to later periods. This effect is well observable by the decreasing remaining time and seats, while increasing the target. Of course, policies with lower targets have a greater probability for reaching the target. It can be more easily obtained by accepting requests early, thus leaving more time for balancing against having no profitable requests in the next time periods.

The average revenues of the target policies are in each case lower than that of the riskneutral policy. The standard deviation of these revenues grows with an increasing target, although when compared with the risk-neutral case, their policies less often fail the targets. This can be explained by comparing the distribution histograms of the revenues of the policies.



Figure 3: Histogram of the distributions of gained revenue in the simulation using policies with revenue targets 1200 (left blue bars), 1400 (right red bars) and the risk-neutral policy (middle green bars).

Figure 3 shows the distribution histograms of 1000 simulation runs of three policies: one with low target 1200, one with high target 1400, and a risk-neutral one maximizing expected revenue.

The distribution associated with the low target has its peak above its target value 1200 and a positive skew. It has only small frequencies for values lower than 1200 but also for values higher than 1500, as its standard deviation from Table 2 also emphasizes. It has two peaks, the first at 1000 and the second at 1300.

The risk-neutral solution shows a negative skewed distribution with a peak at 1500 with a long tail to very low values, though some high revenues at 1800. Compared with the

policy with target 1200, its revenues are more often below 1200; however, given the revenue is greater 1200, it will be better off. Its risk of falling below 1200 remains higher than the risk of the low target policy.

The distribution of the policy with high target 1400 has a negative skew, too. As with the 1200 target policy, it has two peaks. The distribution of revenues from 600 to 1200 increases, then drops at 1300, before it peaks at 1400 and decreases until 1700. Compared with the risk-neutral counterpart, this policy shifts frequency from 1300 to 1400 revenue. The target is achieved mainly at the expense of 1300 revenue and greater than 1500 revenue. Further, it shows also higher frequencies for low revenue than both other policies. Hence, if it fails the target, there is a greater risk of obtaining only low revenue.

The histogram demonstrates that the policy with low target aims at a lower average revenue and smaller variance, but the policy with a higher target, near to the expected revenue of the risk-neutral solution, does not.

The results show that an analysis of the loss tail is important as it gives information about the probability of worst case disasters. In order to evaluate the performance of target revenue policies in more detail, we compare them with the risk-sensitive policies derived from expected utility theory, as in Barz and Waldmann (2007). We select the latter policies for comparison as they result from optimizing the dynamic capacity control model using an exponential utility² and no heuristics. Referring to the recent works of Huang and Chang (2009) and Koenig and Meissner (2009), we view the mean, standard deviation, and CV@R of the policies. The CV@R is a measure for the expected revenue given the revenue is below a certain quantile specified by a confidence level α ; it is the expected value in the $1 - \alpha$ percent of worst cases.

Table 3 compares both types of risk-sensitive policies. Beyond the mean, standard deviation, and CV@R with confidence level 95%, the probability of failing the 1000 revenue target is given. We see that the target policy for 1000 has the least risk failing it. However, it is also observable that the target policies only limit the risk of failing the certain target and do not provide more preferable results in terms of the other measures. The expected utility based policies have higher average revenue. They also have a higher CV@R than the policies with target \geq 1200. The standard deviation increases with higher target and higher

²Instead of searching the policy π for $\max_{pi} \mathcal{E}_{\pi} [\rho_N^{\pi}(c, i)]$, the expected utility approach with exponential utility uses $\max_{pi} \mathcal{E}_{\pi} [\exp(-\gamma \rho_N^{\pi}(c, i))]$, where γ represents a factor for the level risk aversion.

level of risk sensitivity for both types of policies. Discussed already by Figure 3, the CV@R results also show that the target policies do not limit the risk of obtaining only a few revenues in the worst cases. Further, it is interesting that the policies aimed at targets different from 1000 do not work as well for the 1000 target.

Policy	mean	std. dev.	CV@R	$P(\rho < 1000)$
Target 800	1359.0	167.5	978.8	0.027
Target 900	1348.5	168.1	976.8	0.029
Target 1000	1335.4	168.1	975.8	0.019
Target 1100	1326.1	162.3	973.0	0.028
Target 1200	1325.3	152.0	940.2	0.028
Target 1300	1342.6	165.1	873.4	0.056
Target 1400	1347.4	217.7	777.2	0.082
Utility $\gamma = 0.010$	1361.1	152.6	992.0	0.023
Utility $\gamma = 0.005$	1386.6	166.9	978.6	0.025
Utility $\gamma = 0.001$	1405.3	191.4	943.0	0.033

Table 3: Comparison between two risk-sensitive policies: target percentile optimizing and exponential utility function optimizing policies (the risk aversion increases in conjunction with γ). CV@R is for alpha = 95%.

This effect becomes more observable by the distribution histogram of the 1000 revenue target policy and the expected utility policy with high risk aversion $\gamma = 0.005$, as shown in Figure 4. The target policy has a lower average revenue, a slightly lower 95% CV@R, and a higher standard deviation than the exponential utility policy, but it achieves at least a revenue of 1000 in more cases. The frequencies for the low revenue 700 are little higher for the target policy than for the exponential utility policy and clarify the lower CV@R. The target policy has higher frequencies for revenues between 1000 and 1300 and lower ones between 1400 and 1800 than its counterpart. This explains the lower mean revenue.

Figures 3 and 4 show that the target policies dent the distribution slightly below the target. Thereby, the distribution lower and greater the target is influenced. Frequencies below this dent may increase as frequencies for the target do. In particular, distribution lower then the target need not be modified in a favorable manner regarding the lowest revenues, that is to say the worst cases.



Figure 4: Histogram of the distributions of gained revenue in the simulation. The left blue bars show revenue frequencies gained by target policy with target 1000, the right red bars show frequencies by an expected utility optimizing policy using exponential utility function with $\gamma = 0.05$.

5 Conclusions

A risk-averse policy minimizing the failure of a previously defined, certain revenue target has been proposed for a revenue management problem, namely the dynamic capacity control setting. This policy is derived by extending the state space of the Markov decision process formulation of the problem. We have discussed aspects for implementing the policy numerically.

In numerical experiments, we have analyzed the proposed policy and evaluated against risk-neutral and another risk-sensitive policies. We have compared the mean, standard deviation, and conditional-value-of-risk of those policies. The optimal policy for a given target revenue focuses on minimizing the likelihood of the failing of this certain target but does not compensate for other risk measures. The analysis of the revenue distributions of the target revenue aimed policies in numerical experiments disclose how important correct understanding of such policy is when applied. The decision maker must be aware of its limitations, in particular, that it is the policy with lowest probability of failing the target, but the probability of worst outcomes are not eliminated. However, using a low target revenue supports limiting such risk.

The presented approach can be further developed in order to achieve a policy which optimizes conditional-value-at-risk as proposed by Boda and Filar (2006). Furthermore, it also offers the basis for the development of investigating policies balancing out mean revenue versus target achievement.

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