| 1 | |
|----|---|
| 1 | Supra-threshold auditory brainstem response amplitudes in humans: Test-retest reliability, |
| 2 | electrode montage and noise exposure |
| 3 | |
| 4 | |
| 6 | Garreth Prendergast ^{a,*} , Wenhe Tu ^a , Hannah Guest ^a , Rebecca E. Millman ^{a,b} , Karolina |
| 7 | Kluk ^{a,b} , Samuel Couth ^a , Kevin J. Munro ^{a,b} , Christopher J. Plack ^{a,b,c} |
| 8 | |
| 9 | a, Manchester Centre for Audiology and Deafness, University of Manchester, Manchester |
| 10 | Academic Health Science Centre, M13 9PL, UK. |
| 11 | b, NIHR Manchester Biomedical Research Centre, Central Manchester University |
| 12 | Hospitals NHS Foundation Trust, Manchester Academic Health Science Centre, |
| 13 | Manchester, M13 9WL, UK. |
| 14 | c, Department of Psychology, Lancaster University, Lancaster, LA1 4YF, UK. |
| 15 | |
| 16 | Abstract |
| 17 | The auditory brainstem response (ABR) is a sub-cortical evoked potential in which a series |
| 18 | of well-defined waves occur in the first ten milliseconds after the onset of an auditory |
| 19 | stimulus. Wave V of the ABR, particularly wave V latency, has been shown to be |
| 20 | remarkably stable over time in individual listeners. However, little attention has been paid |
| 21 | to the reliability of wave I which reflects auditory nerve activity. This ABR component has |
| 22 | attracted interest recently, as wave I amplitude has been identified as a possible non- |
| 23 | invasive measure of noise-induced cochlear synaptopathy. The current study aimed to |
| 24 | determine whether ABR wave I amplitude has sufficient test-retest reliability to detect |
| 25 | impaired auditory nerve function in an otherwise normal-hearing listener. Thirty normal- |
| 26 | hearing females were tested, divided into equal groups of low- and high-noise exposure. |
| 27 | The stimulus was an 80 dB nHL click. ABR recordings were made from the ipsilateral |
| 2 | 1 |
| 3 | |

| 28 | mastoid and from the ear canal (using a tiptrode). Although there was some variability |
|----------|--|
| 29 | between listeners, wave I amplitude had high test-retest reliability, with an intraclass |
| 30 | correlation coefficient (ICC) comparable to that for wave V amplitude. There were slight |
| 31 | gains in reliability for wave I amplitude when recording from the ear canal (ICC of 0.88) |
| 32 | compared to the mastoid (ICC of 0.85). The summating potential (SP) and ratio of SP to |
| 33 | wave I were also quantified and found to be much less reliable than measures of wave I |
| 34 | and V amplitude. Finally, we found no significant differences in the amplitude of any wave |
| 35 | components between low- and high-noise exposure groups. We conclude that, if the other |
| 36 | sources of between-subject variability can be controlled, wave I amplitude is sufficiently |
| 37 | reliable to accurately characterize individual differences in auditory nerve function. |
| 38 | |
| 39 | |
| 40 | |
| 41 | Keywords |
| 42 | Auditory brainstem response; test-retest reliability; cochlear synaptopathy; summating |
| 43 | potential; electrode montage |
| 44 | |
| 45 | |
| 46 | Highlights |
| 47 | |
| 48 | ABR wave I and V amplitudes have excellent test-retest reliability in humans |
| 49 50 | SP amplitude and SP/AP ratio have poor test-retest reliability |
| 51 52 | Canal tiptrodes result in only slightly increased reliability re. mastoid electrodes |
| 55 54 | No significant differences in amplitudes between low- and high-noise exposed females |

55 **<u>1. Introduction</u>**

56

The auditory brainstem response (ABR) is a well-established diagnostic tool widely used in 57 58 the clinic to assess auditory function (see Hall, 1992, for an overview). The ABR is evoked 59 by transient stimuli, typically clicks or tone bursts, and consists of a series of waves, with 60 wave I reflecting auditory nerve function, and wave V resulting from generators in the 61 rostral brainstem. The threshold and latency of wave V are the most common clinical 62 metrics of the response. However, wave I has also proved valuable, particularly in 63 research studies, as a more direct measure of peripheral auditory function (Schaette and 64 McAlpine, 2011; Santos et al., 2017).

65

Wave I amplitude has attracted considerable interest recently, following the demonstration 66 67 of noise-induced cochlear synaptopathy in the mouse model by Kujawa and Liberman (2009). In the base of the cochlea, up to 50% of synapses between inner hair cells and 68 69 auditory nerve fibers were destroyed after a 2-hour exposure to 100 dB SPL noise (8-16 70 kHz). Post-exposure measures of absolute auditory sensitivity were unaffected but 71 histological analyses confirmed the dramatic loss of cochlear synapses. Post-exposure 72 ABR measures showed unaffected responses close to threshold. However, at medium-to-73 high sound intensities there was a permanent reduction in the amplitude of wave I of the 74 ABR (by 60% at 32 kHz and ~30% at 12 kHz), reflecting decreased auditory nerve activity.

75

These results suggest that wave I of the ABR might have potential as a non-invasive
measure of cochlear synaptopathy in human listeners. However, the evidence for noiseinduced synaptopathy in humans, based on ABR results, is somewhat inconsistent.
Recent work from our laboratory has found no evidence that greater lifetime noise
exposure, which we assume to be a proxy for greater synaptopathy, is associated with a

8

reduction in ABR amplitude for normal hearing listeners (Prendergast et al., 2017) or 81 82 listeners with tinnitus (Guest et al., 2017). An absence of a relation between noise 83 exposure and ABR wave I amplitude has also recently been reported by a number of other 84 laboratories using different normal-hearing cohorts (Spankovich et al., 2017; Grinn et al., 85 2017: Fullbright et al., 2017). Liberman et al. (2016) also reported no significant reduction 86 in wave I amplitude with increasing noise exposure but did find a significantly increased 87 ratio between the summating potential (SP: reflecting hair cell function) and action 88 potential (AP; equivalent to wave I of the ABR, reflecting auditory nerve function). Bramhall 89 et al. (2017) reported that some groups of firearm users exhibited reduced ABR wave I 90 amplitudes consistent with cochlear synaptopathy and Grose et al. (2017) found a reduced 91 wave I/V ratio in noise-exposed listeners relative to controls. There remain many 92 unanswered questions regarding how these studies can best be reconciled and the extent 93 to which high-frequency hearing loss, gender, and homogeneity of noise exposure can account for the differing evidence for this phenomenon in humans. One additional concern, 94 95 despite the clear changes in ABR wave I in the animal model of synaptopathy, is whether 96 the ABR is the best tool for identifying these neural changes in the human listener.

97

10

98 If the early waves of the ABR are to have utility as a diagnostic measure in individual 99 listeners, they must be reliable, with low measurement error. As ABR wave I amplitude 100 tends to be lower than wave V amplitude, the response may be more difficult to measure 101 reliably (Mehraei et al., 2016). However, there is little available evidence that addresses 102 this issue directly. Much work on the test-retest reliability of the ABR focuses on the 103 latency of wave V because of its clinical relevance. Edwards et al. (1982) provided an overview of ABR amplitude and latency reliability across a six month period, using 72 dB 104 105 nHL (72 dB above the normal adult hearing threshold) monaural clicks in 10 listeners. No 106 significant differences emerged between sessions for any wave amplitudes or latencies, or

11

107 for wave I/V ratios. Using a mean-squared-difference approach, it was found that the 108 participant contributed most variability to the measured responses, followed by ear, 109 session (different days), and run (different acquisition on the same day); however, this was 110 only estimated using wave latency. Lauter and Loomis (1986; 1988) tested seven listeners 111 in eight separate weekly sessions and all waves (I-V) were evaluated. The data show high 112 repeatability across the different testing sessions for both amplitude and latency. Rather than a formal assessment of reliability, the approach used the coefficient of variation (CoV; 113 114 standard deviation divided by the mean) as a marker of "stability" and used ANOVAs to 115 determine that between-subject variability was significantly greater than within-subject 116 variability. Munial at al. (2016) evaluated the long-term test-retest reliability of the ABR in 117 50 normal hearing listeners at 3, 6 and 12 month intervals. Only latencies and inter-peak 118 latencies were studied, which demonstrated good reliability overall, although there were 119 differences in the absolute latency of wave I across the different test intervals.

120

13

121 The studies discussed above all used either linear correlations or ANOVAs to estimate the 122 reliability of ABR responses across multiple sessions. These statistical tools are not formal 123 methods of quantifying reliability, unless the ANOVA is set up in an appropriate manner 124 (Zaki et al., 2012; Kim, 2013). A more appropriate method is to use the intra-class 125 correlation coefficient (ICC; Shrout and Fleiss, 1979), which estimates the proportion of the 126 total variance that can be attributed to between-subject variability. Recently, Bidelman et 127 al. (2017) used the ICC to study the test-retest reliability of sub-cortical and cortical 128 auditory evoked potentials. Wave V of the ABR was evaluated, in response to an 80 dB 129 nHL click stimulus, and the amplitude and latency ICCs were 0.65 and 0.76 respectively, 130 reflecting good test-retest reliability.

131

132 The primary motivation for the current study was to determine the test-retest reliability of

5

14

ABR wave I, to evaluate its suitability for measuring auditory nerve function in individual human listeners. There were also a number of secondary questions which the present study was able to address in parallel to the main research guestion. By using two different EEG montages, a scalp-mounted mastoid electrode and a canal tiptrode (a gold-wrapped foam insert which records the electrical potential from the ear canal), we were able to determine the extent to which reliability is improved by recording from closer to the neural generator of wave I. A canal tiptrode is known to produce a larger wave I response than a scalp-mounted mastoid electrode (Bauch and Olsen, 1990), and it was therefore predicted that the canal tiptrode would produce a more reliable response by virtue of an enhanced signal-to-noise ratio. Furthermore, by using a tiptrode (which emphasizes the SP) we were able to measure the reliability of the SP/AP ratio (utilized by Liberman et al., 2016), and thus evaluate the potential clinical utility of this measure for the detection of synaptopathy.

Finally, the study recruited groups of low- and high-noise exposed female listeners to determine whether changes in the ABR or SP/AP are associated with noise exposure in a single-sex cohort in which audiometric function is tightly controlled. It was predicted that high-noise exposed listeners would yield smaller wave I amplitudes, and larger SP/AP ratios, than low-noise exposed controls.

- 2. Methods

- 2.1. Participants and test sessions

Thirty female participants were tested, all with clinically normal audiometric thresholds (see section 2.3 and Fig. 1). Participants were recruited into two equal-sized groups based on noise exposure histories (see section 2.2). The mean age of participants in the lowexposure group was 23.87 years (range, 19-31) and in the high-exposure group was 24.87 years (range, 20-34). The study was approved by the University of Manchester Research Ethics Committee (project number 16206) and informed, written consent was obtained from all participants.

166

19

167 Testing was conducted over three sessions. Noise exposure estimates and pure tone 168 audiometry were performed in the first session to establish eligibility. The second session 169 (Test 1, T1) consisted of the ABR and distortion product otoacoustic emmission (DPOAE) 170 recordings. The third and final session (Test 2, T2) was a replication of session 2 and was completed on a different day to that of session 2. There were no criteria to constrain how 171 many days elapsed between T1 and T2, provided it was at least 12 hours. Each test 172 173 session took approximately 1 hour. The average number of days between test sessions 174 was 3.5 (s.d. = 3.3; range = 1-12) for the low-noise exposure group and 3.3 (s.d. = 2.7; 175 range = 1-8) for the high-noise exposure group.

- 176
- 177

178 **2.2. Noise exposure**

179

Lifetime noise exposure was estimated using a structured interview developed to assess the effectiveness of the UK noise at work regulations (Lutman et al., 2008). The specific implementation used is described fully in Guest et al. (2017). In summary, participants are asked to consider any high-noise (above ~ 80 dBA) environments/activities to which they have exposed themselves over the course of their lifetime. The duration and frequency of

20

185 exposure is estimated from discussion with the participant and entered into the following186 formula:

 $U = 10^{(L-A-90)/10} \times Y \times W \times D \times H / 2080,$

where U is cumulative noise exposure, L is estimated noise exposure level in dBA, A is attenuation of hearing protection in dB, Y is years of exposure, W is weeks of exposure per year, D is days of exposure per week, H is hours of exposure per day, and 2080 corresponds to the number of hours in a working year. One noise exposure unit is equivalent to exposure for 1 year to a working daily level of 90 dBA. For our purposes, we used the raw units of noise exposure (linearly related to total energy of exposure above 80 dBA) and these were log transformed to produce a normal distribution. Each such logarithmic unit is a factor of 10 in terms of lifetime exposure energy. The cut-off between the low- and high-noise exposure groups was a transformed score of 1.

2.3. Pure tone audiometry

Pure tone audiometry was performed in each ear separately at octave frequencies between 0.25 and 8 kHz in accordance with the British Society of Audiology (2011) recommended procedure. Air-conduction thresholds were measured in a sound-attenuating booth using a Kamplex KC50 audiometer coupled to TDH-39P supra-aural headphones. The audiometric criterion for inclusion in the study was audiometric thresholds < 25 dB HL in both ears at all standard audiometric frequencies. High-frequency audiometric thresholds were also acquired at 12 and 16 kHz using Sennheiser HDA 300 headphones.

2.4 DPOAEs

| 213 | DPOAEs were acquired from both ears using the Otodynamics ILO v6 clinical OAE |
|--|---|
| 214 | software interfaced with a laptop. The ILO probe microphone was calibrated daily using a |
| 215 | 1-cc cavity. The frequency ratio of the two primary tones, f_2/f_1 , was 1.22. Responses were |
| 216 | recorded for f_2 frequencies of 1, 1.5, 2, 3, 4, 6, and 8 kHz. The level of both tones was 70 |
| 217 | dB SPL. The cubic distortion product $(2f_1-f_2)$ amplitude was used as a measure of the |
| 218 | DPOAE. Data collection was terminated after 240 low-noise sweeps had been obtained at |
| 219 | each frequency. A signal-to-noise ratio of 3 dB was required for the DPOAE to be identified |
| 220 | as present. 4% of the DPOAEs were not present (1.4% from the low-noise exposure group |
| 221 | and 2.4% from the high-noise exposure group), and these values were excluded from the |
| 222 | average and the calculation of confidence intervals. |
| 223 | |
| 224 | |
| | |
| 225 | <u>2.5 ABRs</u> |
| 225 226 | 2.5 ABRs 2.5.1 Recording procedure |
| 225 226 227 | 2.5 ABRs 2.5.1 Recording procedure |
| 225 226 227 228 | 2.5 ABRs 2.5.1 Recording procedure Data were recorded using an ICS Chartr EP 200 (Otometrics) and insert earphones |
| 225 226 227 228 229 | 2.5 ABRs 2.5.1 Recording procedure Data were recorded using an ICS Chartr EP 200 (Otometrics) and insert earphones supplied with the system. For both montages the positive electrode was placed at Cz. Two |
| 225 226 227 228 229 230 | 2.5 ABRs 2.5.1 Recording procedure Data were recorded using an ICS Chartr EP 200 (Otometrics) and insert earphones supplied with the system. For both montages the positive electrode was placed at Cz. Two different reference electrodes were used; one coupled to the gold-wrapped insert eartip |
| 225 226 227 228 229 230 231 | 2.5 ABRs 2.5.1 Recording procedure Data were recorded using an ICS Chartr EP 200 (Otometrics) and insert earphones supplied with the system. For both montages the positive electrode was placed at Cz. Two different reference electrodes were used; one coupled to the gold-wrapped insert eartip (canal tiptrode) and one standard electrode mounted on the ipsilateral mastoid. An |
| 225 226 227 228 229 230 231 232 | 2.5 ABRs 2.5.1 Recording procedure Data were recorded using an ICS Chartr EP 200 (Otometrics) and insert earphones supplied with the system. For both montages the positive electrode was placed at Cz. Two different reference electrodes were used; one coupled to the gold-wrapped insert eartip (canal tiptrode) and one standard electrode mounted on the ipsilateral mastoid. An electrode placed on the contralateral mastoid served as the common ground. All electrode |
| 225 226 227 228 229 230 231 232 232 | 2.5 ABRs 2.5.1 Recording procedure Data were recorded using an ICS Chartr EP 200 (Otometrics) and insert earphones supplied with the system. For both montages the positive electrode was placed at Cz. Two different reference electrodes were used; one coupled to the gold-wrapped insert eartip (canal tiptrode) and one standard electrode mounted on the ipsilateral mastoid. An electrode placed on the contralateral mastoid served as the common ground. All electrode impedances were below 5 kΩ and data were sampled at 30 kHz. All recordings were |
| 225 226 227 228 229 230 231 232 233 233 | 2.5 ABRS 2.5.1 Recording procedure Data were recorded using an ICS Chartr EP 200 (Otometrics) and insert earphones supplied with the system. For both montages the positive electrode was placed at Cz. Two different reference electrodes were used; one coupled to the gold-wrapped insert eartip (canal tiptrode) and one standard electrode mounted on the ipsilateral mastoid. An electrode placed on the contralateral mastoid served as the common ground. All electrode impedances were below 5 kΩ and data were sampled at 30 kHz. All recordings were performed by the same researcher to obtain consistent electrode placement, and canal |
| 225 226 227 228 229 230 231 231 232 233 234 235 | 2.5 ABRs 2.5.1 Recording procedure Data were recorded using an ICS Chartr EP 200 (Otometrics) and insert earphones supplied with the system. For both montages the positive electrode was placed at Cz. Two different reference electrodes were used; one coupled to the gold-wrapped insert eartip (canal tiptrode) and one standard electrode mounted on the ipsilateral mastoid. An electrode placed on the contralateral mastoid served as the common ground. All electrode impedances were below 5 kΩ and data were sampled at 30 kHz. All recordings were performed by the same researcher to obtain consistent electrode placement, and canal iptrodes were inserted by the same researcher such that the bottom edge of the foam |

28

238 Clicks were 100 µs in duration and presented in alternating polarity at 80 dB nHL (115.5 dB peSPL) at a rate of 11/s. Stimuli were presented to the right ear, without the left ear 239 240 plugged. Signals were amplified with a gain of 50,000 and band-pass filtered between 0.1 241 and 1.5 kHz (with low- and high-pass roll-offs of 12 dB/octave and 6 dB/octave, 242 respectively). Data were collected over a 20-ms epoch and averaged for a minimum of 243 6000 repetitions. In sessions 2 and 3 (T1 and T2), two such recordings were made within a 244 one-hour period (with the electrodes remaining attached between recordings). The grand average waveform, taken over both acquisitions, was used to characterize the response 245 246 on each day. Participants lay in a comfortable position and were asked to remain still during the recordings. Data were acquired in a sound-treated, but not sound-proofed, 247 248 room.

249

250 **2.5.2 Response identification**

251

252 Three waves were identified in each recording: the SP, wave I and wave V. The average waveform for each listener was subjected to an automated peak- and trough-picking 253 254 procedure based on extracting the phase reversals from the first derivative of the time series (Prendergast et al., 2017). Time windows were constructed around waves I and V 255 256 and the largest identified peak within the window was selected. The center of the window was determined by the peak in the grand average ABR waveform using all 30 participants 257 258 and both montages, which were at 1.70 and 5.60 ms for waves I and V, respectively. The 259 edges of the window were set by using standard deviations of ABR latency reported in response to a 70 dB nHL 100 µs click (Issa and Ross, 1995). Standard deviations were 260 0.17 ms for wave I and 0.21 ms for wave V. The bounds of the windows for our analyses 261 262 were ± 3 standard deviations around the peak central values described above. The SP

10

peak was identified as a peak which occurred 0.5-1.5 ms after stimulus onset. If no peak 263 264 was present in this time window, it was defined as the point at which the first differential of the waveform within this window was lowest, i.e. when the rate-of-change was closest to a 265 phase reversal. Waves I and V were calculated as peak-to-trough, with the trough 266 267 constrained to fall within 2/2.5 ms of the identified peak for waves I and V respectively. If 268 multiple troughs were present, the one which gave the largest peak-to-trough amplitude 269 was used. The SP was defined as being peak-baseline rather than peak to trough. The 270 baseline was calculated as the lowest value in the first 1 ms of the waveform (Liberman et 271 al., 2016). To be consistent with Liberman et al. (2016), the AP values used to compute the 272 SP/AP ratio were peak-baseline values rather than the peak-trough values more commonly used to characterize wave I. This made little difference to the consistency of the 273 274 SP/AP ratio across the test sessions. To make this distinction clear the manuscript will use 275 the terms wave I (peak-to-trough) and AP (peak-to-baseline) to differentiate the two 276 measures. These analyses were performed in Python (version 2.7). S1 of Supplementary 277 Materials provides a schematic of how each wave amplitude was calculated.

278

31

The peaks were visually inspected to ensure that they appeared to select waves I, V and the SP. It was confirmed that the automated procedure was performing appropriately and it was not necessary to redefine any of the peaks.

282

283 **2.6. Statistical metrics**

284

ANOVAs were used to determine if there were differences in wave amplitudes as a

function of noise exposure. The CoV was used as a descriptive statistic of overall

variability for the different groups, montages and the waves. For test-retest reliability,

288 Spearman correlation coefficients were used as a descriptive statistic, but the ICC was

32

289 used to formally quantify the reliability of the measures across the test sessions. The ICC 290 estimates the proportion of total variance that is between-subject rather than between-291 measurements. The ICC uses pooled scaling and standard deviations for the full dataset 292 rather than for each group independently, and is more robust than Pearson correlation 293 coefficients for estimating the correlation in small sample sizes. Furthermore, the 294 assumptions of linearity implicit in a Pearson correlation coefficients can lead to high 295 correlations in cases where the ICC is in fact poor (McGraw and Wong, 1996). There are a 296 number of different formulations for the ICC. Here ICC1 (as defined by Shrout and Fleiss, 297 1979) was used when both observations were from the same montage. ICC1 is sensitive 298 to differences in means between the observations and is a measure of absolute 299 agreement. ICC3 was used when comparing observations between montages, which is 300 insensitive to mean differences and the different observations are treated as fixed effects. 301 In all cases, individual responses were treated as single measures rather than considering 302 the reliability of average responses. All statistical analyses were performed in R (R Core 303 Team, 2015).

304

34

305 <u>3. Results</u>

306

307 3.1. Noise exposure

308

The mean log-transformed noise-exposure score for the low-noise exposure group was -0.98 (std=1.05; min=-3.00; max=0.52) and for the high-noise exposure group was 1.55 (std=0.42; min=1.08; max=2.64). The high-noise exposure group had a mean lifetime exposure energy roughly 340 times that of the low-noise exposure group. The difference in exposure between the two groups was due to a combination of both level and duration. The loudest activities reported by the high-noise group were on average 12.5 dB more

35 26

36

315 intense than those of the low-noise group. The high noise group also reported average 316 exposure durations for a single activity which were 2.5 times longer than those of the low-317 noise group. The high-noise group also typically reported more numerous exposure 318 activities, and so the average total lifetime exposure was three times greater than for the 319 low-noise group. The mean exposure for the low-noise group is equivalent, in terms of 320 total energy, to that for an individual who goes to a nightclub or live music event for 1.5 321 hours, once per year, for five years. The mean high-noise exposure is equivalent to going 322 to the same event for three hours, three times per week, every week of the year, for five 323 years. These exposure values are comparable to those reported by Guest et al. (2017) 324 and there is clear separation between the groups. The high-noise exposure group for the 325 current study was less exposed than the highest-exposed participants reported by 326 Prendergast et al. (2017). Prendergast et al. (2017) tested a large cohort and inspection of these data indicate that when recruiting largely from a University population, a log 327 328 exposure value of 1.5 is high for people aged 18-25, with only 12% of people within this 329 group reporting a log exposure score in excess of 1.5.

330

37

331 **3.2. Pure tone audiometry**

332

333 Fig. 1 shows pure tone audiometric thresholds for the test ear (right ear) of the two groups. 334 The groups appear to be well matched and a mixed design ANOVA with within-subject 335 factors of Ear (two levels; left, right) and Frequency (eight levels; 0.25, 5, 1, 2, 4, 8, 12, 336 and 16 kHz) and a between-subjects factor of Group (two levels; low- and high-noise 337 exposure) confirmed that there is no main effect of Group (F[1,28] = 1.15; p>0.05) nor Ear 338 (F[1,28] = 0.22; p>0.05), but there is a significant main effect of Frequency (F[2,58] =339 14.37; p<0.01). Bonferroni-corrected pairwise comparisons indicate that hearing 340 thresholds at 16 kHz are higher than all other frequencies except 0.25 kHz. Thresholds for

38

0.25 kHz are higher than those at 0.5, 1, 2 and 4 kHz. A significant three-way interaction
between Ear x Frequency x Group was found (F[4,122] = 2.68; p<0.05). The high-noise
group shows higher thresholds at 16 kHz compared to the low-noise exposure group in the
left ear but not the right ear. Since ABRs were acquired from the right ear, the groups were
well matched in terms of audiometric thresholds.

346

347 **3.3. DPOAEs**

348

349 Fig. 2 shows average DPOAE amplitudes for the two groups at each frequency in the test 350 ear (right ear). DPOAEs were collected twice, in the same sessions as the ABR data (T1 351 and T2). A mixed design ANOVA was used with three within-subject factors of Ear (two 352 levels; left and right), Frequency (seven levels; 1, 1.5, 2, 3, 4, 6, and 8 kHz) and Test-retest (two levels; T1 and T2), and a between-subject factor of Group (two levels; low- and high-353 noise exposure). There is no significant main effect of Group (F[1,12] = 0.23; p>0.05), Ear 354 355 (F[1,12] = 0.24; p>0.05), nor Test-retest (F[1,12] = 0.18; p>0.05). A significant main effect of Frequency was found (F[2,28] = 11.12; p<0.01). Bonferroni-corrected pairwise 356 357 comparisons indicate that DPOAE amplitudes at 8 kHz are significantly lower than those at 358 1.5, 2, 4, and 6 kHz. No significant interactions were found (all p>0.05).

359

The attenuation of the response at 8 kHz, equivalent between exposure groups, is most likely related to the difficulties of obtaining reliable DPOAEs at this frequency rather than attributable to a deficit in OHC function. Responses at this frequency are affected by standing waves in the ear canal (Richmond et al., 2011) and the reflectance magnitude tends to be greatest at 8 kHz (Keefe et al., 1993). These factors in conjunction are thought to be responsible for the DPOAE amplitude at 8 kHz often being described as "poor" (Richmond et al., 2011; Gorga et al, 1993, 1997).

41

- 43 367
- 368 **3.4. ABRs**
- 369
- 370 **3.4.1. Effects of session, montage, and group**
- 371

372 Fig 3. shows the grand average ABR waveforms across sessions for the two electrode 373 montages and the two groups of listeners (low- and high-noise exposure). The waveforms appear similar for the two groups. S2 of Supplementary Materials shows the individual 374 waveforms of all 15 listeners in each group, for both electrode montages. Fig. 4 shows the 375 376 average wave I and wave V amplitudes for the two groups for each montage and session, together with the I/V amplitude ratios. There is little difference between the groups or 377 378 sessions. As expected, use of the canal tiptrode montage resulted in larger wave I 379 amplitudes and smaller wave V amplitudes than the mastoid electrode. Equivalent 380 information for wave I and V latency is reported in S3 of Supplementary Materials.

381

Fig. 5 shows average SP values for the two groups in each of the sessions and for both montages, and also SP/AP ratios. The SP values are about 50% larger for the canal tiptrode than the mastoid electrode. However, the SP/AP ratios are comparable in size across the two recording montages, with the difference in the montage means ~ 0.02.

386

Mixed design ANOVAs were used to characterize the response amplitudes for each wave of the response, and the ratio measures, separately. Within-subject factors of Test-retest (two levels; T1 and T2) and Montage (two levels; mastoid electrode, canal tiptrode), and a between-subject factor of Group (two levels; low- and high-noise exposure) were included.

392 For wave I, there is no main effect of Test-retest (F[1,28] = 4.16, p>0.05) nor Group

44

(F[1,28] = 0.14, p>0.05). There is a main effect of Montage (F[1,28] = 209.60, p<0.001)
and Bonferroni-corrected post-hoc tests confirm that wave I amplitudes are greater for the
canal tiptrode than the mastoid electrode. There are no significant interactions between
factors.

397

398 For wave V, there is no main effect of Test-retest (F[1,28] = 0.70, p>0.05) nor Group 399 (F[1,28] = 0.33, p>0.05). There is a main effect of Montage (F[1,28] = 120.68, p<0.001)400 and Bonferroni-corrected post-hoc tests confirm that wave V amplitudes are greater for the 401 mastoid electrode than the canal tiptrode. There are no significant interactions between 402 factors. The wave I/V ratios show no significant interactions and no main effect of Test-403 retest (F[1,28] = 0.09, p>0.05) nor Group (F[1,28] = 1.58, p>0.05). As expected, there is a 404 significant main effect of Montage (F[1,28] = 282.52, p<0.001) with Bonferroni-corrected 405 post-hoc tests indicating the canal tiptrode I/V ratios to be significantly greater than the 406 mastoid electrode ratios.

407

For the SP amplitudes, again there is no main effect of Test-retest (F[1,28] = 0.02, p>0.05) nor Group (F[1,28] = 0.48, p>0.05). There is a main effect of Montage (F[1,28] = 55.36, p<0.001) and Bonferroni-corrected post-hoc tests confirm that SP amplitudes are greater for the canal tiptrode than the mastoid electrode. There are no significant interactions between factors. For the SP/AP ratios there are no significant interactions and no significant main effects of Test-retest (F[1,28] = 0.73, p>0.05), Montage (F[1,28] = 1.42, p>0.05), nor Group (F[1,28] = 2.88, p>0.05).

415

416 **<u>3.4.2. Reliability</u>**

417



47

419 waves and wave ratios of interest. The value reported is the mean CoV for each of

- 420 the two sessions calculated independently.
- 421

| | Mastoid electrode | | Ca | Canal tiptrode | |
|----------------|-------------------|------------|-----------|----------------|--|
| | Low noise | High noise | Low noise | High noise | |
| Wave I | 0.25 | 0.32 | 0.23 | 0.28 | |
| Wave V | 0.33 | 0.19 | 0.33 | 0.16 | |
| Wave I/V ratio | 0.23 | 0.30 | 0.22 | 0.29 | |
| SP | 0.55 | 0.67 | 0.52 | 0.57 | |
| SP/AP ratio | 0.42 | 0.57 | 0.41 | 0.39 | |

422 423

424 Table 1 shows the CoV for the different wave amplitudes and ratios for the two groups and the two sessions. A lower CoV represents less relative dispersion of the data about the 425 426 mean. The lowest coefficients are seen for wave V for the high noise exposure group. 427 Overall, CoVs for wave I are similar to those for wave V (all <0.35), and much less than 428 those for the SP. The coefficients for the canal tiptrode are slightly smaller than for the 429 mastoid electrode, by 0.02 and 0.04 for the low- and high-noise exposure groups, 430 respectively. For wave V the high-noise exposure group shows less variability than the low-noise exposure group in both montages. The CoVs for the ratio measures and the SP 431 432 amplitude are comparable across montages, with the means for each montage differing by 433 no more than 0.1 across all three measures (wave I/V ratio, SP, and SP/AP ratio). The 434 high-noise exposure group shows larger wave I/V ratio variability and greater SP and SP/AP ratio variability. 435

436

Fig. 6 shows wave I and wave V amplitudes for both montages in scatter plots, with
session T2 plotted against session T1. The Spearman correlation coefficient is used as a
descriptive summary statistic of this relation. The low- and high-noise exposure groups are

50

51

plotted in different symbols for consistency, but as there are no statistically significant 440 differences between the groups (see section 3.4.1), all correlations and ICCs were 441 computed across all participants. For wave I, the linear correlation between sessions is 442 443 comparable across the two montages (panels A and B), with a difference of just 0.02. For wave V, the correlation coefficients are 0.04 larger for the mastoid electrode (panel C) than 444 445 for the canal tiptrode (panel D). The correlation between sessions is as strong for wave I as for wave V. The bottom panel of Fig. 6 shows the wave I/V ratio for session T2 plotted 446 against that of session T1. The correlations for the I/V ratio are larger for the canal tiptrode 447 (panel F) than the mastoid electrode (panel E), and similar to those for the individual 448 449 waves shown in the upper two panels of Fig. 6.

450

52

Fig. 7 shows scatter plots for the SP amplitudes and SP/AP ratios. The correlations between sessions are much weaker for the SP than for the main ABR waves. The correlation coefficients are larger for the SP in the canal tiptrode montage (panel B) than for the mastoid electrode (panel A). The bottom panel of Fig. 7 shows the SP/AP ratios for session T2 plotted against those for session T1. The correlations for the SP/AP ratio are slightly larger in the canal tiptrode montage (panel D), though both recording locations show much smaller coefficients than the wave I/V ratio.

- 458
- 459

Table 2. ICC values for five ABR amplitude measures, both between sessions (for
 both electrode montages) and between montages. Lower and upper 95% confidence
 intervals are shown in parentheses.

463

| | Mastoid electrode | Canal tiptrode | Between montage |
|--------|-------------------|------------------|------------------|
| Wave I | 0.85 (0.71/0.92) | 0.88 (0.76/0.94) | 0.88 (0.80/0.94) |
| Wave V | 0.80 (0.63/0.90) | 0.87 (0.75/0.94) | 0.90 (0.82/0.95) |

| Wave I/V ratio | 0.84 (0.70/0.92) | 0.89 (0.79/0.95) | 0.85 (0.74/0.92) |
|----------------|--------------------|-------------------|-------------------|
| SP | 0.18 (-0.18/0.50) | 0.40 (0.056/0.66) | 0.47 (0.25/0.67) |
| SP/AP ratio | 0.32 (-0.039/0.60) | 0.46 (0.13/0.70) | 0.31 (0.083/0.54) |
| | | | |

464

ICC values are shown in Table 2, together with 95% confidence intervals. The ICCs are 465 466 largest for waves I, V, and the I/V ratio, and largest for the canal tiptrode montage. These ICC values would generally be described as reflecting excellent repeatability (>0.75; 467 468 Cicchetti, 1994), both within and between montages. The reliability of wave I across the 469 two test sessions is comparable to that for wave V, with all ICC values greater than 0.80. Wave I amplitudes are larger for the canal tiptrode montage, but it does not appear that 470 this is concordant with a substantial increase in reliability over the mastoid electrode 471 472 montage. ICC values for wave I and V latency are reported in S4 of Supplementary 473 Materials.

474

475 The SP and SP/AP ratio measures show much lower reliability. The SP for the mastoid 476 electrode has poor reliability, and although this is improved by using the SP/AP ratio, it still remains lower than the reliability reported for the other waves. The SP values from the 477 canal tiptrodes are more reliable and these are also improved by using a ratio measure, 478 479 although, as indicated by the confidence intervals, there is no statistically significant 480 difference between the reliability of the two montages for any of the measured waves or 481 ratios. However, it is clear that any measure utilising the SP is much less reliable than one 482 using waves I and V. The strongest ICC value of the four measures involving the SP (restricted to reliability estimates within a montage) is 0.46. Comparing this ICC value with 483 484 the weakest ICC from the three measures using waves I and V (0.80) demonstrates that the reliability of measures utilising the SP are significantly poorer than those using waves I 485 and V (z = 2.21; p<0.05). 486

487

56 57

488 <u>4. Discussion</u>

489

490 **4.1 Reliability of ABR measures**

491

492 The primary aim of the current study was to quantify the test-retest reliability of ABR 493 measures, to evaluate whether the ABR is a suitable technique for measuring auditory 494 nerve function in individual human listeners. Although it has been reported that the ABR is 495 stable over long time periods in an individual, much of this evidence relates to wave V. The 496 data presented here indicate that wave I test-retest reliability, and therefore measurement 497 error, is comparable to that of the larger amplitude wave V. Therefore, although wave V is often characterised as robust and reliable, and wave I as small and variable (Mehraei et 498 499 al., 2016), it is clear that wave I has high within-subject reliability in normal-hearing 500 listeners, at least for the stimulus intensity used here. If the other sources of between-501 subject variability (for example, head size, tissue resistance) can be controlled, wave I 502 amplitude is sufficiently reliable to accurately characterize individual differences in auditory 503 nerve function.

504

505 Neither the SP nor SP/AP ratio were reliable. Even when using the canal tiptrode montage, 506 the best-case ICC was 0.46. In the current study these measures clearly have poor test-507 retest reliability, but this may be because of the small SP amplitudes evoked by an 80 dB 508 nHL (115.5 peSPL) click. The click used by Liberman et al. (2016) to evoke the SP had a 509 level of 94.5 dB nHL(130 peSPL), and produced much larger SP amplitudes. However, it is 510 not clear that raising presentation levels to enhance the SP is advisable. Even an 80 dB nHL stimulus is intolerably loud for some listeners (Gu et al., 2012). A stimulus 511 512 presentation level greater than 90 dB nHL (over 120 dB peSPL when presented through 513 ER3A inserts) could risk exceeding recommended daily exposure limits after a few

59 60

514 thousand presentations. Moreover, even such exposure limits may be too permissive, since impulse noise is more damaging than continuous-type noise of equivalent energy 515 516 (Starck et al., 2003). It may also be the case that the SP is inherently unreliable, even if 517 higher stimulus presentation levels are used. Either way, the clinical utility of the SP 518 measure may be limited.

- 519
- 520

521 The SP/AP ratio in the current study used an arbitrary baseline to compute the amplitude 522 of both the SP and the AP components, as described by Liberman et al. (2016). It has 523 been reported previously that peak-baseline measures of wave I amplitude (the AP) are 524 less reliable than peak-trough estimates of amplitude (Stelmack et al., 2003). Therefore, 525 measures such as the SP/AP ratio could benefit from using peak-trough estimates of the 526 AP. However, in the current study this made little difference to the reliability of the SP/AP 527 ratio, which suggests that the variability of the SP was the limiting factor.

- 528
- 529
- 530

531 4.2. Effects of electrode montage

532

One concern when trying to measure small, supra-threshold changes in the auditory nerve 533 534 function of normal-hearing listeners is that scalp-mounted mastoid electrodes are simply 535 not sensitive enough to reliably detect the subtle changes in evoked responses. The 536 results presented in this study indicate that moving the recording site closer to the generator of wave I (the auditory nerve), by placing a tiptrode in the ear canal, produced 537 only a small increase in reliability for waves I and V, although the benefit was greater for 538 539 the SP. The amplitude of wave I increased and that of wave V decreased when using a

62

canal tiptrode relative to a mastoid electrode, as seen in other studies (e.g. Bauch and
Olsen, 1990). However, reliability of the wave amplitude did not appear to be directly linked
to the absolute amplitude of the wave. Wave V was slightly more reliable in the canal
tiptrode montage compared to the mastoid electrode montage, despite having lower
amplitudes on average. Given that the use of canal tiptrodes increases the financial
burden on ABR practitioners and can reduce participant comfort, it is not clear that such
equipment is necessary or advisable for the recording of ABR waves I or V.

547

64

548 **4.3. Relation of ABR measures to noise exposure**

549

550 The final aim of the study was to investigate supra-threshold changes in the ABR in 551 relation to noise exposure. The results presented here, for a group of young females in which low- and high-noise exposed listeners were well-matched for audiometric thresholds 552 and age, indicate no changes in wave I amplitude as a function of noise exposure. There 553 554 is no evidence for noise-induced cochlear synaptopathy. This is consistent with other 555 recent studies in our laboratory which have found no association between noise exposure 556 and wave I amplitude in young listeners with normal audiograms (Prendergast et al., 2017; 557 Guest et al., 2017). The range of noise exposures in the present study allowed for good separation between the groups, although compared with Prendergast et al. (2017) there 558 were fewer listeners with very high exposures, and more listeners with very low exposures. 559 560 It should be noted that an absence of any evidence for cochlear synaptopathy is not the 561 same as evidence for absence of the disorder. It remains unclear how sensitive the ABR is 562 to a loss of low-SR fibers, even in animals (Bourien et al., 2014). Shaheen et al. (2015) suggested that the frequency-following response is a more sensitive identifier of cochlear 563 synaptopathy than the ABR. It may yet prove that in humans, a click-evoked response is 564 565 too crude a measure with which to elucidate subtle supra-threshold, sub-clinical deficits.

65 66

567 Liberman et al. (2016) also reported no significant difference in wave I amplitude between low- and high-noise exposed groups of listeners, although they did find a large difference 568 569 between the groups in the SP/AP ratio. Liberman et al. reported mean SP amplitudes of 570 approximately 0.14 and 0.21 µV, and SP/AP ratios of 0.26 and 0.46, for the low- and high-571 noise exposure groups, respectively. For the canal tiptrode montage in the present study, 572 the SP amplitudes were 0.07 and 0.08 µV, and the SP/AP ratios were 0.22 and 0.26, for the low- and high-noise exposure groups, respectively. Although the present data show a 573 trend in the direction reported by Liberman et al., the effect did not reach significance. The 574 575 click intensity used in the current study was 14.5 dB lower than that used by Liberman et 576 al., and therefore it may be that substantial differences between noise-exposure groups 577 are only observed for more intense presentation levels than used here. Alternatively, there 578 were substantial high-frequency audiometric differences between the groups in the 579 Liberman et al. study, in contrast to the present study in which the groups were closely 580 matched at high frequencies. Hence the populations tested in the two studies may not be 581 directly comparable. One possibility is that high-frequency audiometric loss is a marker for 582 cochlear synaptopathy. For example, only noise exposures that produce high-frequency 583 threshold elevations may have the capacity to cause a substantial loss of cochlear 584 synapses. Another is that SP/AP ratios may be directly influenced by high-frequency 585 sensitivity, in the absence of synaptopathy. It may also be crucial to consider age more 586 carefully, for example, whether the age at which intense noise exposures are experienced 587 is critical, or whether the effects of noise-induced synaptopathy are more easily observed 588 as an accelerated decline in hearing with advancing age.

589

590 <u>5. Conclusions</u>

- 591
 - 68 69

| 70 | |
|----------|--|
| 592 | • For young female listeners with normal hearing, ABR wave I and wave V |
| 593 | amplitudes, and the I/V amplitude ratio, all show excellent test-retest reliability, with |
| 594 | over 80% of the variability in measurement accounted for by between-subject |
| 595 | differences in ABR response. |
| 596 | • The SP amplitude and SP/AP ratio show poor levels of reliability for the 80 dB nHL |
| 597 | click intensity used here. |
| 598 | • Use of a canal tiptrode may result in slightly improved reliability, although a mastoid |
| 599 | electrode is still highly reliable for waves I and V. |
| 600 | No significant differences were found in any ABR measure between low- and high- |
| 601 | noise exposure groups. |
| 602 | |
| 603 | Acknowledgments |
| 604 | This work was supported by the Medical Research Council, UK (MR/L003589/1). Author |
| 605 | H.G. was supported by an Action on Hearing Loss studentship funded by the Marston |
| 606 | Family Foundation. Author S.C. was supported by the Colt Foundation. We thank the |
| 607 | Associate Editor and two anonymous reviewers for constructive comments on earlier |
| 608 | drafts of this manuscript. |
| 609 | |
| 610 | |
| 611 | |
| 612 | |
| 613 | Figure Captions |
| 614 | Fig. 1. Pure tone air conduction audiometric thresholds. Thresholds are shown for the test |
| 615 | ear, with 95% confidence intervals, for the two groups of listeners. $N = 15$ in each group. |
| 616 | |
| 617 | Fig. 2. Distortion product otoacoustic emissions. DPOAEs from a single session (T1) for |
| 71 72 | 24 |

618 the test ear are shown, with 95% confidence intervals, for the two groups of listeners.

619

Fig. 3. Grand average ABR waveforms in response to a 80 dB nHL click. Waveforms are
shown for each group of listeners and for the mastoid electrode and canal tiptrode. 95%
confidence intervals are indicated by the shaded areas.

623

Fig. 4. Mean peak-to-trough amplitudes for wave I and wave V, and mean wave I/V ratios.
Each test session is plotted individually for the two montages and the two groups. Error
bars show 95% confidence intervals.

627

Fig. 5. Mean peak-to-baseline amplitudes for the SP and the SP/AP ratio. Each session is
plotted individually for the two montages and the two groups. Error bars show 95%
confidence intervals.

631

Fig. 6. Test-retest reliability of waves I and V, and I/V ratio. Amplitudes and ratios for the second test session (T2) are plotted against those for the first test session (T1). The data for the mastoid electrode and canal tiptrode are plotted in the left- and right-hand column respectively. Spearman correlation coefficients are reported as a summary statistic. Lownoise exposed listeners are shown in open green circles and high-noise exposed listeners in filled grey circles. The diagonal line represents the ideal relation across both test sessions.

639

Fig. 7. Test-retest reliability of the SP and SP/AP ratio. Amplitudes and ratios for the
second test session (T2) are plotted against those for the first test session (T1). The data
for the mastoid electrode and canal tiptrode are plotted in the left- and right-hand column
respectively. Spearman correlation coefficients are reported as a summary statistic. Low-

74 75

- 644 noise exposed listeners are shown in open green circles and high-noise exposed listeners
- 645 in filled grey circles. The diagonal lines represent the ideal relation (perfect reproducibility)
- 646 between test sessions.

References

- Bauch, C.D., Olsen, W.O., 1990. Comparison of ABR amplitudes with TIPtrode and
 mastoid electrodes. Ear Hear. 11, 463-467.
- Bidelman, G.M., Pousson, M., Dugas, C., Fehrenbach, A., 2017. Test–Retest reliability of
 dual-recorded brainstem versus cortical auditory-evoked potentials to speech. J.
 Am. Acad. Audiol. In press.
- Bourien, J., Tang, Y., Batrel, C., Huet, A., Lenoir, M., Ladrech, S., Desmadryl, G., Nouvian,
 R., Puel, J.L., Wang, J., 2014. Contribution of auditory nerve fibers to compound
 action potential of the auditory nerve. J. Neurophysiol. 112, 1025-1039.
- Bramhall, N.F., Konrad-Martin, D., McMillan, G.P., Griest, S.E., 2017. Auditory brainstem
 response altered in humans with noise exposure despite normal hair cell function.
 Ear Hear. 38, e1-e12.
- British Society of Audiology, 2011. Pure-tone air-conduction and bone-conduction
 threshold audiometry with and without masking. Recommended procedure (British
 Society of Audiology, Reading, UK), 32.
- Cicchetti, D.V., 1994. Guidelines, criteria, and rules of thumb for evaluating normed and
 standardized assessment instruments in psychology. Psychological Assessment. 6,
 284-290.
- 673 Edwards, R.M., Buchwald, J.S., Tanguay, P.E., Schwafel, J.A., 1982. Sources of variability

| in auditory brain stem evoked potential measures over time. Electroencephalogr. |
|---|
| Clin. Neurophysiol. 53, 125-132. |
| |
| Fullbright, A.N.C., Le Prell, C.G., Griffiths, S.K., Lobarinas, E., 2017. Effects of recreational |
| noise on threshold and suprathreshold measures of auditory function. Semin. Hear. |
| 38, 298-318. |
| |
| Gorga, M.P., Neely, S.T., Bergman, B., Beauchaine, K.L., Kaminski, J.R., Peters, J., |
| Jesteadt, W., 1993. Otoacoustic emissions from nornal-hearing and hearing- |
| impaired subjects: distortion product responses. J. Acoust. Soc. Am. 93, 2050-2060. |
| |
| Gorga, M.P., Neely, S.T., Ohlrich, B., Hoover, B., Redner, J, Peters, J., 1997. From |
| laboratory to clinic: a large scale study of distortion product otoacoustic emissions |
| in ears with nornal hearing and ears with hearing loss. Ear Hear. 18, 440-455. |
| |
| Grinn, S.K., Wiseman, K.B., Baker, J.A., Le Prell, C.G., 2017. Hidden hearing loss? No |
| effect of common recreational noise exposure on cochlear nerve response |
| amplitude in humans. Font. Neurosci. 11: 465. |
| |
| Grose, J.H., Buss, E., Hall III, J.W., 2017. Loud music exposure and cochlear |
| synaptopathy in young adults: Isolated auditory brainstem response effects but no |
| perceptual consequences. Tends in Hearing. 21, 1-18. |
| |
| Gu, J.W., Herrmann, B.S., Levine, R.A., Melcher, J.R., 2012. Brainstem auditory evoked |
| potentials suggest a role for the ventral cochlear nucleus in tinnitus. J. Assoc. Res. |
| Otolaryngol. 13, 819-833. |
| |

| 700 | |
|----------|--|
| 701 | Guest, H., Munro, K.J., Prendergast, G., Howe, S., Plack, C.J., 2017. Tinnitus with a |
| 702 | normal audiogram: Relation to noise exposure but no evidence for cochlear |
| 703 | synaptopathy. Hear. Res. 344, 265-274. |
| 704 | |
| 705 | Hall, J. W. (1992). Handbook of auditory evoked responses. Boston: Allyn and Bacon. |
| 706 | |
| 707 | Issa, A., Ross, H.F., 1995. An improved procedure for assessing ABR latency in young |
| 708 | subjects based on a new normative data set. Int. J. Pediatr. Otorhinolaryngol. 32, |
| 709 | 35-47. |
| 710 | |
| 711 | Keefe, D.H., Bulen, J.C., Arehart, K.H., Burns, E.M., 1993. Ear-canal impedance and |
| 712 | reflection coefficient in human infants and adults. J. Acoust. Soc. Am. 94, 2617-2638. |
| 713 | |
| 714 | Kim, HY. (2013). Statistical notes for clinical researchers: Evaluation of measurement |
| 715 | error 1: using intraclass correlation coefficients. Restorative Dentistry & |
| 716 | Endodontics, 38, 98–102. |
| 717 | |
| 718 | Kujawa, S.G., Liberman, M.C., 2009. Adding insult to injury: cochlear nerve degeneration |
| 719 | after "temporary" noise-induced hearing loss. J. Neurosci. 29, 14077-14085. |
| 720 | |
| 721 | Lauter, J.L., Loomis, R.L., 1986. Individual differences in auditory electric responses: |
| 722 | Comparisons of between-subject and within-subject variability I. Absolute latencies |
| 723 | of brainstem vertex-positive peaks. Scand. Audiol. 15, 167-172. |
| 724 | |
| 725 | Lauter, J.L., Loomis, R.L., 1986. Individual differences in auditory electric responses: |
| 86 87 | 29 |

| 88 | |
|-----|---|
| 726 | Comparisons of between-subject and within-subject variability II. Amplitude of |
| 727 | brainstem vertex-positive peaks. Scand. Audiol. 17, 87-92. |
| 728 | |
| 729 | Liberman, M.C., Epstein, M.J., Cleveland, S.S., Wang, H., Maison, S.F., 2016. Toward a |
| 730 | differential diagnosis of hidden hearing loss in humans. PLoS ONE, 19, e0162726. |
| 731 | |
| 732 | Lutman, M.E., Davis, A.C., Ferguson, M.A., 2008. Epidemiological evidence for the |
| 733 | effectiveness of the noise at work regulations. Health and Safety Executive. |
| 734 | Research Report RR669. |
| 735 | |
| 736 | Mcgraw, K., Wong, S.P., 1996. Forming inferences about some intraclass correlation |
| 737 | coefficients. Psychological Methods. 1. 30-46. |
| 738 | |
| 739 | Mehraei, G., Hickox, A.E., Bharadwaj, H.M., Goldberg, H., Verhulst, S., Liberman, M.C., |
| 740 | Shinn-Cunningham, B.G., 2016. Auditory Brainstem Response Latency in Noise as |
| 741 | a Marker of Cochlear Synaptopathy. J. Neurosci. 36, 3755–3764. |
| 742 | |
| 743 | Munjal, S., Panda, N., Pathak, A., 2016). Long Term Test-Retest Reliability of Auditory |
| 744 | Brainstem Response (ABR) and Middle Latency Response (MLR). Glob. J. Oto. 1, |
| 745 | 555559. |
| 746 | |
| 747 | Prendergast, G., Guest, H., Munro, K.J., Kluk, K., Leger, A., Hall, D.A., Heinz, M.G., Plack, |
| 748 | C.J., 2017. Effects of noise exposure on young adults with normal audiograms I: |
| 749 | Electrophysiology. Hear. Res. 344, 68-81. |
| 750 | |
| 751 | R Core Team . R Foundation for Statistical Computing; Vienna, Austria: 2015. R: a |
| | |

Language and Environment for Statistical Computing.

- Richmond, S.A., Kopun, J.G., Neely, S.T., Tan, H., Gorga, M.P., 2011. Distribution of
 standing-wave errors in real-ear sound-level measurements. J. Acoust. Soc. Am.
 129, 3134-3140.
- Santos, M., Marques, C., Nobrega Pinto, A., Fernandes, R., Coutinho, M.B., Almeida E
 Sousa, C., 2017. Autism spectrum disorders and the amplitude of auditory
 brainstem response wave I. Autism Research. 10, 1300-1305.
- Schaette, R., McAlpine, D., 2011. Tinnitus with a normal audiogram: physiological
 evidence for hidden hearing loss and computational model. J. Neurosci. 31, 1345213457.
- Shaheen, L.A., Valero, M.D., Liberman, M.C., 2015. Towards a diagnosis of cochlear
 neuropathy with envelope following responses. J. Assoc. Res. Otolaryngol. 16, 727 768 745.
- Shrout, P.E., Fleiss, J.L., 1979. Intraclass correlations: uses in assessing rater reliability.
 Psychol. Bull. 86, 420-428.
- Spankovich, C., Griffiths, S.K., Lobarinas, E., Morgenstein, K.E., de la Calle, S., Ledon, V.,
 Guerico, D., Le Prell, C.G., 2017. Temporary threshold shift after impulse-noise
 during video game play: laboratory data. Int. J. Audiol. S53-65.

577 Starck, J., Toppila, E., Pyykko, I., 2003. Impulse noise and risk criteria. Noise and Health

| 778 | 5, 63-73. |
|-----|-----------|
| | - |

| 7 | 7 | 9 |
|---|---|---|
| / | / | 9 |

| 780 | Stelmack, R.M., Knott, V., Beauchamp, C.M., 2003. Intelligence and neural transmission |
|-----|---|
| 781 | time: a brain stem auditory evoked potentials analysis. Personality and Individual |
| 782 | Differences. 34, 97-107. |
| 783 | |
| 784 | Zaki, R., Bulgiba, A., Nordin, N., & Azina Ismail, N., 2013, A systematic review of statistical |

| /04 | Zaki, N., Bulyiba, A., Norulli, N., & Azina Ismail, N., 2013. A systematic review of statistica |
|-----|---|
| 785 | methods used to test for reliability of medical instruments measuring continuous |
| 786 | variables. Iranian Journal of Basic Medical Sciences, 16, 803–807. |





Mastoid Electrode



Canal Tiptrode









SP / AP



Session T2 Amplitude (µV)

| | ACCIAN | |
|---|---------|--------|
| J | COSIDII | ιλάτιο |

Mastoid Electrode



